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A NEW METHODOLOGY FOR FLOOD FREQUENCY ANALYSIS WITH OBJECTIVE DETECTION AND MODIFICATION OF OUTLIERS/INLIERS

Bу

Krishan P. Singh, Ph.D., Principal Scientist Masahiro Nakashima, Graduate Research Assistant



Champaign, Illinois

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ABSTRACT: Prerequisites for deriving satisfactory estimates of design floods are the objective detection and modification of outliers/inliers at desired levels of significance and a versatile technique of flood frequency analysis. Objective detection and modification of outliers/inliers has been accomplished by developing statistics for outliers/inliers at both the high and low end of the flood spectrum. An inlier at the high end is defined as a variate (generated or observed) which is lower than indicated by the trend of the rest of the data, and an inlier at the low end is higher than indicated by the rest of the data. The tests developed in this study can be used to check for outliers/inliers at different levels of significance, such as 0.01, 0.05, 0.1, 0.2, 0.3, and 0.4.

Three transformations for converting an observed flood series to an approximately normally distributed series were tested on flood series at 28 gaging stations in Illinois. The transformations considered were the power, Wilson-Hilferty, and 3-parameter log transformation. Analyses of the transformed series indicate that power transformation is the best of the three tested. The observed flood series is converted to a quasi normally-distributed series with the power transformation and, then, the statistical tests are applied for detection and modification of any outliers/inliers at various levels of significance.

Flood frequency methodologies (normal distribution after power transformation, log Pearson type III distribution, and mixed distribution) were tested on flood series observed at 37 gaging stations in Illinois. These analyses indicate that 1) regionalization of skew values alone, as recommended by the Water Resources Council in their Bulletin 17, is not a satisfactory solution to flood frequency problems, 2) the outlier criterion as recommended in Bulletin 17 is too severe, 3) an observed flood series needs to be checked for both inliers and outliers, 4) the power transformed series can have kurtosis lower or higher than 3 (it is 3 for a normal distribution) and the kurtosis correction can be applied if the transformed series is symmetrical, 5) the 37 power-transformed series exhibited asymmetry insofar as 5th and higher order odd moments were not zero, 6) only the mixed distribution can account for asymmetry displayed by the transformed series, and 7) the mixed distribution applied to series after detection and modification of outliers yields design flood estimates which exhibit regional consistency.

The versatile flood frequency analysis with the mixed distribution, coupled with objective detection and modification of outliers at various levels, provides a very satisfactory solution to the flood frequency problem. The method has been written as an efficient computer program.

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INTRODUCTION

"An accurate estimate of the flood potential is a key element to an effective nationwide flood damage abatement program. To obtain both a consistent and accurate estimate of flood losses requires development, acceptance and widespread application of a uniform, consistent and accurate technique for determining flood-flow frequencies."

The above excerpt from the Foreword of Bulletin 17A of the Hydrology Committee, U.S. Water Resources Council, stresses the need of a uniform, consistent and accurate technique for flood frequency analyses. From their analyses and research in the last two decades, the Council published bulletins (1967, 1976, and 1977) containing guidelines for determining flood frequency. They have recommended the fitting of log-Pearson type III, or LP3, distribution to observed annual flood peaks. The method of moments is used to determine the statistical parameters of the distribution from station data. Generalized relationships are used to define the skew coefficient for short record stations. Methods are proposed for treatment of some flood record problems encountered. The problem of outliers is recognized and it is dealt with in respect to outliers at both the low and high end. For the existence of low outliers, the criterion is

$$\left|\frac{x-\bar{x}}{s}\right| > [2.5 + 1.2 \log (n/10)](1 - 0.4 g_r)$$
(1)

in which $x = \log Q$, Q is an annual flood, x and s are the mean and standard deviation of log-transformed floods, n is the sample size, and g_r is the regional skew coefficient. The generalized skew coefficient for Illinois (with the exception of the lower portion of southern Illinois) is -0.4

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(figure 1). The test statistic or the left-hand side of equation 1 needs to be greater than 3.45 and 3.87 for n = 25 and 50, respectively. Use of this criterion is equivalent to rejection at the 1 percent level of significance. When one or more outliers are identified, they are deleted from the record and the remaining record is treated as an incomplete record. If a high outlier is suspected, a comparison with historical flood data and flood information at nearby sites is made. If such information is available, a plotting position is assigned to each outlier and the procedure for historic floods is used; otherwise the outlier is retained as it is in the basic computations.

Previous Study

A study on the regional and sample skew values in flood-frequency analyses and the effect of outliers on the distribution parameters was conducted at the State Water Survey (Singh, 1980). In this study, the storm, basin, stream, soil, floodplain, and other relevant factors were investigated for 62 basins in areas drained by the Sangamon, Rock, and Little Wabash Rivers in Illinois (figure 1), to understand the variation in skew values from a number of annual flood series. Various flood frequency analyses were conducted for the annual flood series at each of the 62 study basins. The main conclusions drawn from this study were: 1) the criterion for a low outlier, as given in Bulletin 17A of the Water Resources Council, is too severe and needs modification — it yielded no outliers in any of the 62 series; 2) one or two very low floods greatly decrease the skew value and distort the fitted distribution curve — such low floods were found in about 30 percent of the flood series analyzed; 3) a high outlier can be confirmed

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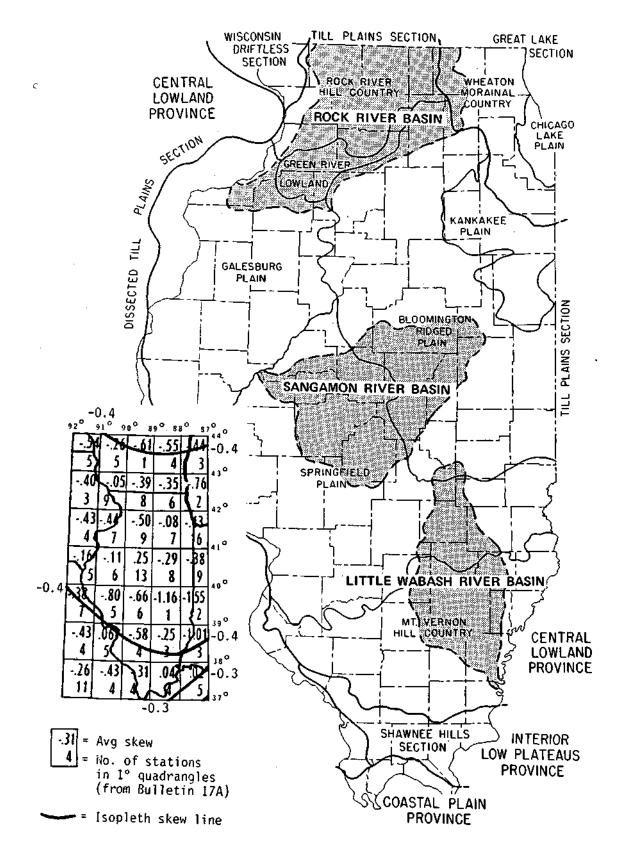


Figure 1. Study basins and physiographic divisions in Illinois, and skew coefficients from Bulletin 17A

by statistics of the storm producing it; 4) some better tests need to be developed for identifying or perceiving low and high outliers; 5) the methodology developed in this study for modifying outliers is generally satisfactory - this methodology depends on specifying the floods perceived as outliers by the analyst; 6) the modification of outliers developed for the LP3 changes both the standard deviation and skew, and regionalization of both the parameters may be needed instead of the skew alone; 7) the observed floods should be plotted using the best statistical plotting position instead of the commonly used Weibull plotting position, for checking the fit of the derived distribution curve with the observed data; 8) the standard deviation and skew appear to be correlated with basin and stream characteristics; 9) a change in the flow-section characteristics when river discharge begins inundating the floodplain introduces a new storage element that can affect the distribution shape of the observed floods; and 10) the distribution parameters (mean, standard deviation, and skew) below the junction of two major tributaries are affected greatly by the degree of concurrency of tributary flood peaks in time and flow magnitude.

The most important conclusions from this study were the need for developing tests to detect outliers/inliers at various levels of significance and better flood-frequency methods, and the inappropriateness of regionalizing skew values and using them in flood-frequency analyses without consideration of atypical hydrologic and hydraulic conditions.

Present Study

The main objectives of the study presented in this report are: 1. Development of statistical tests for outliers and inliers: An

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inlier at the high end is a flood which is lower than indicated by the trend of the rest of the data and an inlier at the low end is a flood which is higher than indicated by the rest of the data (figure 2). These tests should check for outliers/inliers at different levels of significance, such as 0.01, 0.05, 0.1, 0.2, 0.3 and 0.4.

2. The statistical tests will be developed for the normally distributed series because of fewer number of distribution parameters. The observed flood series will be transformed to a series distributed as N (μ , σ^2). The available transformations will be tested to choose the best.

3. Various methods of analyzing floods will be reviewed and their advantages and disadvantages examined. Their theoretical development, practical use, and any basic assumptions will be considered.

4. The flood-frequency methods will be computerized in a general package which will include testing for inliers/outliers and modification of any inliers/outliers detected at various significance levels. The results obtained with the use of 30 or more annual flood series from the Rock, Sangamon, and Little Wabash River sub-basins will be compared to determine the best method.

All of the objectives of this study have been met. Statistical tests for outliers/inliers at various levels of significance have been developed from extensive use of random number generators. The transformation technique that consistently and efficiently converts an observed flood series to a normally distributed series has been found. A new flood frequency methodology has been developed. It is much better than the others tested. The methodology yields flood estimates at various recurrence intervals with outliers/inliers detected and modified at various levels of significance.

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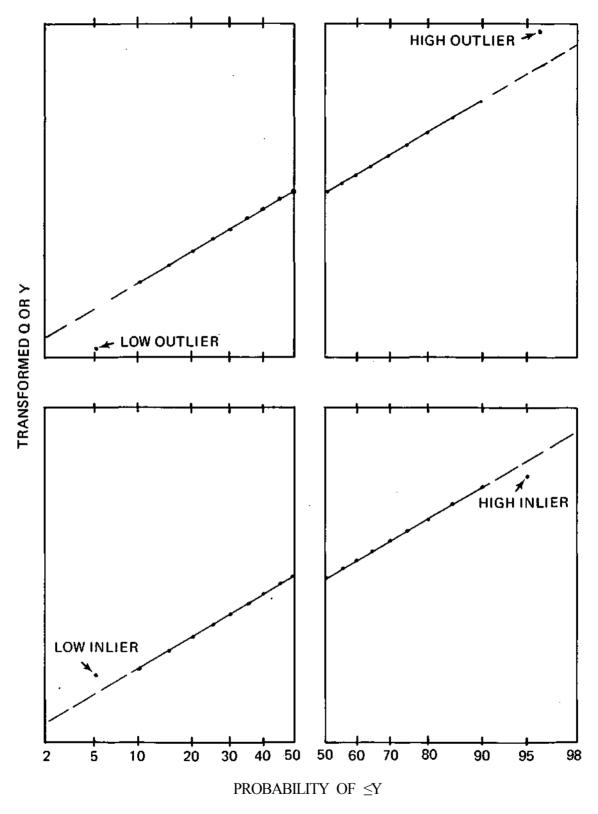


Figure 2. Definition sketch for low and high outliers and inliers

Acknowledgments

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Ganapathi Ramamurthy, part-time graduate research assistant, helped in analyses of various transformation methods. John W. Brother, Jr. supervised the preparation of illustrations.

STATISTICAL TESTS FOR OUTLIERS AND INLIERS

An outlier in a set of data is defined as an observation or subset or observations, which appears to be inconsistent with the remainder of that set of data (Barnett and Lewis, 1978). The inconsistency can be interpreted as the observation being either significantly higher or lower at the high end (or lower or higher at the low end) than the value indicated by the rest of the data; the observation will be termed as an outlier or an inlier, respectively. The outlier can depart considerably from the assumed underlying distribution curve but the inlier departs by a lesser amount because the next observation can replace an inlier. In conventional flood-frequency analyses, it has been a matter of subjective judgment on the part of the analyst whether or not he picks up some observation for scrutiny. As stated earlier in the text, the criterion given for outlier detection in Bulletin 17 of the U.S. Water Resources Council is too severe. Literature search did not show the existence of statistical tests for checking outliers (at higher than 5% level) and inliers at different probability levels of their occurrence. The development of suitable statistical tests, detailed hereafter in this section, was an important part of this study.

Generation of Normally Distributed Random Numbers

Four methods or algorithms for generating normally distributed random numbers were tested extensively regarding their suitability, stability, and effectiveness in generating such numbers. A brief background of these methods is given here.

Box and Muller Method (BAMM)

Box and Muller (1958) presented a method of generating normally distributed random numbers, X_1 and X_2 , with zero mean and unit variance:

$$X_{1} = (-2 \ln U_{1})^{\frac{1}{2}} \cos 2\pi U_{2}$$
(1)

$$X_{2} = (-2 \ln U_{1})^{\frac{1}{2}} \sin 2\pi U_{2}$$
 (2)

in which U_1 and U_2 are random numbers drawn from a uniform or rectangular distribution function, U (0, 1), and ln is the natural logarithm. X_1 and X_2 are a pair of independent random variables such that

$$f(X_1, X_2) = f(X_1) f(X_2)$$
(3)

According to Box and Muller, this scheme should generate normal random numbers which are more reliable in the two extreme tails of the distribution. *The Polar Method (PLRM)*

Box, Muller, and Marsaglia (Knuth, 1969) presented a method, commonly known as the polar method, for calculating two independent normally distributed variables, X_1 and X_2 , from two independent random numbers from a uniform distribution, U (0,1). Computation of these variables follows the procedure (Knuth, 1969) given below.

- 1) Generate two independent random variables, U_1 and U_2 , uniformly distributed between 0 and 1. Set $V_1 \leftarrow 2U_1-1$ and $V_2 \leftarrow 2U_2-1$. Then, V_1 and V_2 are uniformly distributed between -1 and +1.
- 2) Set $s \leftarrow v_1^2 + v_2^2$
- 3) If $S \ge l$, return to step 1.
- 4) Set X_1 and X_2 according to the following two equations:

$$X_1 = V_1 \sqrt{(-2 \ln S)/S}$$
 (4)

$$X_2 = V_2 \sqrt{(-2 \ln S)/S}$$
 (5)

According to Knuth, the polar method is easy to computerize and has essentially perfect accuracy.

Inverse Normal Function Method (INFM)

International Mathematical Statistics Library (IMSL) has a normal or Gaussian random deviate generator which interprets the random numbers distributed as U(0, 1) to be cummulative probabilities and computes the corresponding normal deviates through an inverse normal function subroutine. The subroutine computes X. so that:

$$U_{i} = Gauss (X_{i})$$

$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{X_{i}} \exp \left(-\frac{t^{2}}{2}\right) dt$$
(6)

Central Limit Theorem Method (CLTM)

The normally distributed random numbers can also be generated by the application of the central limit theorem. It states that the sum of a large number of components tends to the normal distribution as the number of components (regardless of their initial distribution) increases without limit (Ang and Tang, 1975). Therefore, the sum of a fixed number of uniform deviates on the interval (0, 1) should be distributed as gaussian. According to Cramer (1946), the mean and standard deviation of $f_n = \sum_{i=1}^n U_i$ are n/2 and $\sqrt{n/12}$, and f approaches the normal deviate rapidly as n increases. To generate standard normal deviates distributed as N (0, 1), n must be 12 for unit variance, and then $f_n = \sum_{i=1}^{12} U_i - 6$ for zero mean. Then, the deviate f_n constitutes normal deviate X.

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Evaluation of Random Number Generators

Suitability of the random number generators was evaluated by two methods: the consistency of the statistics derived from 10 samples of generated normal random deviates of size 1,000 to 50,000 and the consistency of the statistics derived from 20 to 1000 samples of size 15 to 100.

Random sampling distribution theory aids in finding distribution parameters of the 3 statistics (mean, variance or standard deviation, and skew – their population values are 0, 1, and 0) being used in comparative evaluation of the 4 algorithms. Denoting mean, standard deviation and skew by $\overline{\mathbf{X}}$, s, and g, the expected value and variance of the 3 statistics are given by the following equations (Cramer, 1946):

$E(\bar{X}) = 0$	(7)
$Var(\bar{X}) = \frac{1}{n}$	(8)

$$E(s) = 1$$
 (9)

$$Var(s) = \frac{1}{2n}$$
(10)

$$E(s^2) = 1$$
 (11)

$$Var(s^2) = \frac{2}{n-1} \text{ or } \frac{2}{n}$$
 as n becomes large (12)

$$E(g) = 0$$
 (13)

$$Var(g) = \frac{6n (n-1)}{(n-2)(n+1)(n+3)} \text{ or } \frac{6}{n} \text{ as } n \text{ becomes large}$$
(14)

A. Consistency of Statistics with Sample Size 1000 to 50,000

The intent was to investigate the variation of the mean and standard deviation of some statistics from the respective population values with respect to the length of the generated sequences from each of the 4 algorithms. The procedure, applied to each algorithm, can be considered in 4 steps. 1) Generate a sequence of 50,000 deviates.

i=1

- 2) Compute statistics: mean, standard deviation, and skewness for each of the 12 sample sizes of 1,000, 2,000,..., and 50,000 deviates from the beginning of the sequence.
- Repeat steps 1 and 2 10 times, giving 10 values of the 3 statistics for each of the 12 sample sizes.
- Compute the mean and standard deviation from the 10 values of each statistic for each of the 12 sample sizes.

$$AV(\text{statistic}) = \sum_{i=1}^{10} (\text{statistic})_{i}/10$$
(15)

$$i=1$$

$$STD(\text{statistic}) = \left[\sum_{i=1}^{10} ((\text{statistic})_{i}-AV(\text{statistic}))^{2}/9\right]^{0.5}$$
(16)

The AV in equation 15 can be compared with expected values from equations 7, 9, 11, and 13 which are 0, 1, 1, and 0. The STD in equation 16 corresponds to \sqrt{Var} . The STD values for the 3 statistics and 12 sample sizes are given below.

<u>Sample size, n</u>	$\sqrt{\frac{1}{n}}_{STD}(\bar{X})$	$\sqrt{\frac{1}{2n}}$ STD(s)	$\sqrt{\frac{6}{n}}$ STD(g)
1,000	0.03162	0.02236	0.07745
2,000	0.02236	0.01581	0.05477
3,000	0.01826	0.01291	0.04472
5,000	0.01414	0.01000	0.03464
7,000	0.01195	0.00845	0.02928
10,000	0.01000	0.00707	0.02449
15,000	0.00817	0.00577	0.02000
20,000	0.00707	0.00500	0.01732
25,000	0.00632	0.00447	0.01549
30,000	0.00517	0.00408	0.01414
40,000	0.00500	0.00354	0.01224
50,000	0.00447	0.00316	0.01095

Evaluation of Statistics

Mean: The expected value of the mean of the random deviates, N (0, 1), is zero according to equation 7. The values of AV of the mean from the 4 algorithms are plotted with respect to sample size in figure 3a. It is evident that PLRM and CLTM yield AV vs n curves that are closer to zero than the other two. The values of STD of the mean are graphed in figure 3b together with the curve corresponding to equation 8, i.e., $STD(\vec{X}) = \sqrt{1/n}$. The curves from PLRM and INFM lie below the equation 10 curve, practically for the whole range of n.

<u>Standard Deviation</u>: The expected value of the standard deviation, s, of the random deviates, N(0, 1), is 1 according to equation 9. The values of AV of standard deviation from the 4 algorithms are plotted with respect to sample size in figure 4a. The curves show that PLRM is the best, closely followed by CLTM and BAMM. However, the STD curves together with the $\sqrt{\frac{1}{2n}}$ curve (figure 4b) show that INFM is the best, PLRM and CLTM are equally good, and BAMM is the worst. The overall rating, considering both AV and STD, in the decreasing order of preference is PLRM, CLTM, INFM, and BAMM.

<u>Skewness</u>: The expected value of the skew for deviates, N (0, 1), is zero according to equation 13. The values of AV of the skew from the 4 algorithms are plotted with respect to sample size in figure 5a. It is evident that CLTM and BAMM are better than PLRM which is better than INFM. The comparison of STD curves with $\sqrt{6/n}$ curve (figure 5b) shows that all algorithms are similar for n larger than 10,000.

B. Consistency of Statistics with Sample Size 15 to 100

The aim was to analyze the variation in the mean and standard deviation of $\bar{x},\ s^2,$ and g for small sample sizes but with the number of samples

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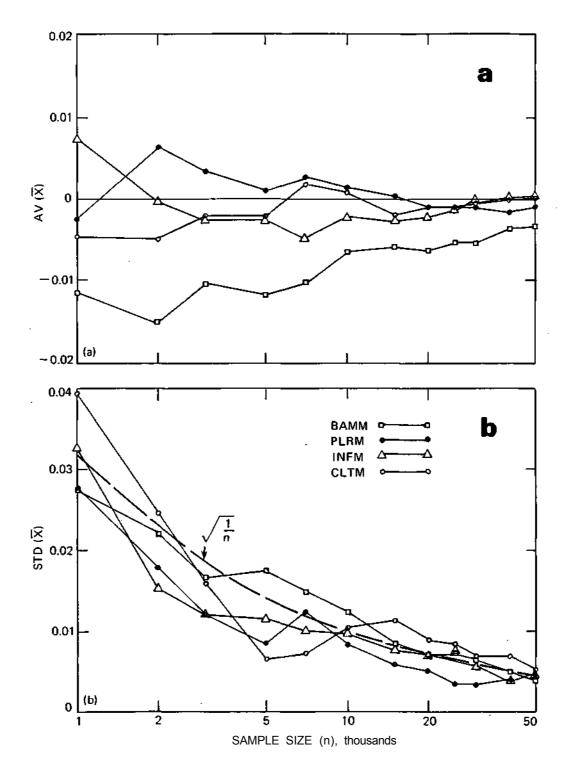


Figure 3. AV(X) and STD(X) versus sample size

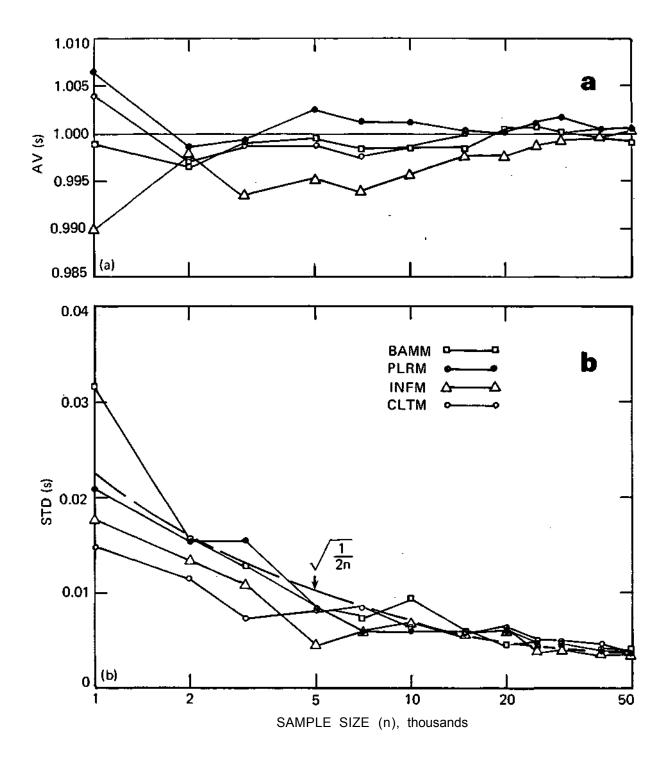


Figure 4. AV(s) and STD(s) versus sample size

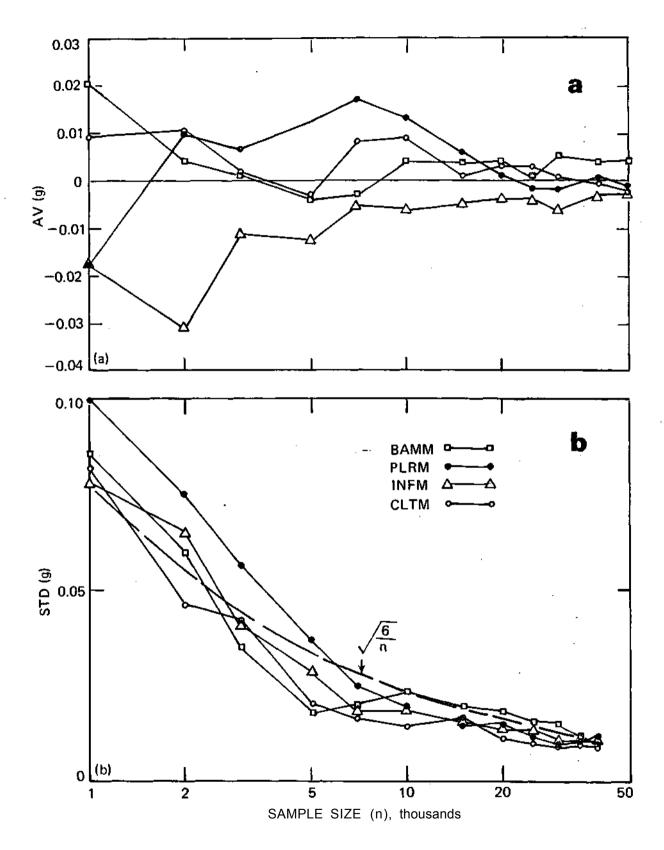


Figure 5. AV(g) and STD(g) versus sample size

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varying from 20 to 1000. The procedure applied to each of the 4 algorithms can be considered in 4 steps.

- 1) Generate a sequence of 50,000 to 75,000 deviates.
- 2) Dissect the sequence into sizes of 15, 25, 40, 50, 75, and 100 resulting in 20, 50, 100, 300, 500, and 1000 samples of all sizes with some exceptions.
- 3) Compute 3 statistics, mean $\bar{\mathbf{X}}$, variance \mathbf{s}^2 , and skewness g for each sample.
- Compute the mean and standard deviation or variance of each of the 3 statistics for each of the 6 number of samples of size 15, 25, 40, 50, 75, and 100.

The values of AV, STD, and Var of the three statistics for different sample sizes and number of samples were calculated from the 4 algorithms. The expected values and standard deviations or variances of the 3 statistics were also computed from equations 7, 8, and 11 to 14.

Evaluation of Statistics

The evaluation of AV and STD or VAR of mean, variance, and skew for 20, 50, 100, 300, 500, and 1000 samples of sizes 15, 25, 40, 50, 75, and 100 is explained by the example of sample size 15 in Table 1. The best rating is 4 assigned to AV, STD, OR VAR from samples closest to the E, \sqrt{Var} or Var from equation 7, 8 and 11 to 14 for the value of n under consideration. The ratings for 6 values of N (where N is the number of samples of size n) are added to give the total. It is evident from Table 1 that BAMM consistently underestimates statistics AV(X), AV(s²), and AV(g). The combined overall ratings (sum of the 6 totals) for the BAMM, PLRM, INFM, and CLTM are 63, 109, 95, and 93, respectively. The BAMM algorithm does not perform as well as others.

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N	BAMM		PLRM		INFM		CLTM		E,VAR, or √VAR
$AV(\widetilde{X})$		*		*		*		*	
20	-0.0290	2	-0.0257	3	0.0681	1	-0.0108	4	0.0000
50	-0.0128	2	0.0158	1	-0.0008	4	-0.0058	3	0.0000
100	-0.0371	1	-0.0112	3	0.0032	4	-0.0122	2	0.0000
300	-0.0150	l	0.0117	3	-0.0069	4	-0.0149	2	0.0000
500	-0.0176	I	-0.0006	4	0.0022	2	0.0014	3	0.0000
1000	-0.0059	1	-0.0013	_2	-0.0003	4	0.0011	3	0.0000
		8		16		19		17	
$\underline{STD}(\overline{X})$									
20	0.2945	2	0.2821	3	0.2142	1	0.2531	4	0.2582
50	0.2705	4	0.2939	2	0.2312	3	0.2092	Ł	0.2582
100	0.2772	2	0.2809	1	0.2545	4	0.2486	3	0.2582
300	0.2711	1	0.2690	2	0.2515	3	0.2517	4	0.2582
50 0	0.2654	3	0.2658	2	0.2477	1	0.2625	4	0.2582
1000	0.2680	1	0.2640	3	0.2501	2	0.2534	4	0.2582
		13		$\frac{3}{13}$		14		20	
AV(s ²)									
20	0.9159	2	1.0017	4	1.1513	1	1.0162	3	1.0000
50	0.9133	3	0.9841	4	1.1313	1	0.9053	2	1.0000
100	0.9221	1	1.0090	4	1.0463	3	0.9304	2	1.0000
300	0.9693	2	1.0324	1	0.9890	3	1.0067	4	1.0000
500	0.9694	. 1	1.0019	3	1.0014	4	0.9903	2	1.0000
1000	0.9891	3	0.9937	4	1.0122	2	0.9783	1	1.0000
1000	0.0001	12		20	1.0122	14	0.9705	14	1.0000
$\frac{VAR(s^2)}{2}$									
$\frac{VAR(S)}{20}$	0.0831	1	0.1468	4	0.1696	2	0.1312	3	0.1429
20 50	0.0331		0.1400	3	0.1843	2 2			0.1429
100		1		4		23	0.1163	4 2	
	0.1061	1	0.1163		0.1714		0.1117		0.1429
300	0.1278	1	0.1368	3	0.1411	4	0.1330	2 3	0.1429
500	0.1210	1	0.1369	2	0.1423	4	0.1375		0.1429
1000	0.1326	$\frac{1}{6}$	0.1385	$\frac{3}{19}$	0.1422	$\frac{4}{19}$	0.1347	$\frac{2}{16}$	0.1429
AV(a)		v		• /		• /		10	
<u>AV(g)</u> 20	-0.1248	2	0.0149	4	-0.1695	1	-0.0485	3	0.0000
20 50	-0.1248 -0.2157		-0.1029	4	-0.1695	4	-0.0485	3	0.0000
100	-0.0910	1	-0.1029	4	0.0109	4 3	0.1018	2	0.0000
	-0.0910	1	0.0050	-	-0.0171	2	0.0784		0.0000
300 500	-0.0213	I	0.0030	4 2	-0.0001	4		3 3	0.0000
1000		1			-0.0055		0.0125		0.0000
1000	-0.0163	$\frac{1}{7}$	0.0159	$\frac{2}{18}$	~0.0033	$\frac{4}{18}$	0.0126	<u>3</u> 17	0.0000
STD(g)		•		-				-	
$\frac{310(g)}{20}$.	0.5534	3	0.5612	4	0.5304	I	0.6147	2	0.5801
50	0.6015	3	0.5608	4	0.6580	2	0.7215	ĩ	0.5801
100	0.6226	3	0.5938	4	0.6327	2	0.6729	1	0.5801
300	0.5929	1	0.5770	4	0.5733	2	0.5743	3	0.5801
500	0.5849	4	0.5872	3	0.5913	2	0.5547	1	0.5801
1000	0.5880	3	0.5791	4	0.5901	2	0.5531	1	0.5801
1000	0.000	17	0.3771	23	0.9701	$\frac{2}{11}$	1,000		0.5001

Table 1. Evaluation of AV and STD or VAR of \bar{X} , s^2 , and g for Sample Size 15

1723119*refers to the rating, 4 is the highest and 1 is the lowest

N denotes the number of samples of size n which is 15 in this table

The information contained in Table 1 for sample size 15 is given for all the sample sizes 15, 25, 40, 50, 75 and 100 for the evaluation of AV, STD or VAR of the mean, variance, and skew in Table 2. The overall ratings for the 4 algorithms are:

	BAMM	PLRM	INFM	CLTM
$AV(\overline{X})$	49	105	96	100
STD(X)	84	64	83	89
AV(s ²)	71	113	92	74
VAR(s ²)	70	114	81	85
AV(g)	59	100	85	76
STD(g)	82	_90	97	51
	415	586	534	475

Thus, the PLRM algorithm seems to be the best in generating normal random numbers, with statistical attributes closely resembling samples drawn from a population distributed as N (0,1). This algorithm was used in developing departure distribution tables for detection of any outliers and/or inliers at the low and/or high end of the observed flood series.

Determination of Departure Distributions

Departure is defined here as the standard normal deviate corresponding to the plotting position of the high or low point of the series under consideration, minus the sample standard deviate for that point. The magnitude of these departures for outliers and inliers at the two extreme ends of the series needs to be determined at various probability levels. The theoretical departure distribution depends on m, n, and α in the general plotting position formula:

$$p = \frac{m - \alpha}{n + 1 - 2\alpha} \tag{17}$$

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n 15	BAMM	PLRM						
15			INFM	CLTM	BAMM	PLRM	INFM	CLTM
15	$\underline{AV}(\overline{X})$				$\overline{\text{STD}(\bar{X})}$			
	8	16	19	17	13	13	14	20
25	11	16	16	17	16	10	15	19
40	9	18	15	18	12	10	14	14
50	7	18	17	18	15	7	18	10
75	8	19	17	16	15	10	10	15
00	<u>6</u> 49	<u>18</u> 105	<u>12</u> 96	$\frac{14}{100}$	<u>13</u> 84	<u>14</u> 64	<u>12</u> 83	<u>11</u> 89
	<u>AV(s²)</u>				VAR(s ²)	_		
15	12	20	14	14	6	19	19	16
25	11	22	16	11	9 ·	21	13	17
40	13	20	15	12	16	17	12	15
50	13	19	16	12	12	21	12	15
75	13	18	15	14	13	18	15	14
100	<u>9</u> 71	$\frac{14}{113}$	<u>16</u> 92	$\frac{11}{74}$	<u>14</u> 70	$\frac{18}{114}$	$\frac{10}{81}$	<u>8</u> 85
	<u>AV(g)</u>				<u>STD(g)</u>			
15	7	18	18	17	17	23	11	9
25	12	15	21	12	16	20	16	8
0	10	19	11	10	16	11	17	6
0	11	15	12	12	. 9	14	19	8
5	9	17	13	11	10	11	17	12
0	<u>10</u> 59	$\frac{16}{100}$	<u>10</u> 85	$\frac{14}{76}$	<u> 14</u> <u> 82</u>	<u>-11</u> 90	$\frac{17}{97}$	<u>8</u> 51

Table 2. Evaluation of AV, STD or VAR of X, s , and g for Sample Sizes 15 to 100 $\,$

Note: 1000 samples of certain sample sizes were not generated or processed.

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in which a is a shift parameter, m is the rank, and n is the sample size. Such distributions have been derived in this study by generating large samples of departures for different m, n, and a>ranking them in an ascending order of magnitude, and determining the magnitude of departure at various probability levels and rank of outlier or inlier.

Generation of Departures

The following procedure was used in generating departures:

- Generate 100,000 standard normal deviates with each of the four algorithms and dissect them into 1000 samples of 100 deviates each.
- 2) Pick one sample for each n (i.e., 10, 15, 20, 25, 30, 40, 50, 60, 75, and 100) starting from the beginning of each sample of size 100; this gives 1000 samples of 10 different sizes from 10 to 100.
- 3) Rank each of the 1000 samples of n size in an ascending order of magnitude and store 1000 values of the 5 lowest and 5 highest deviates in 10 series; each series corresponds to a high or low point. There are 10 series for each of the 10 sizes of n, and the size of each series is 1000.
- Normalize each series by subtracting the mean from each deviate and dividing by the standard deviation.

5) Compute departures from 1000 normalized deviates in a series: Departure, $\Delta_{k,i}$ = ith theoretical standard normal deviate -kth normalized deviate for high/low location i (18) in which ith theoretical standard normal deviate equals that which corresponds to the probability from equation 17 with m = i (i=1, 2, ..., 5 and n-4, n-3,..., n) and n = 10, 15,..., 100;

and for the kth normalized deviate k = 1, 2, ..., 1000; and with proper a values determined in steps 6 through 8.

- 6) In order to determine the appropriate value of α for different i and n values, generate 1000 departures by the step 5 with each of the following 6 values of $\alpha - 0.00$, 0.25, 0.32, 0.38, 0.43, and 0.50. Compute the means of 1000 departures for each of the 6 values of a.
- 7) Interpolate the α values which make the means zero. These values of a were practically the same for the generating algorithms PLRM, INFM, and BAMM (values from CLTM were consistently lower) for the first highest and first lowest (similarly for the second highest and second lowest, and so on) rank for a given value of n. The results at the end of this step are given below:

<u>n</u>	<u>a valu</u>	es for the h	ighest and	lowest rank 1	to 5
	1	2	3	4	5
10	0.425	0.474	0.493	0.506	
15	0.414	0.465	0.484	0.494	0.505
20	0.408	0.456	0.482	0.490	0.505
25	0.405	0.448	0.477	0.494	0.500
30	0.406	0.443	0.464	0.482	0.491
40	0.404	0.439	0.460	0.472	0.484
50	0.403	0.439	0.455	0.469	0.474
60	0.402	0.439	0.452	0.469	0.472
75	0.403	0.441	0.450	0.456	0.461
100	0.402	0.440	0.451	0.456	0.460

 Plot the a values versus n for the 5 ranks and draw smooth curves. The values from the smooth curves are as shown on the next page.

<u>n</u>	Smoothed a	values for	the highest	and lowest	rank
	1	2	3	4	5
10	0.425	0.474	0.492	0.506	0.511
15	0.414	0.464	0.485	0.498	0.506
20	0.408	0.455	0.478	0.491	0.501
25	0.406	0.448	0.472	0.486	0.496
30	0.404	0.443	0.467	0.481	0.491
40	0.403	0.440	0.459	0.473	0.482
50	0.403	0.440	0.454	0.467	0.475
60	0.403	0.440	0.451	0.462	0.469
75	0.403	0.440	0.450	0.458	0.463
100	0.403	0.440	0.450	0.456	0.460

Compute departures from 1000 normalized deviates for each of the 100 series (5 series for low and 5 series for high ranks for each of the 10 sample sizes) developed from the PLRM algorithm in step 4 and use the a values derived in step 8 to compute $\Delta_{k,i}$ in step 5, then go to step 10.

- 10) Rank the 1000 departures in each of the 100 series, and obtain values of departures corresponding to probability, or rank/1000, equal to 0.01, 0.02, 0.05, 0.10, 0.15, 0.20, 0.25, 0.30, 0.35, 0.40, 0.45, 0.50, 0.55, 0.60, 0.65, 0.70, 0.75, 0.80, 0.85, 0.90, 0.95, 0.98 and 0.99. These departures are defined as (Δ)_{1,m,n} where the subscript 1 denotes the number of probability levels 1 to 23; m refers to rank of low values 1 to 5 and high values 6 to 10, 1 is the lowest and 10 is the highest; and n denotes the sample size number 1 to 10, 1 for size 10 and 10 for size 100.
- 11) Generate 40 samples of departures $(\Delta)_{\underline{1},\underline{m},\underline{n}}$ from 40 generated sequences of 100,000 standard normal deviates each with the PLRM

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algorithm. The mean of each of the 2300 departures $(\underline{1} \times \underline{m} \times \underline{n} = 23 \times 10 \times 10 \text{ or } 2300)$ was obtained from the corresponding 40 values for each $(\Delta)_{\underline{1},\underline{m},\underline{n}}$. The mean departures are designated as $(\overline{\Delta})_{\underline{1},\underline{m},\underline{n}}$.

Development of Compact Departure Table

Barnett and Lewis (1978) give critical values for 1% and 5% tests of discordancy for a single outlier in a normal sample, using the deviation from the sample mean divided by the sample standard deviation as the test statistic. The corresponding test statistic is deviate corresponding to the plotting position of the higher outlier minus the departure $(\overline{\Delta})_{\underline{1},\underline{m},\underline{n}}$. The comparison of the test statistics is given below:

Sample size, n	p = 0.	.01	p = 0.	p = 0.05		
<u> </u>	Barnett & Lewis	This study	Barnett & Lewis	This study		
20	2.88	2.89	2.56	2.55		
30	3.10	3.11	2.74	2.74		
40	3.24	3.24	2.87	2.86		
50	3.34	3.35	2.96	2.96		
60	3.41	3.43	3.03	3.03		
100	3.60	3.59	3.21	3.21		

Test statistics developed in this study for p = 0.01 and p = 0.05 are practically the same as given by Barnett and Lewis. However, the test statistics for inliers are not available in the literature. The test statistics for outliers of rank 2 to 5 and at other than 0.01 and 0.05 values of p are also not available in the literature.

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The table of 2300 departures, $(\bar{\Delta})_{\underline{1},\underline{\mathbf{m}},\underline{\mathbf{n}}}$, was reduced to a compact table containing only 230 values. The following procedure was used in developing the compact table.

1. It was considered desirable to restrict the number, NO, of inliers/ outliers at both the low and high end of flow spectrum, in relation to the sample size.

Sample size, n	NO
40 - 100	5
30-39	4
25-29	3
20-24	2
15-19	1

2. For the 5th outlier/inlier, the mean departure for n in the range of 40 to 100 was obtained by calculating the mean of $(\vec{\Delta})_{\underline{1},\underline{\mathbf{m}},\underline{\mathbf{n}}}$ for $\underline{\mathbf{n}} = 6$, 7, 8, 9, and 10 (or for n = 40, 50, 60, 75, and 100). Similarly, for the 1st outlier/inlier the mean departure for n in the range of 15 to 100 was obtained with n = 2, 3, 4, 5, 6, 7, 8, 9, and 10 (or for n = 15, 20, 25, 30, 40, 50, 60, 75, and 100). This reduces the departure table to $(\overline{\Delta})_{\underline{1},\underline{\mathbf{m}}}$ in which 1 = 1, 2,..., 23 and m = 1, 2,..., 10. The resulting compact table of departures is shown in Table 3.

The standard deviation of departures for a particular range of n was calculated in a similar manner as the mean departure for that range in step 2. The standard deviations are given in Table 4. For p = 0.3 or 0.7, recommended for detection and modification of outliers/inliers in the later part of this report, the standard deviation varies from 0.0008 to 0.0061 and 0.0013 to 0.0050 for low and high outliers, respectively, and from

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TABLE 3. DEPARTURES AT DIFFERENT PROBABILITY LEVELS FOR LOW AND HIGH OUTLIERS

P	Low 1 15 - 100	Low 2 20-100	Low 3 25-100	Low 4 30-100	Low 5 40-100	High 5 40-100	High 4 30-100	High 3 25-100	High 2 20-100	High 1 15-100
.01	6893	4950	4122	3633	3270	3769	4289	5109	6541	-1.0538
.02	6297	4452	3703	3242	2923	3271	3706	4423	5614	8993
.05	5325	3694	3033	2641	2373	2563	2900	3414	4328	6835
.10	4410	2991	2425	2107	1880	1951	2205	2580	3225	5003
.15	3747	2495	2008	1734	1547	1551	1746	2030	2514	3826
.20	3184	2091	1670	1432	1270	1236	1387	1614	1966	2951
.25	2687	1734	1372	1169	1040	0970	1083	1257	1514	2215
.30	2210	1408	1103	0933	0819	0737	0818	0942	1120	1586
.35	1766	1097	0846	0710	0622	0524	0573	0655	0762	1024
.40	1319	0795	0598	0501	0432	0322	0348	0387	0425	0511
.45	0879	0493	0357	0290	0247	0128	0131	0131	0109	0030
.50	0435	0193	0118	0085	0059	.0063	.0081	.0111	.0196	.0428
.55	.0028	.0114	.0122	.0123	.0128	.0247	.0288	.0356	.0495	.0878
.60	.0515	.0428	.0375	.0339	.0318	.0432	.0497	.0599	.0790	.1321
.65	.1036	.0764	.0638	.0563	.0517	.0621	.0711	.0843	.1087	.1761
.70	.1606	.1123	.0922	.0807	.0732	.0817	.0934	.1096	.1400	.2212
.75	.2243	.1519	.1234	.1070	.0966	.1030	.1168	.1364	.1731	.2675
.80	.2969	.1974	.1586	.1369	.1228	.1265	.1426	.1663	.2089	.3166
.85	.3868	.2509	.2005	.1724	.1540	.1535	.1720	.1999	.2496	.3718
.90	.5030	.3213	.2541	.2172	.1933	.1865	.2088	.2411	.2989	.4376
.95	.6809	.4309	.3366	.2853	.2534	.2348	.2627	.3000	.3686	.5285
.98	.8906	.5566	.4318	.3646	.3215	.2887	.3204	.3655	.4412	.6220
.99	1.0292	.6433	.4979	.4182	.3682	.3233	.3583	.4073	.4882	.6787

Note: 15-100, ..., 40-100 denote the range of sample size, n

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TABLE 4. STANDARD DEVIATION OF DEPARTURES AT DIFFERENT PROBABILITY LEVELS

Ρ	Low 1 15-100	Low 2 20-100	Low 3 25-100	Low 4 30-100	Low 5 40-100	High 5 40-100	High 4 30-100	High 3 25-100	High 2 20-100	High 1 15-100
.01	.0062	.0041	.0112	.0142	.0115	.0071	.0134	.0193	.0296	.0588
.02	.0076	.0034	.0082	.0101	.0088	.0056	.0094	.0100	.0197	.0396
.05 .10	.0073 .0070	.0031 .0048	.0037 .0013	.0050 .0021	.0048 .0025	.0020 .0003	.0030 .0008	.0079 .0035	.0106	.0212 .0099
.15	.0073	.0052	.0019	.0010	.0013	.0016	.0009	.0012	.0026	.0067
.20	.0068	.0059	.0022	.0008	.0008	.0016	.0014	.0017	.0014	.0036
.25	.0064	.0055	.0025	.0011	.0009	.0018	.0020	.0024	.0008	.0019
.30	.0061	.0052	.0027	.0015	.0008	.0017	.0025	.0024	.0014	.0016
.35 .40	.0062 .0056	.0047 .0043	.0032 .0034	.0016 .0021	.0009 .0013	.0019 .0018	.0025	.0029 .0036	.0019 .0024	.0022 .0032
.45	.0054	.0045	.0033	.0023	.0017	.0017	.0026	.0039	.0027	.0034
.50	.0047	.0040	.0030	.0021	.0018	.0017	.0027	.0039	.0036	.0039
.55	.0043	.0040	.0029	.0023	.0017	.0018	.0026	.0040	.0042	.0041
.60	.0041	.0034	.0025	.0024	.0016	.0019	.0024	.0037	.0046	.0039
.65	.0025	.0027	.0020	.0020	.0020	.0015	.0021	.0036	.0048	.0037
.70	.0015	.0021	.0015	.0020	.0020	.0013	.0017	.0033	.0050	.0042
.75	.0030	.0016	.0015	.0020	.0023	.0011	.0017	.0029	.0047	.0047
.80	.0048	.0016	.0011	.0018	.0021	.0008	.0014	.0025	.0047	.0049
.85	.0067	.0035	.0017	.0018	.0018	.0017	.0011	.0021	.0040	.0058
.90	.0098	.0066	.0038	.0020	.0014	.0026	.0024	.0022	.0029	.0061
.95	.0198	.0118	.0068	.0039	.0028	.0043	.0053	.0050	.0023	.0058
.98	.0357	.0200	.0113	.0076	.0046	.0073	.0107	.0095	.0045	.0055
.99	.0493	.0265	.0180	.0116	.0076	.0116	.0136	.0118	.0068	.0053

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0.0015 to 0.0021 and 0.0014 to 0.0025 for low and high inliers, respectively. These small values of standard deviation justify the use of a compact table. However, the table of 2300 departures can be as easily used in the computer program, if so desired.

The distributions of departures for the lowest 5 events are graphed in figure 6 and for the highest 5 events in figure 7.

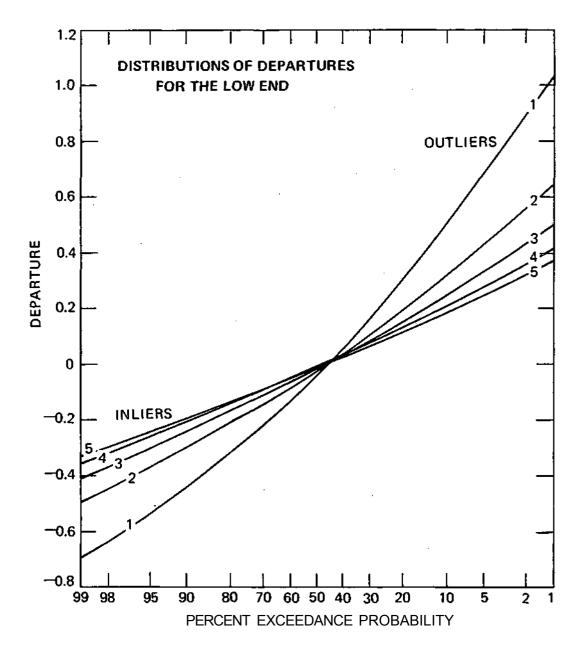


Figure 6. Distributions of departures for the low end

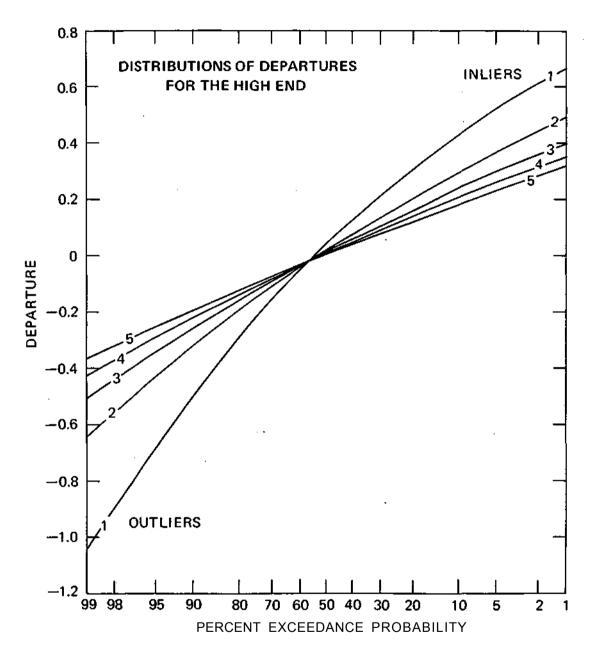


Figure 7. Distributions of departures for the high end

METHODS OF NORMALIZING DATA

The tests for determining outliers/inliers can be easily developed and applied for normally distributed samples because of the minimum number of distribution parameters, i.e., the mean and the standard deviation. Three methods of transforming an observed flood series to a series distributed as N (μ , σ^2), where u is the mean and σ^2 is the variance, were tested extensively on flood series observed at 28 gaging stations in Illinois. The methods are: power transformation, Wilson-Hilferty transformation, and 3-parameter lognormal transformation.

Power Transformation

Box and Cox (1964) suggested a transformation for normalizing a data set:

$$y_{i} = \frac{Q_{i}^{\lambda} - 1}{\lambda}, \quad \lambda \neq 0$$
 (1)

and

$$y_i = \log Q_i, \ \lambda = 0 \tag{2}$$

in which Q_{i} is the annual flood from a sample of size n, λ is a constant of transformation, and i = 1, 2, ..., n. It is a general power transformation and the logarithmic, reciprocal and square-root transformations can be considered as its special cases. The constant λ can be obtained with one of the following three criteria:

1. Maximum log-likelihood (ML) estimator of λ , when $\lambda \neq 0$, can be obtained from (Singh, 1980)

$$L_{\max}(\lambda) = -1/2 \ n \ \log \ \hat{\sigma}_y^2(\lambda) + \log \ J(\lambda;Q)$$
(3)

and $\log J(\lambda;Q) = (\lambda - 1) \sum_{\substack{\Sigma \\ i=1}}^{n} \log Q_i$ (4)

A plot of L_{max} (λ) versus λ can indicate the ML estimate of λ . A computer algorithm was developed for determining $L_{max}(\lambda)$.

2. Zero coefficient of skew criterion can be met by computing the skew g for y series with different values of λ ,

$$g = n \sum_{i=1}^{n} (y_i - \bar{y})^3 / [(n - 1)(n - 2) s_y^3]$$
 (5)

in which $\mathbf{\bar{y}}$ and s_y are the mean and standard deviation of the y series. Value of λ which makes g = 0 can be interpolated from the λ values giving a little positive and a little negative g. A computer program for calculating λ which yields g equal to zero was added to the ML algorithm.

3. Minimization of |g| + |5th| criterion is based on the premise that 3rd and higher order odd moments are zero in the case of a theoretical normal distribution. A computer program was added to the ML algorithm for calcualting λ value which minimized the sum of the absolute values of the skew and the 5th,

$$5th = n^{3} \sum_{i=1}^{n} (y_{i} - \bar{y})^{5} / [(n-1)(n-2)(n-3)(n-4) s_{y}^{5}]$$
(6)

Wilson-Hilferty Transformation

A standard deviate, x, can be calculated from Q, $\bar{\mathbf{Q}}$, and s_{Q} :

$$x_{i} = \frac{\log Q_{i} - (\overline{\log Q})}{s (\log Q \text{ series})}$$
(7)

If the underlying distribution is log-Pearson type III, or LP3, x is the gamma standard deviate that can be converted to the normal standard deviate by the

Wilson and Hilferty (1931) transformation

$$y_{i} = \frac{6}{g} \left[\left(\frac{gx_{i}}{2} + 1 \right)^{1/3} + \frac{g^{2}}{36} - 1 \right]$$
(8)

A computer program was developed for converting the Q series to x series, and for calculating the skew of the x series. Two subprograms were added to obtain values of g (with the first estimate equal to sample g for the x series) so that the y series has zero skew and to obtain a value of g that minimized the sum of the absolute values of g and 5th of the y series. These subprograms used a reiterative procedure to obtain satisfactory values of g to meet zero skew and min [|g| + |5th|] criteria.

Three-Parameter Lognormal Transformation

The following transformation was considered for normalizing the data,

$$y_{i} = \log (Q_{i} + a)$$
 (9)

in which a is a constant, positive, negative or zero. By a fast-converging reiterative process, the value of a was determined for the following three criteria:

- 1. skew g = 0
- 2. minimize [|g| + |5th|]
- 3. kurtosis = 3.0

A computer program was developed for calculating the values of a to meet the above criteria.

Test Data and Results

Annual flood series at 28 gaging stations from a previous report (Singh, 1980) were used in testing the suitability of the three transformations.

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These stations were selected and arranged in three categories:

- 14 gaging stations with flood series having no significant high or low outliers/inliers.
- II. 7 gaging stations with flood series having outliers/inliers at the high end.
- III. 7 gaging stations with flood series having outliers/inliers at the low end.

The 28 stations are listed in Table 5 together with observed high and low floods and their modified values as determined in a previous study (Singh, 1980). Category I flood series has no significant high or low outliers/ inliers but one outlier/inlier at either low or high end was considered for checking any effect of minor modification in values of these outliers/inliers. For both categories II and III, one and two outliers/inliers were considered separately.

Power Transformation Results

The results are presented in Table 6. Criteria A, B, and C denote g = 0, ML estimate, and min [|g| + |5th|], respectively. The TS1 and TS2 are test statistics, given by

$$TSL = \sum_{i=1}^{n} (\Delta Z_i)^2 \text{ with } \alpha = 0$$
(10)

$$TS2 = \prod_{i=1}^{n} (\Delta Z_i)^2 \text{ with } \alpha = 0.38$$
(11)

$$\Delta Z_{i} = (Z_{0})_{i} - (Z_{c})_{i}$$
(12)

 $(Z_{c})_{i} = (y_{i} - \bar{y}) / y_{s}$ (13)

formations	3	of	Evaluation	in	Used	Basins	5.	TABLE
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	No.	Stream and gaging station	USGS No.	n	<u>Obs</u>	erved flood i			ed flood in	
					L1	L2 H1	H2	L1	L2 H1	H2
		A. Well Behaved Flood Series				·				·
	1	Sangamon River at Mahomet	05 571000	29	1020	14600		1060	13682	
	2	Sangamon River at Monticello	05 572000	68	704	19000		1110	19567	
	3	Salt Creek near Greenview	05 582000	35	3440	41200		3224	46989	
	4	Pecatonica River at Freeport	05 435500	63	1910	18400		1704	20082	
	5	Pecatonica River at Shirland	05 437000	32	3490	16600		2946	18282	
	6	Rock River at Rockton	05 437500	37	6340	30000		5402	29249	
	7	Kishwaukee River near Perryville	05 440000	37	2020	16400		1579	17754	
	8	Green River near Geneseo	05 447500	40	1340	12100		1323	11533	
	9	Kishwaukee River at Belvidere	05 438500	37	935	10300		716	11994	
	10	S. Br. Kishwaukee R. near Fairdale	05 439500	37	1010	8460		819	7559	
	11	Elkhorn Creek near Penrose	05 444000	37	530	6770		563	6872	
	12	Rock Creek near Morrison	05 445500	32	765	5770		720	5429	
-35-	13	Green River at Araboy	05 447000	37	480	6120		340	5327	
	14	Mill Creek at Milan	05 448000	37	450	9300		399	11253	
		B. Flood Series with High Outliers/	Inliers							
	1	S.F. Sangamon River near Nokomis	05 574000	26		8600			9340	
						8600	6000		6439	4440
	2	S.F. Sangamon River near Kincaid	05 575500	33		21500			17082	
						21500	13700		20079	15808
	3	Sangamon River at Riverton	05 576500	64		68700			42826	
						68700	41000		36170	34880
	4	Salt Creek near Rowell	05 578500	34		24500			16258	
						24500	12400		19949	14911
	5	Kickapoo Creek near Lincoln	05 580500	32		14800			17558	-
	-			-		14800	13800		12230	10066
	6	Sangamon River at Oakford	05 583000	59		123000			81230	
	2					123000	46300		58656	52683
	7	Skillet Fork near Wayne City	03 380500	48		51000			26591	
	,	shiring fork hear harme erey		10		51000	22800		28673	24441
						01000	22000		20070	

TABLE 5. (concluded)	TABLE	5.	(concluded)	
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No.	Stream and gaging station	USGS	s No.	n	Obs	served	flood	in cfs	Modif	ied flo	od in	cfs
					L1	L2	H1	Н2	L1	L2	Н1	H2
	C. Flood Series with Low Outliers/I	nlier	S	<u> </u>					i			
1	Flat Branch near Taylorville	05	574500	27	457				452			
					457	660			1088	1360		
2	S.F. Sangamon River near Rochester	05	576000	27	971				910			
					971	1230			1211	1568		
3	Spring Creek at Springfield	05	577500	29	217				154			
					217	225			377	479		
4	Lake Fork near Cornland	05	579500	29	152				432			
					152	548			394	507		
5	Leaf River at Leaf River	05	441000	37	233				309			
					233	459			215	335		
6	Little Wabash River below Clay City	03	379500	62	1440				2319			
					1440	2920			1899	2439		
7	Little Wabash River at Carmi	03	381500	37	3320				3516			
					3320	4180			5953	6567		

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TABLE 6. Evaluation of Normalization of a Flood Series by Power Transformation

			W	ith no m	odifica	ation			With H1	lmodif	i e d			With L	1 modif	i e d	
# US	GS No.	*	λ	g	5 t h	TS1	TS2	λ	g	5 t h	TS1	TS2	λ	g	5 t h	TS1	TS2
١.	Flood se	eries	with	no sign	ificant	high	or low	outlier	s/inlie	rs							
1	05571000	A B C	0.075 0.066 0.095	-0.011	-0.204 -0.295 -0.002	0.445	0.320	0.103 0.090 0.134	-0.016	-0.282 -0.401 0.001	0.459	0.353	0.053	-0.001 -0.010 0.022	-0.244	0.447	0.330
2	05572000	-	0.180	-0.001 -0.015	-0.330	0.517	0.420 0.421	0.174	0.000	-0.306	0.511 0.513	0.401 0.402	0.105 0.095	0.000 -0.013 -0.013	0.098	0.464 0.464 0.464	0.472 0.473
3	05582000	В		0.000 -0.009 -0.097	0.653		0.834		0.000 -0.001 -0.110	0.853	0.889 0.889 0.907	0.807		0.000 -0.011 -0.091		0.848 0.847 0.858	0.778
4	05435500	A B C	0.050 0.044 0.053	-0.007	- 0 . 0 2 2 - 0 . 0 6 1 - 0 . 0 0 2	0.666	0.763		0.000 -0.003 -0.009	0.035	0.603 0.603 0.604	0.665	0.075 0.067 0.089	0.000 -0.010 0.016			0.633
- 5	05437000	A B C	0.532 0.414 0.474			0.916	1.018		0.000 -0.061 -0.062		0.790	0.838		0.000 -0.085 -0.025	-0.337	0.740 0.750 0.740	0.801
6	05437500	A B C		0.000 -0.124 -0.060	0.198	0.927 0.937 0.958			0.000 -0.028 -0.054	0.138		1.055		0.000 -0.033 -0.038	0.033	0.717 0.725 0.726	0.740
7	05440000	A B C		0.000 -0.123 -0.011	0.557	1.368	1.546 1.572 1.544		0.000 -0.112 -0.029		1.229	1.389		0.000 -0.132 -0.003	-0.675	1.193 1.199 1.192	1.355
8	05447500	A B C	0.743 0.731 0.709	0.000 -0.016 -0.045	0.267	0.629	0.367 0.369 0.377		0.000 -0.030 -0.026		0.555	0.329		0.000 -0.015 -0.043	0.267	0.628 0.630 0.638	0.369
9	05438500	A B C	0.332 0.248 0.318			1.157	1.320	0.206	0.000 -0.059 -0.037	-0.125	0.998	1.101	0.371 0.284 0.371	0.000 -0.084 0.000		0.973 0.974 0.973	
10	05439500	A B C		0.000 -0.171 -0.119	-0.343	1.137			0.000 -0.212 -0.061	-0.815	1.264	1.381		0.000 -0.177 -0.115	-0.418	0.987 1.044 1.004	1.066

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TABLE 6 . continued

	With no mc	dification		With H1 modi	fied	With L1	modi-vied	
# USGS No. *	λg	5th TS1	τς2 λ	g 5th	TS1 Ts2	λg	5th TS1	TS2
I. Flood series w	with no signif:	icant high or	low outliers	s/inliers cor	tinued			
11 05444000 A B C	0.842 0.000 0.742 -0.113 0.327 -0.017	-0.671 0.470	0.391 0.734	-0.112 -0.62	2 0.434 0.351 27 0.457 0.367 2 0.434 0.349	0.736 -0.114	0.130 0.455 -0.643 0.480 0.002 0.455	0.405
12 05445500 A B C	0.142 -0.001 0.144 0.002 0.107 -0.053	0.525 0.846	0.564 0.191	0.001 0.35	7 0.784 0.523 0 0.784 0.523 5 0.786 0.525		0.416 0.875 0.461 0.875 0.007 0.877	0.577
13 05447000 A B C	0.872 0.000 0.803 -0.099 0.810 -0.089	-0.082 0.953	0.813 0.897	-0.155 -1.02	0 0.746 0.701 24 0.836 0.790 0 0.746 0.701		0.703 0.847 -0.167 0.892 0.007 0.879	0.724
14 05148000 A B C	3.272 0.000 0.229 -0.060 0.235 -0.052	-0.043 0.872	0.912 0.190	-0.046 0.16	8 0.744 0.718 3 0.752 0.728 03 0.762 0.738		0.268 0.822 -0.115 0.828 0.001 0.824	0.848
With II. Flood serie	no modificati es with high o			th H1 modifie	d	With H1 and	H2 modified	
	-0.493 0.130		0.366 -0.501	0.130 0.94	240.4360.33920.4590.353040.4360.339	-0.436 0.114	-0.090 0.463 0.713 0.476 0.002 0.462	0.395
2 05575500 A B C	0.058 0.000 0.055 -0.006 0.037 -0.039		0.286 0.118	-0.019 -0.13	70.4260.294L00.4270.295000.4270.294	0.057 -0.001 0.054 -0.006 0.047 -0.019	0.112 0.450	0.276
3 05576500 A B C		3.460 3.132 5.089 3.154 0.001 3.266	2.656 0.593	0.000 0.75	31.1580.88031.1580.880001.1970.916	0.691 -0.061	0.189 0.618 -0.303 0.643 -0.001 0.624	0.498
	-0.079 0.000 -0.063 0.020 -0.114 -0.043	0.309 1.069 0.452 1.072 0.005 1.070	1.099 0.025	-0.010 -0.06	071.2161.330511.2141.329021.2161.330	-0.047 -0.001 -0.035 0.013 -0.062 -0.017	0.192 1.019	1.077
	-0.095 0.000 -0.080 0.016 -0.164 -0.073		0.859 -0.135	0.025 0.81	40.8230.79740.8250.796010.8420.826	0.101 0.000 0.078 -0.019 0.051 -0.041	0.131 0.929	0.990

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TABLE 6. (concluded)

				W	ith no n	nodific	ation			With H1	modif	i e d		Wit	hH1 a	nd H2 r	modifie	d
	#	USGS No.	*	λ	9	5th	TSI	TS2	λ	9	5th	T\$1	TS2	λ	g .	5th	TSI	T\$2
	11	. Flood s	erie	s with	high o	utliers	/inlie	rs con	ıtinued									
	6	05583000	A B C		0.000 0.017 -0.290		2.555	2.349		0.000 -0.143 -0.011	-0.735	1.009 1.044 1.005	1.219	0.531	0.000 -0.109 -0.056	-0.352	0.767	0.831
	7	03380500	A B C	0.115	-0.001 0.004 -0.053	0.712	1.365	1.093 1.094 1.092	0.317 0.284 0.359	-0.049	-0.476 -0.869 -0.006	0.863	0.813		0.000 -0.039 0.050		0.816	0.726
				W	ith no m	nodific	ation			With L1	modif	ied		Wit	hL1 a	nd L2 r	nodifie	d
	11	I. Flood	seri	ies witl	n low o	utliers	s/inlie	ers										
	1	05574500	A B C	0.356	0.000 -0.023 -0.006		0.678			0.001 -0.024 -0.005	-0.170				- 0 . 0 0 1 - 0 . 0 0 4 - 0 . 0 4 8	0.328	0.571 0.571 0.580	0.461
	2	05576000	A B C	0.194	0.001 -0.032 -0.052	0.164	0.683	0.527 0.531 0.535		0.000 -0.033 -0.046	0.110	0.673 0.677 0.679	0.515	0.100	0.000 -0.019 -0.067	0.366	0.720 0.722 0.733	0.612
-39-	3	05577500	A B C		0.000 -0.021 0.087		0.837	0.723	0.195 0.184 0.243	-0.024	-0.851 -1.056 0.000	0.840	0.689	- 0 . 1 5 9 - 0 . 1 2 9 - 0 . 1 1 5	0.036	-0.109		
	4	05579500	A B C		-0.001 0.006 0.194	-1.893	1.072	0.851	0.063 0.058 0.134	-0.007	-0.790 -0.848 0.002	0.668	0.531		0.000 -0.011 1.113		0.708	
	5	05441000	A B C	0.439	0.000 -0.159 -0.027	-0.736	1.013	1.088 1.108 1.080		0.000 -0.155 -0.035	-0.639	1.042 1.086 1.038	1.208	0.456	0.001 -0.166 -0.024	-0.810	0.920 0.952 0.914	1.037
	6	03379500	A B C	0.154	0.001 -0.015 0.064		1.027	0.959		-0.001 -0.012 0.008		0.941	0.967		0.000 -0.018 0.033		0.935	
	7	03381500	A B C	0.026	-0.001 0.004 0.133	-1.401	1.722	1.441	0.002 0.002 0.059	0.001	- 1 . 3 2 2 - 1 . 3 2 2 - 0 . 0 0 7	1.709	1.388	- 0 . 5 3 7 - 0 . 5 0 7 - 0 . 5 0 0	0.036	-0.355 -0.066 0.001	0.518	

* Criterion used for derivingλ

$$(p_0)_i = \frac{i - \alpha}{n + 1 - 2\alpha}$$
(14)

and $(Z_o)_i$ is obtained from $(p_o)_i$ by $a \ p \rightarrow Z$ subroutine for normal distribution. <u>Category I</u>: The following inferences can be made from the results in Table 6 for flood series without significant high or low outliers/inliers.

1. TS2 is lower than TS1 for about 50% of the basins. A lower value of the test statistic shows an overall better fit.

2. For a basin, the TS1 or TS2 values for the three criteria A, B, and C are quite close to each other with (or without) modification of any outliers/ inliers.

3. Minimum values of the 5th are obtained with the criterion that minimizes [|g| + |5th|].

4. For a given basin, the values of λ for the three criteria are not much different from each other, but the ML estimate of λ is generally somewhat smaller than those with g = 0 and min [|g| + |5th|].

5. The values of λ for the three cases: with no modification of any outlier/inlier, with modification of highest inlier/outlier, and with modification of lowest outlier/inlier, are not much different from each other when the flood series are well behaved, i.e., they do not have significant high and low outliers/inliers.

<u>Category II</u>: The following inferences can be made from the results in Table 6 for flood series with high outliers/inliers.

1. TS2 is lower than TS1 for about two-thirds of the basins. A lower value of the test statistic shows an overall better fit. Thus, the use of $\alpha = 0.38$ seems better than $\alpha = 0.00$.

2. For a basin, the TS1 or TS2 values for the three criteria A, B, and C are rather close to each other with (or without) modification of

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outliers/inliers. However, for a given criterion, these values vary considerably from each other when obtained with and without modification of outliers/inliers for three basins, each with a rather severe high outlier; e.g., in the case of basin 3 and criterion A, TS2 decreases from 2.636 with no modification of outliers, to 0.880 with modification of H1, and to 0.478 with modification of H1 and H2. The TS1 or TS2 values, after modification of H1 and H2, lie in the general range of 0.4 to 1.0, the same as for category I.

3. Minimum values of the 5th are obtained with the criterion that minimizes [|g| + |5th|].

4. For a given basin, the values of X for the three criteria, after modification of outliers/inliers, are not much different from each other but the ML estimate of X is more often somewhat smaller than those with the other criteria.

5. The value of X changes with modification of any outlier/inlier and the magnitude of change depends on the severity of the outlier/inlier.

<u>Category III</u>: The following inferences can be made from the results in Table 6 for flood series with low outliers/inliers.

1. TS2 is generally less than TS1 with and without modification of any outliers/inliers. A lower value of the test statistic indicates an overall better fit. Thus, the use of a = 0.38 seems better than a = 0.00.

2. For a basin, the TS1 or TS2 values for the three criteria A, B, and C are rather close to each other with (or without) modification of outliers/ inliers. However, for a given criterion, these values vary considerably from each other when obtained with and without modification of outliers/inliers for three basins, each with 1 or 2 rather severe low outliers. The TS1

or TS2 values, after modification of L1 and L2, lie in the general range of 0.4 to 1.0, the same as for categories I and II.

3. Minimum values of the 5th are obtained with the criterion that minimizes [|g| + |5th|].

4. For a given basin, the values of λ for the three criteria, after modification of outliers/inliers, are not much different from each other.

5. The value of λ changes with modification of any outlier/inlier and the magnitude of change depends on the severity of the outlier/inlier.

Wilson-Hilferty Transformation Results

The results are presented in Table 7. Criteria A, B, and C denote transformation as expressed by equation 8 ($g_s = skew$ of x series in equation 7 and g = skew of y series in equation 8); iterative modification of g_s so that g becomes zero (g_s equals the value of g used in equation 8 so that skew of y series becomes zero); and iterative modification of g_s so that [|g| + |5th|] of y series becomes minimum, respectively. The 5th, TS1, and TS2 are the same as defined under power transformation or earlier.

<u>Category I.</u> The following inferences can be made from the results in Table 7 for flood series without significant high or low outliers/inliers.

1. TS1 is lower than TS2 for about 50% of the basins. A lower value of the test statistic shows an overall better fit.

2. The TS1 or TS2 values for the three criteria A, B, and C are quite close to each other with (or without) modification of any outliers/inliers in the case of 11 basins, but for 3 basins these values are considerably higher with A than with B or C.

TABLE 7. Evaluation of Normalization of a Flood Series by Wilson-Hilferty Transformation

				,	With no	modific	ation			With H	1 modi	fied			With I	L1 modi	ified	
	#	USGS No.	×	9 _s	9	5th	TSI	TS2	9 _s	g	5th	TSI	TS2	9 _s	8	5th	TS1	TS2
	I	. Flood	seri	es with	no sign	nificant	high	or low	outlie	rs/inlie	rs							
	1	05571000		- 0 . 0 9 8 - 0 . 1 4 9 - 0 . 1 9 0	0.000	- 0 . 9 6 8 - 0 . 2 0 7 0 . 0 0 1	0.446	0.319	- 0 . 1 2 9 - 0 . 2 0 5 - 0 . 2 6 5	0.000		0.460	0.351	- 0 . 0 7 7 - 0 . 1 2 0 - 0 . 1 5 5	0.000	-0.382 -0.171 0.001	0.447	0.329,
	2	05572000	В	- 0 . 3 0 8 - 0 . 3 6 5 - 0 . 4 1 5		- 0 . 7 8 1 - 0 . 3 6 4 - 0 . 0 0 1	0.535	0.415	- 0 . 3 0 0 - 0 . 3 5 3 - 0 . 3 9 8		-0.730 -0.339 0.000	0.530	0.398	- 0 . 1 4 3 - 0 . 2 0 5 - 0 . 1 8 6	0.000	0.097	0.463 0.461 0.459	0.473
	3	05582000	В	- 0 . 0 7 3 - 0 . 1 1 4 - 0 . 3 4 0	0.000	0.720	0.832 0.884 0.903	0.833	-0.010	-0.002 0.000 -0.111	0.864	0.889	0.807	- 0 . 0 9 7 - 0 . 1 4 8 - 0 . 0 1 0	0.000	0.688	0.844 0.849 0.857	0.777
	4	05435500	В	-0.086	-0.020 0.000 0.004	-0.002	0.665	0.761	- 0 . 0 3 1 - 0 . 0 4 6 - 0 . 0 3 3		0.054	0.603	0.664			- 0 . 2 8 0 - 0 . 0 9 7 0 . 0 0 1	0.576	0.629
- 4 3 -	5	05437000	В	-0.7i6	-0.190 0.000 -0.030	-0.208	0.91:1	0.957	-0.509	-0.157 0.000 -0.063	0.384	0.782	0.786	-0.420 -0.806 -0.770	-0.215 0.000 -0.023			0.731
	6	05437500	В		0.000	0.337	0.927		-0.283	- 0 . 0 7 2 0 . 0 0 0 - 0 . 0 5 4	0.292		1.033	-0.359	-0.091 0.000 -0.038	0.236		0.714
	7	05440000	В	-1.056	-0.276 0.000 -0.013	0.071	1.316	1.369	-0.963	-0.257 0.000 -0.036	0.215	1.165	1.193	-1.132	-0.302 0.000 -0.002	0.014	1.133	
	8	05447500	В	-0.939		0.591	0.985	0.673	-1.109 -1.006 -0.989	0.217 0.000 -0.029	0.287	1.338 0.886 0.851	0.597	- 1 . 0 9 1 - 0 . 9 4 4 - 0 . 9 1 3	0.373 0.000 -0.054	0.584	1.946 0.993 0.931	0.679
	9	05438500	В	-0.690		0.075	1.128	1.229	-0.560	-0.145 0.000 -0.040	0.241	0.973	1.031	- 0 . 3 7 0 - 0 . 7 8 4 - 0 . 7 8 2	- 0 . 2 0 2 0 . 0 0 0 - 0 . 0 0 1	0.004	0.978 0.940 0.940	0.969
	10	05439500	В	-1.246	-0.418 0.000 -1.154	1.495	1.198	1.046	-1.470	-0.479 0.001 -0.090	0.726	1.087	0.912	- 0 . 8 6 6 - 1 . 2 8 8 - 1 . 1 9 8	-0.001	1.549		0.926

TABLE 7. continued

			W	ith no n	nodifica	ation			With H1	modif	i e d			With L	l modif	ied	
#	USGS No.	*	g _s	g	5th	TSI	TS2	9 _s	9	5th	TSI	TS2	9 _s	9	5th	TSI	TS2
Ι.	Flood se	ries	with no	o signi	ficant	high o	r low	outlier	s/inlier:	s cor	ntinuec	ł					
11	05444000	A B C	- 1 . 1 0 6 - 1 . 2 9 6 - 1 . 2 9 0	0.000	-1.815 0.096 0.004		0.366		0.000	-1.696 0.171 -0.031	0.592	0.354	-1.279	-0.256 0.000 -0.015	0.128	0.486 0.596 0.574	0.373
12	05445500		- 0 . 2 1 0 - 0 . 1 9 1 - 0 . 1 4 4		0.712 0.508 -0.003		0.573	- 0 . 2 3 7 - 0 . 2 5 3 - 0 . 2 1 8		0.479 0.346 -0.001	0.801	0.537		0.038 0.000 -0.043	0.420	0.909 0.893 0.832	0.592
13	05447000	В	- 1 . 2 7 9 - 1 . 2 8 2 - 1 . 2 1 8	-0.001		1.632	1.416	-1.521	0.000	- 2 . 3 7 3 0 . 7 7 4 - 0 . 0 0 5	1.154	0.937	- 1 . 4 2 0 - 1 . 3 4 4 - 1 . 2 9 1	1.627 0.000 -0.780	2.325	10.43 1.682 1.420	1.435
14	05'!48000		-0.639	0.000		0.878	0.872	- 0 . 3 2 8 - 0 . 5 2 6 - 0 . 4 2 4	0.000		0.752	0.694	-0.675	-0.174 0.000 -0.041	0.255	0.837	0.808
			V	Vith no	modific	ation			With H	1 modi	fied		W	ith H1 a	and H2 r	modifie	d
11	. Flood	seri	es with	high o	utliers	/inlie	rs										
1	05574000	A B C	1.267	0.000	2.417 -0.059 -0.002	0.508	0.264	1.000 1.292 1.289		2.445 -0.040 -0.001	0.507	0.255	0.723 1.001 1.088	0.000	2.070 -0.097 -0.002		0.284
2	05575500	A B C	-0.134	-0.021 0.000 -0.039	0.344	0.476 0.478 0.477	0.289	-0.294	0.000	-0.429 0.035 -0.001	0.432	0.290	-0.130	- 0 . 0 2 2 0 . 0 0 0 - 0 . 0 2 0	0.163	0.452	0.276
3	05576500	A B C		3.179 0.000 -0.261		40.28 4.313 3.882	3.833	-1.065	2 . 4 5 3 - 0 . 0 0 1 - 0 . 0 9 4	55.64 1.193 0.003	2.345	1.997	- 1 . 5 2 7 - 1 . 2 8 0 - 1 . 2 7 0	1.948 0.000 -0.029	0.328	32.07 1.527 1.471	1.238
4	05578500	A B C	0.207	0.051 0.000 -0.045	0.315	1.076 1.068 1.080	1.090	-0.033 -0.089 -0.092	0.000	-0.125 -0.006 0.000		1.328	0.052 0.121 0.162	0.030 0.000 0.018	0.111	1.020 1.018 1.021	1.075
5	05580500	A B C	0.104 0.184 0.320	0.045 0.000 -0.075	0.489	0.853 0.852 0.889	0.857	0.192 0.311 0.468	0.072 0.000 -0.094	0.648	0.826 0.828 0.894	0.791	-0.184	-0.046 0.000 -0.041	0.247	0.931 0.928 0.929	0.983

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TABLE 7. (concluded)

				Wi	th no m	odifica	tion			With H1	modifi	ed		Wit	h Hl an	d H2 ma	odified	
	# USGS	No.	*	g _s	9	5th	TS1	TS2	9 _s	9	5th	TS1	TS2	9 _s	9	5th	TS1	TS2
	II.	Flood	ser	ies with	high	outlier	s/inli	ers co	ntinued									
	6 0	5583000		-0.520 -0.434 -0.220	0.000	4.078		2.873 2.523 2.399	-0.819 -1.275 -1.251	0.000		0.763	0.675		-0.293 0.000 -0.080	0.695		0.581
	7 0	3380500	В	-0.293 -0.261 -0.214	-0.001	0.698	1.439 1.405 1.379	1.260	-0.711 -0.787	0.064	-0.538 0.000	0.854 0.927	0.734	-0.633 -0.692	0.051	-0.459 0.001	0.829	0.683
	III.	Flood	ser	Wi ies with	th no m 1 low ou			rs		With L1	modif	ied		Wi	th L1 an	nd L2 m	odified	d
	1 0	5574500		-0.822 -0.796 -0.785	0.000	0.150	0.845 0.814 0.803	0.549	-0.829 -0.800 -0.789	0.000	0.146	0.817	0.551	-0.032 -0.052 0.025		0.360	0.572 0.571 0.581	0.460
-45-	2 0	5576000	A B C	-0.331 -0.454 -0.390	0.000	0.432		0.524	-0.367 -0.487 -0.432		0.392	0.694	0.510	-0.236	-0.057 0.000 -0.067	0.516	0.724 0.722 0.726	0.606
	3 0	5577500	A B C	-0.291 -0.377 -0.491	0.000		0.847	0.722	-0.458 -0.533 -0.647		-1.427 -0.872 0.002	0.869	0.696	0.187 0.363 0.265		0.276 -0.367 0.000		0.642
	4 0	5579500	A B C	-0.961 -0.779 -0.939		0.338 -2.193 -0.001		0.968	-0.096 -0.126 -0.269	-0.022 -0.001 0.106	-0.797	0.670	0.532	-0.182 -0.226 -0.372		-1.145 -0.876 0.002	0.713	0.561
	5 0	5441000	A B C	-0.746 -1.310 -1.292	0.000	0.118	0.904	0.775	-1.253	-0.363 0.000 -0.034	0.204	0.985	0.902	-1.369	-0.395 0.000 -0.015		0.836	0.681
	6 0	3379500	A B C	-0.305 -0.364 -0.448		-0.582		0.956	-0.129 -0.188 -0.202		-0.348 -0.066 0.000	0.939	0.961	-0.230 -0.303 -0.348	-0.055 0.000 0.035	-0.677 -0.258 0.001	0.936	0.899
	7 0	3381500	В	-0.055 -0.040 -0.137	0.000	-1.208 -1.444 0.002	1.774	1.442	-0.004 -0.003 -0.096	0.000	-1.314 -1.332 0.001	1.709	1.388	0.651 0.729 0.683	0.084 0.000 0.051	-0.448	0.562 0.612 0.576	0.391

* Criterion used for deriving g_s .

3. The values of the 5th with criterion A are generally much higher than with ML estimate of λ . Though the criterion C minimized values of the 5th, it decreases skew very significantly (-1.154 and -1.176) below zero for two basins.

4. The values of g for the three criteria are significantly different from each other for 4 out of the 14 basins.

5. The absolute values of g with Wilson-Hilferty transformation are considerably higher than with the ML estimate of λ . Thus, the power transformation brings a flood series closer to the normal distribution than the Wilson-Hilferty transformation.

<u>Category II.</u> The following inferences can be made from the results in Table 7 for flood series with high outliers/inliers.

1. TS2 is lower than TS1 for about two-thirds of the basins.

2. For USGS No. 05 576500 with a significant high outlier, the Wilson-Hilferty transformed series yields g = 3.179 compared to 0.107 with the power transformation and ML estimate of λ . The absolute value of g for other basins is also somewhat higher than with the power transformation.

3. Values of the 5th with Wilson-Hilferty transformation are higher than with the power transformation.

4. The values of g for a given criterion but considering no modification and modification of H1 or H1 and H2 differ significantly for two basins out of seven. With power transformation, there are no significant differences.

<u>Category III.</u> The following inferences can be made from the results in Table 7 for flood series with low outliers/inliers.

1. TS2 is generally less than TS1. A lower value of test statistic indicates a better overall fit. Thus, the use of a = 0.38 seems better than

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 $\alpha = 0.00.$

2. The absolute value of g for the y series with the Wilson-Hilferty transformation is generally higher than with the power transformation.

3. Minimum values of the 5th are obtained with the criterion that minimizes [|g| + | 5th|].

Three-Parameter Lognormal Transformation

The results are presented in Table 8. Criterion A corresponds to a value of a which reduces skew of y series (equation 9) to zero; criterion B denotes the value of a that minimizes [|g| + |5th|]; and criterion C refers to the value of a which makes kurtosis equal to 3.0. The 5th, TS1, and TS2 have been defined earlier. It is evident from Table 8 that no results are obtained for nine basins out of a total of 28 basins analyzed. This transformation is not suitable for converting a flood series to approximately a normal distribution.

Category I: Some inferences of interest are:

1. Making the kurtosis = 3.0 (criterion C) increases tremendously the absolute values of g and 5th, and to some extent TS1 and TS2.

2. Out of 14 basins with flood series having no significant low or high outliers/inliers, results for all the criteria and modification were obtained for only seven basins.

Category II: Inferences for Category I apply to this category also.

<u>Category III</u>. The same remarks as for Category II apply to this category. Complete results are obtained, however, for five out of seven basins.

Selection of Transformation

The power transformation is considered the best of the tested

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TABLE 8. Evaluation of Normalization of a Flood Series by 3-Parameter Lognormal Transformation

				W	'ith no m	modifica	ation			With H	1 modif	i e d			With L1	modif	ied	
;	ب ا	JSGS No.	*	а	g .	5th	TSI	TS2	a	a	5th	TS1	T\$2	а	g .	5th	TS1	T\$2
	١.	Flood	serie	s with	no sigr	nificant	high	or low	outlier	s/inlie	rs							
	1	05571000	A B C	284.0 373.0 313.0		-0.198 0.000 -4.576	0.456	0.344	4 0 2 . 0 5 4 8 . 0 3 6 1 . 0	0.000 0.037 -0.261		0.480	0.392	228.0 303.0 328.0	0.000 0.022 -0.605		0.457	0.351
	2	05572000	A B C	900.0 1007.0 -57.8	0.000 0.024 .0.571		0.506		862.0 950.0 -56.3	0.000 0.020 0.554		0.493	0.433 0.440 2.312	526.0 458.0 -67.2	0.000 -0.016 0.708	0.000	0.508 0.504 3.490	0.536
	3	05582000	A B C	722.0 219.0 -239.0	0.000 -0.097 0.4941		0.894	0.355 0.848 2.090	56.1 -798.0 -206.0	0.000 -0.111 0.417	0.000	0.893	0.809 0.807 1.698	94.0 69.3 -229.0	0.000 -0.089 0.472	0.013	0.863 0.858 1.921	0.795
-	4	05435500	A B C	289.0 311.0 378.0	0.000 0.004 -0.799		0.683	0.788		0.000 -0.009 0.775	0.000	0.610	0.677 0.675 4.346	4 4 1 . 0 5 2 5 . 0 3 5 6 . 0	0.000 0.014 -0.763		0.605	0.675
8-	5	05437000	A B C		-0.008 -0.037 0.702	0.004	0.938	1.043 1.049 3.956	4345.0 _227.0	-0.060 0.571			0.882 2.419					
	6	05437500		3256.0 1619.0 -283.0		0.325 0.015 4.934	0.971	1.064	3765.0 210S.0 -298.0	0.000 -0.053 0.687	0.000	0.992	1.076 1.096 3.829	4947.0 3657.0 -278.0		0.000	0.750 0.756 2.626	0.782
	7	05444000	A B C															
	8	05447500	A B C	276.0	- 0 . 7 0 7	- 5 . 1 8 3	1.627	1.584	268.0	- 0 . 7 1 4	- 5 . 2 3 7	1.670	1.636	283.0	-0.713	- 5 . 2 4 0	1.651	1.609
	9	05438500	A B C	1614.0 1562.0	0.000 -0.009				1213.0 989.0	0.000 -0.034	0.185 0.000			1908.0 1882.0	0.000 -0.003			1.275 1.274
	10	05439500	A B C	-115.0	0.006	3.111	4.104	4.161										

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TABLE 8. continued

			V	Vith no	modific	ation			With H	1 modi	fied			With L	1 modi	fied	
	USGS No.	*	а	9	5th	TS1	T\$2	а	g	5th	TSI	TS2	а	Ģ	5th	TSI	TS2
I	. Flood	series	with	no sign	ificant	high	or low	outlier	s/inlie	rs co	ntinueo	t					
11	05444000																
		B C	-32.4	-0.437	-3.068	1.020	0.913	-31.2	-0.429	-2.994	0.989	0.874	-40.4	-0.396	- 2 . 7 3 0	1.020	0.909
12	05445500				0.500			403.0			0.740			0.000		0.821	
		B C			0.031 -3.828				-0.038 -0.498		0.746 0.807			-0.044 -0.542		0.830 0.879	
13	05447000																
		B C	120.0	-0.785	- 5 . 2 6 5	2.082	2.039	112.0	-0.811	- 5 . 4 5 0	2.261	2.302	120.0	-0.823	- 5 . 7 7 1	2.141	2.123
14	05448000			0.000	0.333			6 S 3 . 0	0.000		0.827		965.0	0.000		0.924	
		B C		-0.055 0.712	0.015 4.859				-0.075 0.651	-0.003 4.907	0.836 3.179			-0.054 0.714			
			W	'ith no n	nodifica	tion			With H	1 modif	i e d		W i	th H1 ar	nd H2 m	odifie	b
	II. Flood	lserie	es with	n high	outlier	s/inli	ers										
1	05574000												-274.0		-0.193		
		B C												0.022 -0.118			
2	05575500			0.000			0.293		0.000			0.331	213.0			0.444	
		B C		-0.041 0.337			0.294 0.700		-0.011 -0.625			0.329 1.071		-0.010 -0.485		0.445 0.797	
3	05576500		7719.0				3 1.863										
		B C		-0.216 -0.667	0.061 -4.654		2.058 2.355										
4	05578500	A C	-236.0	0.000	0.342	1.001	1.006	116.0	0.000	-0.005	1.248	1.369	-141.0	0.000		0.975	1.022
		B C	-326.0 329.0		-0.004 -4.664			119.0 506.0	0.001 -0.736	0.000 -5.428		1.370 1.769		-0.012 -0.674		0.965 1.455	
5	05580500	0 A	-336.0	0.000	0.507	0.827	0.823	-528.0	0.000	0.694	0.790	0.740	431.0	0.000	0.231	0.959	1.027
				-0.077 -0.716	0.000 -4.997		0.810 1.531	-700.0 359.0		0.097 -4.705	0.785 1.410			-0.040 -0.728		0.949 1.752	

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TABLE 8. (concluded)

		With no modification				With H1 modified				With H1 and H2 modified						
# USGS	No. *	а	a	5th	TSI	TS2	а	ø	5th	TSI	TS2	а	٥	5th	TSI	т\$2
11.	=lood s	eries with	high	outlier	s/inli	ers co	ntinued									
6 0558	В	2301.0	-0.278	-0.001	2.415											
7 0338	C 0500 A			-0.875	2.105		2723.0	0.000	-0.345	0.944	0.963	2291.0	0.000	-0.294	0.860	0.837
	B C			0.037 -1.583	1.225 1.125		3309.0 699.0		0.007 -5.397			2679.0 588.0	0.040 -0.697		0.878 2.085	
		Wi	th no n	nodifica	ition			With L1	modifi	e d		Wit	h L1 ar	nd L2 m	odifie	d
III. F	lood se	ries with	low o	utliers	/inlie	rs										
1 0557	4500 A B C	1690.0			0.590	0.398 0.400 0.377		-0.010	0.000	0.590	0.399	-42.0	0.000 -0.047 -0.641	0.000	0.574 0.578 1.074	0.468
2 0557	6000 A	1203.0	0.000	0.451	0.696	0.571	1291.0	0.000	0.425	0.686	0.554	598.0	0.000	0.525	0.744	0.649
	B C					0.574 1.426	1025.0 -122.0		0.000 3.706			289.0 -162.0	-0.069 0.524		0.747 2.348	
3 0557	7500 A B		0.000	-0.653		0.716 0.727	246.0 343.0		-0.743 0.023			-164.0 -124.0		-0.438 -0.003		
	C			-3.988			62.7						-0.538			
4 0557	9500 A B C	877.0	0.171	-1.584 0.011 -1.113	0.898	0.731	98.4 234.0		-0.774 0.000	0.677	0.560	179.0 333.0 61.2		-0.330 0.006	0.705	0.581
5 0544	-		0.007	-1.113	1.005	0.010	02.2	-0.393	- 3 . 4 0 7	0.072	0.344	01.2	-0.304	- 3 . 4 0 2	0.009	0.000
6 0337	9500 A B	2205.0	0.039		1.015	1.011	1005.0	0.003	-0.019	0.985	1.036	1603.0 1792.0	0.020		0.970	0.990
7 0338	1500 A			-4.978 -1.425		1.417	604.0 19.5 723.0		-1.328	1.707		553.0	-0.667	-4.988	2.334	2.306

* Criterion used for deriving a

transformations for converting an observed flood series so that it approximates a normally distributed series. The pertinent reasons are:

1. Power transformed series are more stable and consistent even when some outliers/inliers are present.

2. Power transformed series derived with λ from any of the three criteria have similar statistical properties. The maximum log-likelihood method of determining λ . can be used and it is free from bias that may be attributed to g = 0 and min [|g| + |5th|] criteria.

3. Overall results obtained from the flood series analyzed with the power transformation are much better than from the Wilson-Hilferty transformation. The 3-parameter lognormal transformation is unsuitable for general use.

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METHODS OF ESTIMATING DESIGN FLOODS

Estimation of various recurrence-interval floods was performed basically with three methods: power transformation, log-Pearson type III, and mixed distribution. The background and rationale of these methods are investigated.

Power Transformation Method

1. The observed annual flood series, $Q_{\rm i},$ is normalized using the power transformation to $y_{\rm i}$ series

$$y_{i} = (Q_{i}^{\lambda} - 1)/\lambda \qquad \lambda \neq 0$$
(1)

in which the parameter λ is determined by the maximum log-likelihood method.

2. The mean, $\overline{\mathbf{y}}$, and standard deviation, s_y , of the normalized series are calculated from

$$\overline{y} = \frac{1}{n} \sum_{i=1}^{n} y_i$$

$$s_y = \sqrt{\left(\sum_{i=1}^{n} (y_i - \overline{y})^2\right)/(n-1)}$$
(2)
(3)

3. For a desired recurrence interval of T years, the probability of nonexceedance is (1 - 1/T). A standard normal deviate z_T corresponding to this probability is obtained from a p-to-z subroutine (or it can be interpolated from a normal probability table).

$$Q_T = (\lambda y_T + 1)^{1/\lambda}$$
, where $y_T = \overline{y} + z_T s_y$ (4)

Effect of Kurtosis $\neq 3$

The y. transformed series has a skew very close to zero but the kurtosis, kt, may not equal 3 as for a normal distribution:

$$kt = \frac{n^2 \sum_{i=1}^{n} (y_i - \overline{y})^4}{(n-1)(n-2)(n-3) s_y^4}$$
(5)

The normal distribution is compared with the symmetric platykurtic (kt<3) and symmetric leptokurtic (kt>3) distributions in figure 8. The normal distribution function can be modified to express these variations. The following description is based on Box and Tiao (1973).

The standard normal distribution function may be written as

$$p(x) = k \exp(-\frac{1}{2}|x|^{q})$$
 with q=2 (6)

By allowing q to take values other than 2 with the following expression

$$q = 2/(1+\beta); -1 < \beta \le 1,$$
 (7)

the class of exponential power distribution functions can be written in the general form

$$p(y|\theta,\phi,\beta) = k \phi^{-1} \exp\left(-\frac{1}{2} \left|\frac{y-\theta}{\phi}\right|^{2/(1+\beta)}\right)$$
(8)

$$k^{-1} = \Gamma(1 + \frac{1+\beta}{2}) 2^{1+\frac{1}{2}} (1+\beta)$$
(9)

in which $-\infty < y < \infty$, $\phi > 0$, $-\infty < 0 < \infty$, and $-1 < \beta \le 1$. In equation 9, 0 is a location parameter and \emptyset is a scale parameter. It can be shown that

$$E(y) = \Theta$$
 (10)

$$Var(y) = \sigma^{2} = 2^{(1+\beta)} \left\{ \frac{\Gamma[3/2 \ (1+\beta)]}{\Gamma[1/2 \ (1+\beta)]} \right\} \phi^{2}$$
(11)

The parameter 3 can be regarded as a measure of kurtosis indicating the extent of variation from the normal distribution. In particular, the distribution is normal and double exponential when $\beta=0$ and $\beta=1$, respectively, and the distribution tends to the rectangular distribution as

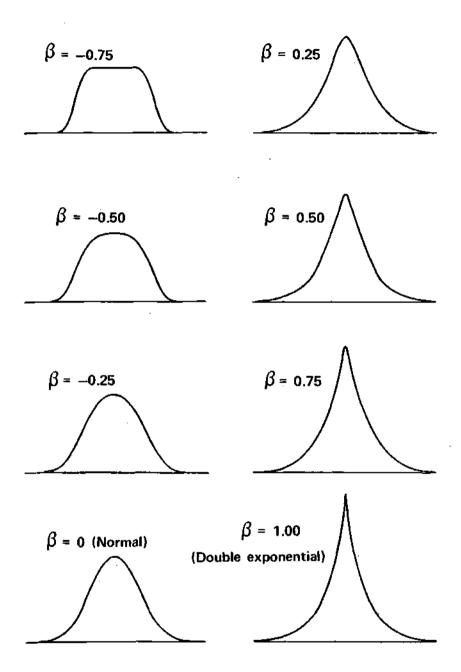


Figure 8. Platykurtic ($\beta < 0$) and leptokurtic ($\beta > 0$) distributions

 β tends to -1. The kurtosis, kt, and β are related by the following expression

$$kt = \frac{\Gamma[5/2 \ (1+\beta)] \ \Gamma[1/2 \ (1+\beta)]}{\{\Gamma[3/2 \ (1+\beta)]\}^2}$$
(12)

Values of kt corresponding to various 3 values, as obtained from equation 12 are:

β	-1.000*	-0.950	-0.900	-0.850	-0.800	-0.750	-0.700	-0.650
kt	1.800	1.807	1.824	1.851	1.884	1.923	1.968	2.017
β	-0.600	-0.550	-0.500	-0.450	-0.400	-0.350	-0.300	-0.250
kt	2.070	2.127	2.188	2.253	2.322	2.394	2.469	2.548
β	-0.200	-0.150	-0.100	-0.050	0.000	0.050	0.100	0.150
kt	2.631	2.718	2.808	2.902	3.000	3.102	3.208	3.319
β	0.200	0.250	0.300	0.350	0.400	0.450	0.500	0.550
kt	3.433	3.553	3.677	3.805	3.939	4.078	4.222	4.372
β	0.600	0.650	0.700	0.750	0.800	0.850	0.900	0.950
kt	4.527	4.688	4.856	5.029	5.209	5.396	5.590	5.791
β kt	1.000 6.000	*is the	limitir	ig case				

The kurtosis effect correction can be made in the Q_T values by modifying the z_T values. These z_T values were computed by numerical integration of equation 8 with θ =0 and \emptyset =1. These are given in Table 9 for 41 values of β lying in the range - 1 * to +1 and 6 values of T: 10, 25, 50, 100, 500, and 1000 years (or corresponding p values of 0.90, 0.96, 0.98, 0.99, 0.998, and 0.999). The various recurrence-interval floods can, thus, be computed with and without correction for kurtosis. For Q_T without correction for kurtosis, the z_T values are taken for kt=3.0. In the case of correction for kurtosis, the β value is interpolated for the sample kt, and the corresponding z_T are taken from Table 9 and used in equation 4.

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g	10	25	50	100	500	1000
-1.00	1.386	1.593	1.663	1.697	1.725	1.729
-0.95	1.384	1.592	1.665	1.708	1.762	1.777
-0.90	1.378	1.594	1.679	1.736	1.817	1.841
-0.85	1.372	1.600	1.699	1.769	1.875	1.908
-0.80	1.366	1.608	1.721	1.803	1.935	1.977
-0.75	1.360	1.618	1.744	1.839	1.996	2.047
-0.70	1.355	1.629	1.768	1.875	2.056	2.117
-0.65	1.350	1.640	1.792	1.911	2.117	2.187
-0.60	1.345	1.651	1.815	1.946	2.178	2.257
-0.55	1.340	1.661	1.838	1.982	2.238	2.328
-0.50	1.335	1.672	1.861	2.016	2.298	2.398
-0.45	1.330	1.682	1.883	2.050	2.358	2.468
-0.40	1.326	1.691	1.904	2.083	2.418	2.538
-0.35	1.321	1.700	1.925	2.116	2.477	2.608
-0.30	1.315	1.709	1.945	2.148	2.535	2.677
-0.25	1.310	1.717	1.965	2.179	2.594	2.747
-0.20	1.305	1.725	1.984	2.210	2.651	2.816
-0.15	1.299	1.732	2.002	2.240	2.709	2.885
-0.10	1.293	1.739	2.020	2.269	2.766	2.954
-0.05	1.288	1.745	2.037	2.298	2.822	3.022
0.00	1.282	1.751	2.054	2.326	2.878	3.090
0.05	1.275	1.756	2.070	2.354	2.934	3.158
0.10	1.269	1.761	2.085	2.381	2.989	3.226
0.15	1.263	1.765	2.100	2.407	3.044	3.293
0.20	1.256	1.770	2.114	2.433	3.098	3.361
0.25	1.249	1.773	2.128	2.458	3.152	3.428
0.30	1.243	1.776	2.141	2.482	3.205	3.494
0.35	1.236	1.779	3.154	2.506	3.258	3.561
0.40	1.229	1.782	2.166	2.529	3.311	3.627
0.45	1.222	1.784	2.178	2.552	3.363	3.692
0.50	1.214	1.786	2.189	2.574	3.414	3.758
0.55	1.207	1.787	2.200	2.596	3.465	3.823
0.60	1.200	1.788	2.210	2.617	3.516	3.888
0.65	1.192	1.789	2.220	2.637	3.566	3.952
0.70	1.185	1.789	2.229	2.657	3.616	4.016
0.75	1.177	1.790	2.238	2.677	3.665	4.080
0.80	1.169	1.789	2.247	2.695	3.714	4.143
0.85	1.162	1.789	2.255	2.714	3.762	4.206
0.90	1.154	1.788	2.262	2.732	3.810	4.269
0.95	1.146	1.787	2.269	2.749	3.857	4.331
1.00	1.138	1.786	2.276	3.766	3.904	4.393

Values of z_{T} for Recurrence Interval, T, of

Log-Pearson Type III Distribution Method

The Water Resources Council (1976, 1977) has recommended the following technique for fitting the log-Pearson type III, LP3, distribution to an observed annual flood series, Q_i , and for computing floods at desired recurrence intervals.

1. Compute mean $\overline{\mathbf{x}}$

$$\overline{\mathbf{x}} = \sum_{i=1}^{n} \mathbf{x}_i / \mathbf{n}$$
(13)

in which $x = \log_{10}Q$ and n = number of years or sample size.

2. Compute standard deviation, s

$$s = \left[\sum_{i=1}^{n} (x_i - \bar{x})^2 / (n-1)\right]^{0.5}$$
(14)

3. Compute skew coefficient, g

$$g = n \sum_{i=1}^{n} (x_i - \bar{x})^3 / [(n-1)(n-2) s^3]$$
(15)

4. Compute flood of recurrence interval T years, $Q_{\rm T}$

$$\log Q_{T} = \bar{x} + ks \tag{16}$$

in which k is a factor that is a function of g and the selected recurrence interval (or exceedance probability). Values of k can be obtained from a table.

Because of the errors inherent in estimating the third moment from a small sample, a regional analysis is recommended for deriving a suitable value of regional skew coefficient, g_r . The weighted skew, g_w ,

$$g_w = g w + (1 - w) g_r$$
(1/)

is used with sample $\overline{\mathbf{x}}$ and s; the weight w equals (n-25)/75. When n equals or exceeds 100, w equals unity.

Mixed Distribution Method

This is based on the mixed distribution concept and considers the observed floods (or their logarithms) to belong to two populations with means μ_1 and μ_2 , variances σ_1^2 and σ_2^2 and relative weights *a* and $1-\alpha$ of the two component distributions which may be both lognormal, normal, or any other distribution type, or a mixture of two types. The mixed distribution method developed from various studies (Singh, 1968; Singh and Sinclair, 1972; and Singh, 1974) is based on the following equations:

$$p\{x\} = \alpha p_1\{x\} + (1-\alpha) p_2\{x\} \qquad 0 \le \alpha \le 1$$
(18)

$$p_{1}\{x\} = \frac{1}{\sigma_{1} \sqrt{2\pi}} \int_{-\infty}^{x} \exp\left[-\frac{(x' - \mu_{1})^{2}}{2\sigma_{1}^{2}}\right] dx'$$
(19)

$$p_{2}{x} = \frac{1}{\sigma_{2} \sqrt{2\pi}} \int_{-\infty}^{x} exp \left[-\frac{(x' - \mu_{2})^{2}}{2\sigma_{2}^{2}} \right] dx'$$
 (20)

in which p is the probability of being equal to or less than x. The component distributions are taken as log-normal (x = log Q) in the above equations. The mixed distribution parameters are linked to sample μ , σ^2 , and g values according to the following equations (Cohen, 1967):

$$\mu = a\mu_1 + (1-a) \mu_2$$
(21)

$$\sigma^{2} = a\sigma_{1}^{2} + (1-a)\sigma_{2}^{2} + a(1-a)(\mu_{2} - \mu_{1})^{2}$$
(22)

$$g = [3\alpha(1-\alpha)(\mu_1 - \mu_2)(\sigma_1^2 - \sigma_2^2) + \alpha(1-\alpha)(1-2\alpha)(\mu_1 - \mu_2)^3]/\sigma^3$$
(23)

Evaluation of Parameters

The distribution parameters for the mixed distribution were obtained by using the Generalized Reduced Gradient Method, a nonlinear programming algorithm; the computer program for which was available from the University of Illinois. Two nonlinear objective functions, i.e., minimization of $\Sigma(\Delta z)^2$ and $\Sigma[\Delta z]$, Δz equals the difference between the standard deviate corresponding to the observed probability equal to (m - 0.38)/(n + 0.24) and that fitted corresponding to p from equation 18, were considered subject to the following constraints:

$$1 = a_1 + a_2$$
 (24)

$$\mu = a_1 \mu_1 + a_2 \mu_2 \tag{25}$$

$$\sigma^{2} = a_{1}\sigma_{1}^{2} + a_{2}\sigma_{2}^{2} + a_{1}a_{2}(\mu_{2} - \mu_{1})^{2}$$
(26)

$$g\sigma^{3} = a_{1}m_{1}(3\sigma_{1}^{2} + m_{1}^{2}) + a_{2}m_{2}(3\sigma_{2}^{2} + m_{2}^{2})$$
 (27)

$$kt\sigma^{4} = a_{1}(3\sigma_{1}^{4} + 6m_{1}^{2}\sigma_{1}^{2} + m_{1}^{4}) + a_{2}(3\sigma_{2}^{4} + 6m_{2}^{2}\sigma_{2}^{2} + m_{2}^{4})$$
(28)

in which $\mathbf{m}_1 = \mu_1 - \mu$, $\mathbf{m}_2 = \mu_2 - \mu$, $\mathbf{a}_1 = \alpha$, and $\mathbf{a}_2 = 1 - \alpha$. Use of $\Sigma |\Delta z|$ was found to give more consistent solutions than $\Sigma (\Delta z)^2$.

The asymmetry of the power-transformed or log-transformed flood series, as evidenced by the kt being lower or higher than 3 and by the 5th moment being significantly different from zero, is accommodated easily by the mixed distribution concept.

Various Recurrence-Interval Floods

These floods are calculated by a reiterative process. For the desired recurrence interval, value of p is obtained from (1 - 1/T). Starting from a given or assumed value of x, z_1 and z_2 are calculated from:

$$z_1 = \frac{x - \mu_1}{\sigma_1}$$
 and $z_2 = \frac{x - \mu_2}{\sigma_2}$ (29)

The corresponding p_1 and p_2 are obtained from the z-to-p subroutine. The p is calculated from p_1 and p_2 with equation 18. If this p is equal

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to, or within a specified tolerance of, the p corresponding to the desired recurrence interval, the value of x yields the logarithm of Q_T . Otherwise, by an iterative process, a value of x is determined that meets the p criterion.

NEW FLOOD-FREQUENCY METHODOLOGY

A new flood frequency methodology has been developed and computerized. It detects objectively the outliers and inliers at various probability levels and modifies them if needed. The computer program prints 2- to 1000-yr floods from power transformation, both with and without kurtosis correction, from log-Pearson type III distribution, both with sample skew and weighted skew, and from the mixed distribution, for levels 0, 1, 2, 3, 4, 5, and 6. The level 0 corresponds to processing of data without any testing for outliers and/or inliers. Levels 1 through 6 correspond to outlier-inlier probability pairs of .01, .99; .05, .95; .10, .90; .20, .80; .30, .70; and .40, .60. The relevant information on statistics of the three methods and the given and modified values of outliers/inliers are also printed at all the levels. The salient features of this new methodology are described in the rest of this section.

Compact Departure Table, Probability Levels and Windows

A compact departure table, containing the test value, at 6 probability levels is given in Table 10. The low 1 through 5 denote the lowest to the 5th low value from the low end and the high 1 through 5 denote the highest to the 5th high value from the high end. Considering level 1 and low 1, a departure Δ of -0.689 or less (in the algebraic sense) will indicate an inlier at 1% or less nonexceedance probability (or 99% or higher exceedance probability), and a departure of 1.029 or more will indicate an outlier at 99% or higher nonexceedance probability (or 1% or lower exceedance probability). Similarly, for level 1 and high 1, a departure of 0.679 or more

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			Test values of departures							
Window	р	Outlier/	Low 1	Low 2	Low 3	Low 4	Low 5			
		Inlier	15-100	20-100	25-100	30-100	40-100			
1	<.01	Inlier	<689	<495	<412	<363	<327			
	>.99	Outlier	>1.029	>0.643	>0.498	>0.418	>0.368			
2	<.05	Inlier	<532	<369	<303	<264	<237			
	>.95	Outlier	>0.681	>0.421	>0.337	>0.285	>0.253			
3	<.10	Inlier	<441	<299	<243	<211	<188			
	>.90	Outlier	>0.503	>0.321	>0.254	>0.217	>0.193			
_										
4	<.20	Inlier	<318	<209	<167	<143	<127			
	>.80	Outlier	>0.297	>0.197	>0.159	>0.137	>0.123			
5	<.30	Inlier	<221	<141	<110	<093	<082			
	>.70	Outlier	>0.161	>0.112	>0.092	>0.081	>0.073			
C	< 10	Tralian	< 100	< 000			< 0.12			
6	<.40 >.60	Inlier Outlier	<132 >0.052	<080 >0.043	<060 >0.037	<050 >0.034	<043 >0.032			
	2.00	OUCTICE	20.002	20.015	20.007	20.001	20.002			
			High 1	High 2	High 3	High 4	High 5			
			15-100	20-100	25-100	30-100	40-100			
1	<.01	Outlier	<-1.054	<654	<511	<429	<377			
T	>.99	Inlier	>0.679	>0.488	>0.407	>0.358	>0.323			
2	<.05	Outlier	<683	<433	<341	<290	<256			
	>.95	Inlier	>0.529	>0.369	>0.300	>0.263	>0.235			
3	<.10	Outlier	<500	<322	<258	<221	<195			
5	>.90	Inlier	>0.438	>0.299	>0.241	>0.209	>0.186			
4	<.20	Outlier	<295	<197	<161	<139	<124			
	>.80	Inlier	>0.317	>0.209	>0.166	>0.143	>0.126			
5	<.30	Outlier	<159	<112	<094	<082	<074			
-	>.70	Inlier	>0.221	>0.140	>0.110	>0.093	>0.082			
6	<.40	Outlier	<051	<043	<039	<035	<032			
	>.60	Inlier	>0.132	>0.079	>0.060	>0.050	>0.043			

Note: 15 - 100,...., and 40 - 100 denote the range of sample size n in years.

indicates an inlier at 99% or higher nonexceedance probability (or 1% or lower exceedance probability), and a departure of -1.054 or lower (in the algebraic sense) indicates an outlier at 1% or lower nonexceedance probability (or 99% or higher exceedance probability). Thus, the probability pairs for outliers and inliers have the connotation of the same relative severity.

The concept of the levels and windows is clarified in figure 9 in which departures for the high 1 are plotted on normal probability paper. For the outliers, window 1 contains \triangle values ≤ -1.054 , window 2 contains \triangle values such that $-1.054 < \Delta \leq -0.683$, and so on for windows 3 through 6. For the inliers, window 1 contains \triangle values ≥ 0.679 , window 2 contains \triangle values such that $0.679 > \Delta \geq 0.529$, and so on for windows 3 through 6. The departures for the low 1 are plotted on normal probability paper in figure 10. If some outliers and/or inliers are found in window 1, the same are modified to respective values at level 1, and the procedure is followed sequentially from one window to the other. If no outliers and/or inliers are detected in a particular window, no modifications are done, and the program moves to the next window after developing and printing distribution statistics and flood estimates.

Plotting Position

The plotting position for the observed floods has been a matter of considerable controversy. A general formula for computing plotting position (Harter, 1971) is:

p = (m - a)/(n - a - b + 1)(1)

in which m is the rank order of flood values arranged in an ascending order

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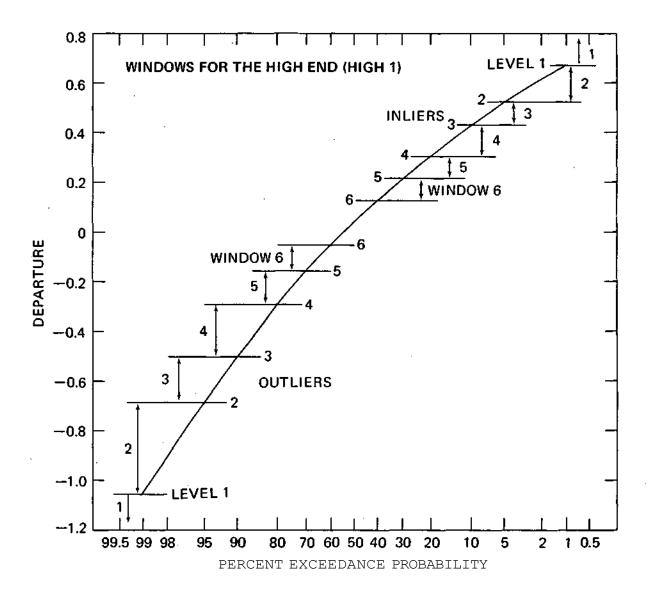


Figure 9. Levels and windows for the outliers/inliers at the high end

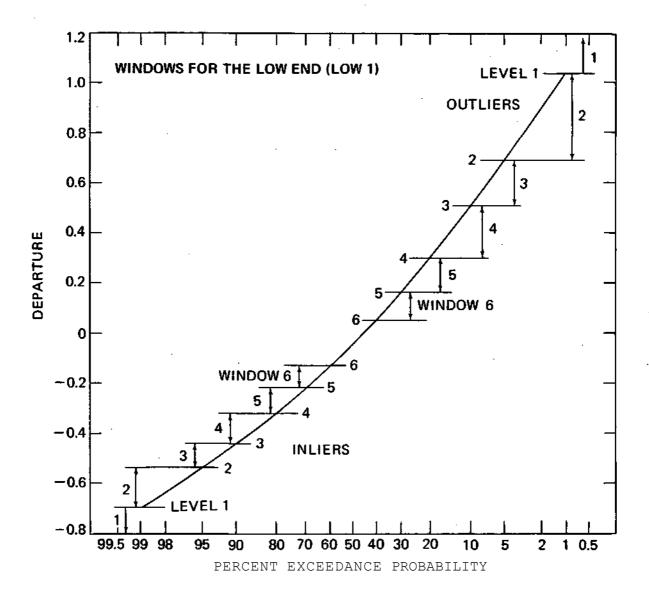


Figure 10. Levels and windows for the outliers/inliers at the low end

of magnitude, p is the probability of nonexceedance, and a and b depend on the distribution. For a symmetrical distribution, a equals b, and equation 1 can be rewritten as:

$$p = (m - \alpha)/(n + 1 - 2\alpha)$$
(2)

The commonly used Weibull plotting position

$$p = m/(n+1) \tag{3}$$

is obtained by putting $\alpha = 0$. However, an α value of about 0.38 (Cunnane, 1978; Blom, 1958) is the best for the normal distribution. Cunnane states that the Weibull plotting formula is exact when the distribution is uniform and that the Gringorten formula, with $\alpha = 0.44$, is satisfactory for exponential distributions.

In calculating $\sum |\Delta z|$ for evaluating the mixed distribution parameters by the nonlinear programming algorithm, the Δz is obtained from

 $\Delta z = (z \text{ corresponding to standard normal deviate for observed})$

probability p)-(z fitted from the mixed distribution with

$$\mathbf{p} = a\mathbf{p}_1 + (1 - a) \mathbf{p}_2$$
(4)

The observed probability p is obtained from

$$p = \frac{m - 0.38}{n + 0.24}$$
(5)

in which $a = b = \alpha = 0.38$.

The Flow Chart

The detection and modification of outliers and/or inliers as well as flood frequency analysis follow the flow chart given in figure 11. Some relevant explanations to clarify the methodology and the computer program are given in the following few pages. The sequence numbers correspond to the numbers attached to various boxes in the flow chart.

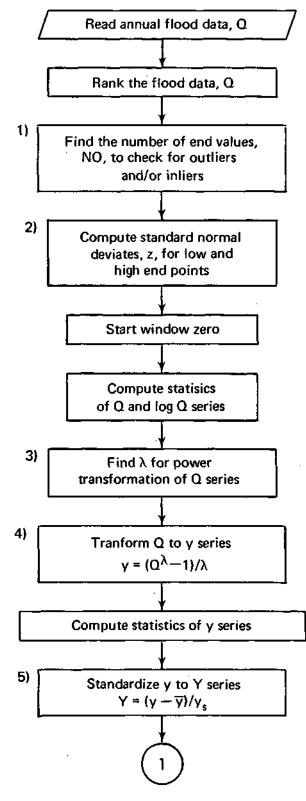


Figure 11. Flow chart for the computer program for flood frequency analyses

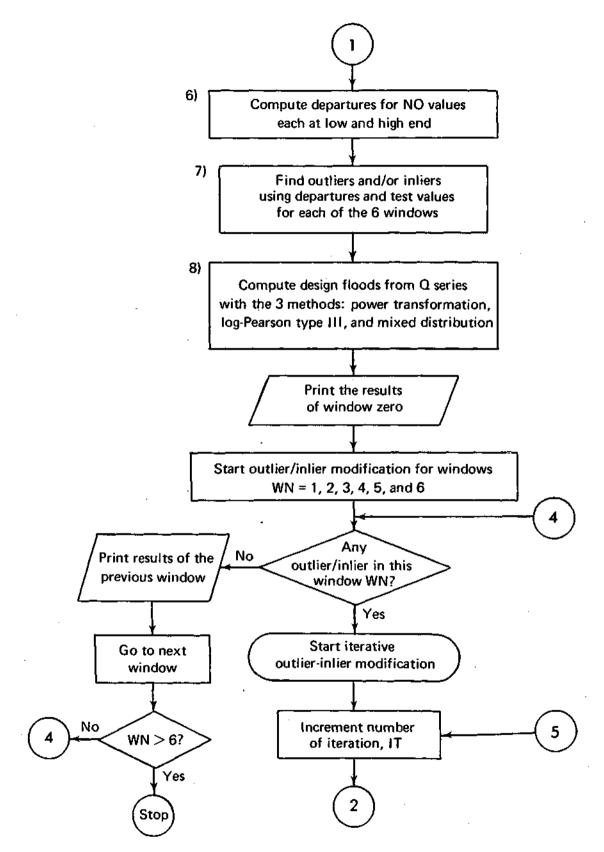


Figure 11. —Continued

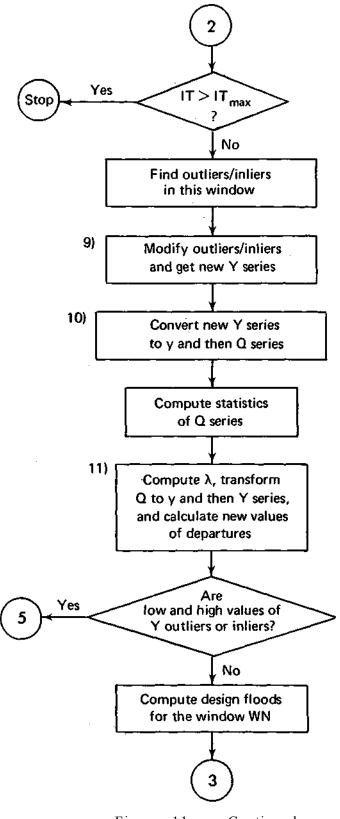
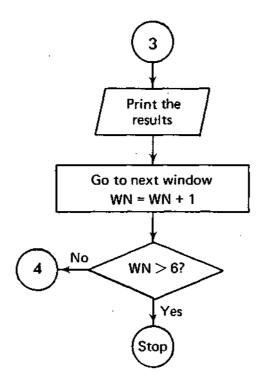


Figure 11. —Continued





1.) Number of low as well as high floods, NO, can be provided as an input information or computed from some expression such as NO = [n/10] where n equals the number of floods in the sample series and NO = 5 for $n \ge 50$.

2.) Standard normal deviate of rank m, $z_{\rm m}$, is computed by converting probability p , obtained from

$$p_{m} = (m - \alpha)/(n + 1 - 2\alpha)$$
 (2)

with a interpolated from the smoothed a values for the rank m and sample size n (see step 8 of Generation of Departures, p. 22 to 23) to z with the p-to-z subroutine, assuming standard normal distribution.

3.) The parameter λ is computed by the maximum log-likelihood method from the given flood series.

4.) The given flood series is transformed to y series by

$$y_{i} = \frac{Q_{i}^{\lambda} - 1}{\lambda}$$
, $i = 1, 2, ..., n$ (6)

and the process is termed normalization by power transformation.

5.) The y series is standardized to Y series with

$$Y_{i} = \frac{y_{i} - \overline{y}}{y_{s}}$$
(7)

in which $\frac{1}{\mathbf{y}}$ and y_s are the mean and standard deviation of the transformed series, y.

6.) The departures, $\Delta_{\!m}$ for NO values at the low end, as well as at the high end, are obtained from

$$\Delta_{m} = z_{m} - Y_{m}$$
 for $m = 1, ..., NO$ (8)

7.) Outliers and inliers, if any, are detected in each of the 6 windows according to 6 probability levels using Δ_m and test values in Table 10.

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8.) The floods corresponding to 2-, 10-, 25-, 50-, 100-, 500-, and 1000-year recurrence intervals are computed with three methods: 1) power transformation with and without kurtosis correction; 2) log-Pearson type III distribution, with sample as well as with weighted skew; and 3) mixed distribution; without any modification of outliers and/or inliers, i.e. with window 0.

9.) Modification process is illustrated in figure 12 as an example. Consider 2 low and 2 high values (NO=2) as candidates for outliers. No outliers are detected in window 1; 1 high and 1 low outliers are detected in window 2; and 2 high and 1 low outliers are detected in windows 3 through 6. The values of Y_m for the detected outliers and/or inliers in a window are changed to (z_m minus departure) values for that window. This gives the new Y series in which the Y outlier/inlier values have been replaced by the corresponding threshold values.

10.) The new Y series is destandardized with \overline{y} and y_{g} from equation 7, to get new y series.

$$y = \overline{y} + y_{S} Y \tag{9}$$

The new y series is detransformed with A from step 3 for the previous Q series, to obtain the new Q series (with values of Q changed only for the outliers/inliers).

11.) The new value of X is computed for the new or modified Q series. With this X, the modified Q series is transformed to y series, which in turn is standardized to Y series, and values of departures are obtained for the NO points at both low and high end.

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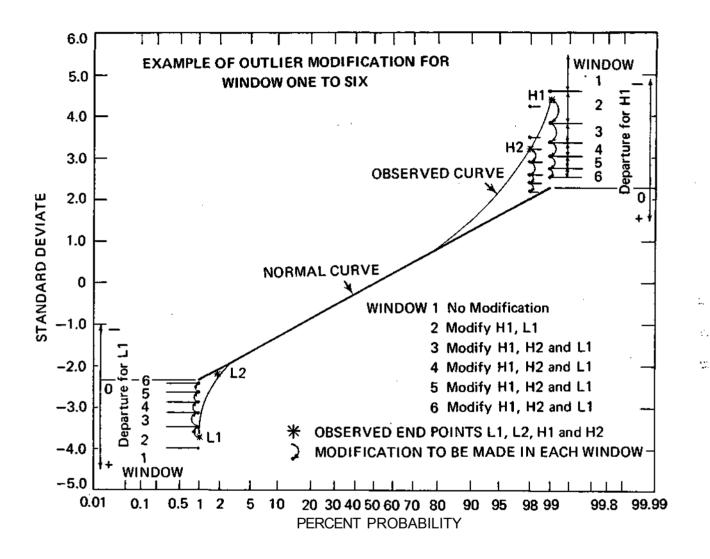


Figure 12. An example of outlier/inlier modification

An Example

The new methodology of detection and modification of any outliers/inliers and flood frequency analysis is explained by an actual example of observed floods for the Sangamon River at Oakford (USGS No. 05 583000, drainage area 5093 sq miles, n = 62 years).

Ranked discharge, Q, data in cfs

3,480	3,800	4,630	5 , 670	5,960	6,430	6 , 720	8,400
9,100	10,000	10,500	10,500	11,000	11,300	12,000	13,300
15,100	15 , 700	16,500	16,500	17,200	19,000	19,400	20,000
20,300	20,800	20,800	21,200	21,400	21,500	22,200	23,700
24,100	24,700	25,000	25,200	25,600	26,200	28,400	29,100
30,100	30,200	31,200	31,400	33,300	33,800	34,600	34,700
36,300	36,300	37,600	37,900	38,000	38,300	42,300	42,800
42,900	44,700	45,800	46,300	55 , 900	123,000		

According to Singh (1980), the maximum flood of 123,000 cfs which occurred on May 20, 1943 was caused by a 50-year storm, covering most of the basin, over very wet antecedent soil moisture conditions, giving a runoff factor of 2.2 times that for the next 4 high floods caused by 10- to 25-year storms.

This gaging station also suffers from junction problem caused by two major tributaries. Salt Creek and Sangamon River join 9 miles upstream of Oakford. Drainage areas above the gaging stations on these tributaries are 1804 square miles (5 miles upstream of confluence) and 2618 square miles (49 miles upstream of confluence), respectively. The relevant

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statistics for the first 5 top floods for the concurrent record of 1942-1979 at Oakford and corresponding floods at Greenview (Salt Creek) and Riverton (Sangamon River) are given below.

Flood	d at Oakfo	ord gage	Salt C	reek near	Greenview	Sanga	mon at Ri	verton
Rank	Date	Peak	Rank	Date	Peak	Rank	Date	Peak
1	5-20-43	123,000	1	5-19-43	41,200	1	5-19-43	68 , 700
2	4-15-79	55,900	3	4-13-79	30,500	2	4-12-79	44,200
3	4-25-73	45,800	6	4-25-73	21,000	6	4-24-73	27,000
4	4-26-44	44,700	5	4-25-44	22,000	3	4-25-44	30,600
5	6-25-74	42,900	2	6-24-74	38,100		6-24-74	

For the flood peaks from Greenview and Riverton to coincide at Oakford, the peak at Riverton should occur a day before that at Greenview and the peak at Greenview should occur about 6 hours before that at Oakford. Concurrent maximum floods at Greenview and Riverton produce the maximum flood at Oakford. An analysis of all the floods at the 3 stations indicates that for floods exceeding a 2-year flood, there is only one chance out of 5 that the tributary floods will be in phase to produce a high flood at Oakford.

Statistics of Q data

<u>Series</u>	Mean	<u>s</u>	<u>8</u>	<u>kt</u>	5th moment
Q	25480	17821	2.739	16.573	93.480
log Q	4.311	0.307	-0.562	3.421	-3.027

It is evident that log transformation makes the series much closer to normal.

Parameter for power transformation (determined by the maximum log-likelihood method), λ = 0.254

Power transformed series, $y = (Q^{\lambda} - 1)/\lambda$

Series	Mean	<u>s</u>	g	<u>kt</u>	5	th moment
У	45.816	8.580	0.018	3.8	20	3.900
y values	(L1 to L5):	27.304,	28.010,	29.654,	31.428,	31.879
y values	(H5 to H1):	55.814,	56.184,	56.350,	59.305,	73.331
Standardized	series, Y =	(y - y)/y	s			
Y values	(L1 to L5):	-2.158,	-2.075,	-1.884,	-1.677,	-1.624
Y values	(H5 to H1)	: 1.165,	1.208,	1.228,	1.572,	3.207
<i>Theoretical stand</i> values o:		viates, co	orrespond	ing to	5 low	and 5 high
z values	(Ll to L5):	-2.342,	-1.958,	-1.739,	-1.580,	-1.454
z values	(H5 to H1)	: 1.454,	1.580,	1.739,	1.958,	2.342
A compari	ison of Y and	l z value	s indicat	tes that	Y(Ll) i	s an inlier,
Y(L2 to L5) a:	re outliers,	Y(H5 to	H2) are	inliers,	and Y(H	l) is an outlier.

Outlier/Inlier detection:	Outliers	and	inliers	are	denoted	by	0	and	I.
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Window	<u>L1</u>	<u>L2</u>	<u>L3</u>	<u>L4</u>	<u>L5</u>	<u>H5</u>	<u>H4</u>	<u>H3</u>	<u>H2</u>	<u>H1</u>
1							I	I		
2						I	I	I	I	0
3						I	I	I	I	0
4					0	I	I	I	I	0
5		0	0	0	0	I	I	I	I	0
6	I	0	0	0	0	I	I	I	Т	0

Floods in cfs for recurrence interval (years)												
Method	2	10	25	50	100	500	1000					
PT, kt=3.0	21,738	47,712	61 , 422	71 , 717	82,029	106 , 247	116 , 843					
PT, sample kt	21,738	46,495	62 , 345	75 , 439	89,513	126,090	143,606					
LP3, sample g	21,857	48,001	60 , 903	70,090	78,851	97 , 730	105,279					
LP3, weighted g	21,649	48,439	62 , 293	72,434	82,331	104,456	113 , 630					
Mixed distrib.	23,411	43 , 791	57 , 244	71 , 571	89 , 279	140,073	166 , 295					

Design floods with original discharge data

Fifty- to 1000-year floods with correction for sample kurtosis (3.820, which is greater than 3.00) become progressively higher than those with no kurtosis correction as the recurrence interval increases.

Window 1

Inliers detected: 3rd and 4th high points

After 2 iterations, the modified values of standardized series, Y, are

Y values (L1 to L5): -2.157, -2.074, -1.883, -1.676, -1.623

Y values (H5 to H1): 1.160, 1.226*, 1.336*, 1.565, 3.191

The H3 or Y_{60} is modified as explained below.

 Y_{60} is increased from 1.228 so that it just gets into the next (or the second) window by making it equal to 1.739 (theoretical standard normal deviate) - 0.407 (departure for H3, window 1) × (1-0.01); or 1.336. Factor 0.01 reduces the number of iterations and just caries H3 or Y_{60} into the next lower window.

Modified data after destandardization and detransformation:

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New X value = 0.252

Power transformed data:

y (L1 to L5): 27.011, 27.706, 29.323, 31.067, 31.510

y (H5 to H1): 54.981, 55.539, 56.464, 58.398, 72.109 Statistics of y series:

	Mean		<u>s</u>		<u>g</u>	<u>kt</u>	<u>5th m</u>	oment
	45.202	1	8.43	5 (0.017	3.788	3.7	63
Stan	dardize	ed se	eries,	Y				
	Y (L1	to L	5): -	2.156,	-2.074,	-1.882,	-1.676,	-1.623
	Y (H5	to	H1):	1.159,	1.226,	1.335,	1.564,	3.190
New	departu	ires	are:					

△ (L1 to L5): -0.185, 0.116, 0.144, 0.095, 0.169

 \vartriangle (H5 to H1): 0.294, 0.355, 0.403, 0.394, -0.848

A check with test departures shows no inliers/outliers in window 1.

Design floods with modified Q series (window 1)

Method	2	10	25	50	100	500	1000
PT, kt=3.0	21 , 761	47 , 882	61 , 700	72 , 086	82 , 498	106 , 973	117 , 692
PT, sample kt	21,761	46 , 701	62 , 607	75 , 729	89,814	126 , 341	143,806
LP3, sample g	21,872	48 , 179	61,219	70 , 526	79 , 420	98,644	106 , 354
LP3, weighted g	21 , 671	48 , 602	62 , 567	72,802	82,803	105 , 198	114,498
Mixed distrib.	23 , 391	44,008	57 , 588	71,987	89 , 772	140 , 791	167 , 152

Windows 2 through 6

Outliers are modified similarly for the successive windows 2 through 6. The results of this analyses are presented in Table 11 (contains

	TABLE 11.	. Sample	e Compute	er Output	with th	ne New Me	thodolog	У
STATION	NO. 55830)00 SAN	IGAMON RI	iver at 0	akf0rd			
DRAINAG				s of Recc		(1910-1	979)	
L	EVEL NO.	0	1	2	3	4	5	6
М	ETHOD			100-Year	Flood i	ln cfs		
	ransform, PT							
	t = 3.0	82,029	82,498	78 , 988	75 , 264	72 , 347	70 , 357	69,059
	ample kt	89,513	89,814	83,010	76 , 541	71 , 464	68,673	67 , 080
Log Tra			=	==				CT 1 1 1 1 1 1 1 1 1 1
	Sample skew	78,851	79,420	77,082	74,442	72,444	70,037	67,484
	Neighted skew		82,803	81,409	79,623	78,185	77,061	76,063
	Distrib., MD	89,279	89,772	82,699	75,229	69 , 725	69 , 535	70,023
Ty	pe No.		Observ	ved and M	odified	Floods :	in cfs	
Lo	w 1*	3,480	3,480	3,480	3,480	3,480	2,927	2,353
	2*	3,800	3,800	3,800	3,800	3,800	3,800	3,800
	3*	4,630	4,630	4,630	4,630	4,630	4,630	4,748
	4*	5,670	5,670	5,670	5,670	5,670	5,670	5,763
	5*	5,960	5,960	5,960	5,960	6,142	6,448	6,654
Hi	gh 5*	44,700	44,700	46,285	46,909	47 , 461	47 , 649	47,927
	4*	45,800	46,403	48,920	49 , 564	50 , 127	50 , 322	50,591
	3*	46,300	49,331	52 , 305	52 , 933	53 , 542	53 , 703	53 , 999
	2«	55 , 900	55 , 900	56 , 753	57 , 476	58 , 223	58,451	58,810
	1*	123,000	123,000	106,954	92 , 964	81,678	75 , 145	70,771
METHOD	STATISTICS			Values	of Stat	tistics		
PT	mean	45.816	45.201	63.589	93.389	131.737	180.224	230.869
	std dev	8.580	8.435	13.449	22.218	34.276	50.703	68.820
	skew	.018	.017	001	022	043	055	062
	kurtosis,kt	3.820	3.788	3.447	3.149	2.895	2.792	2.746
	5th moment	3.900	3.763	2.612	1.651	.856	.416	.115
	lambda	.254	.252	.300	.352	.397	.437	.468
LP3	mean	4.311	4.312	4.312	4.311			
	std dev		.308	.307				
	sample skew		556		655			
	kurtosis,kt				3.202		3.317	
	5th moment	-3.027	-3.006	-3.682	-4.227	-4.606	-5.653	-7.327
MD	weight 'a'	.648	.652	.640	.605	.515	.380	.347
	mul	4.221	4.222	4.207	4.185	4.135	4.046	4.032
	mu2	4.477	4.479	4.497	4.505	4.498	4.470	4.455
	sigmal	.338	.338	.330	.321	.306	.294	.315
	sigma2	.121	.122	.118	.123			
	Test Stat	4.718	4.568	3.925	3.634	3.396	2.906	3.000
* H	ligh & low fl	oods con	sidered	for outli	er dete	ction and	d modifi	cation

* High & low floods considered for outlier detection and modification

100-year floods for windows 0 through 6, successive modification of low and high values and sample statistics) and in Table 12 (contains 2-, 10-, 25-, 50-, 100-, 500-, and 1000-year floods for the various methods for windows 0 through 6; 0 window corresponds to no modification of Q values). The observed and modified floods (5th window) as well as the fitted mixed distribution curve are shown in figure 13.

TABLE 12. Sample Computer Output with the New Methodology

STATION NO. 5583000 SANGAMON RIVER AT OAKFORD DRAINAGE AREA 5093.0 Sq Mi Years of Record 62 (1910-1979)

VARIOUS RECURRENCE-INTERVAL FLOODS

METHOD #	Fl	ood in c	fs for F	Recurrenc	e Interv	vals (Yea	ars)
	2	10	25	50	100	500	1000
PT, kt=3.0 0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	21,738 21,738 21,857 21,649 23,411	47,712 46,495 48,001 48,439 43,791	61,422 62,345 60,903 62,293 57,244	71,717 75,439 70,090 72,434 71,571	89,513 78,851	106,247 126,090 97,730 104,456 140,073	143,606 105,279 113,630
PT, kt=3.0 1 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	21,761 21,761 21,872 21,671 23,391	47,882 46,701 48,179 48,602 44,008	61,700 62,607 61,219 62,567 57,588	72,086 75,729 70,526 72,802 71,987	89,814 79,420	106,973 126,341 98,644 105,198 140,791	143,806 106,354 114,498
PT, kt=3.0 2 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	21,994 21,994 21,999 21,736 23,750	47,211 46,551 47,795 48,358 43,789	60,075 60,663 60,198 61,952 55,334	69,581 71,644 68,894 71,829 67,236	83,010	-	123,242 101,138 111,318
PT, kt=3.0 3 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	22,191 22,191 22,084 21,759 23,879	46,341 46,126 47,217 47,918 43,775	58,211 58,416 58,918 61,058 53,865	66,835 67,517 66,972 70,517 63,160	75,264 76,541 74,442 79,623 75,229	97,377 89,841 99,545	102,471 106,350 95,733 107,633 134,446
PT, kt=3.0 4 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	22,358 22,358 22,146 21,777 23,612	45,612 45,762 46,743 47,540 44,352	56,702 56,548 57,914 60,314 54,078	64,652 64,164 65,495 69,445 61,641	72,347 71,464 72,444 78,185 69,725	87,571 86,507 97,135	91 , 789
PT, kt=3.0 5 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	22,484 22,484 22,271 21,807 23,451	45,139 45,404 46,392 47,424 44,336	55,687 55,380 56,882 59,895 54,384	63,169 62,230 63,817 68,707 -61,915	70,357 68,673 70,037 77,061 69,535	86,256 82,580 82,195 94,894 88,063	88,212
PT, kt=3.0 6 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	22,588 22,588 22,454 21,869 23,351	44,850 45,170 46,096 47,438 44,054	55,035 54,659 55,831 59,618 54,303	62,207 61,087 62,056 68,108 62,069	69,059 67,080 67,484 76,063 70,023	84,100 79,822 77,632 92,726 89,838	90,301 84,922 81,145 99,202 99,332

= level number

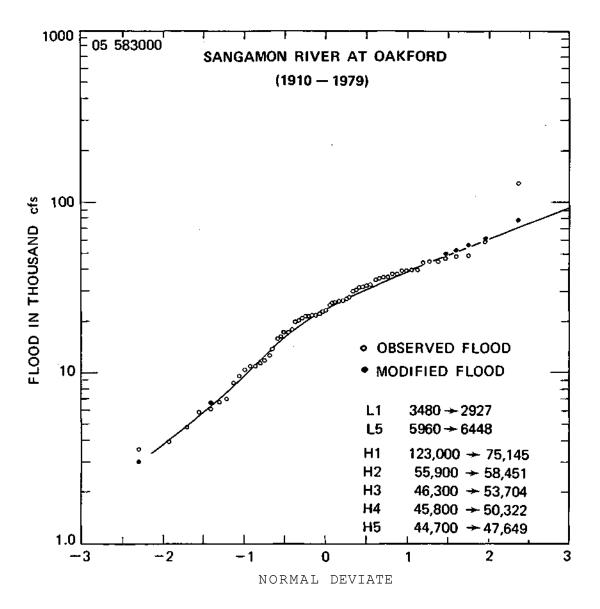


Figure 13. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Sangamon River at Oakford

FLOOD FREQUENCY ANALYSES

The developed flood frequency methods were applied to 37 observed annual flood series for drainage basins with area varying from 11 to 9551 square miles and with records of about 20 to 67 years. The gaging stations and drainage basins above these stations lie in the major river basins of the Sangamon, Rock, and Little Wabash Rivers. The information on USGS gaging station number; the name of the stream and the gaging station; length of record, n, in years; the drainage area, A, in square miles; the main channel length, L, in miles; and the main channel slope, S, in ft/mi are given in Table 13 for each of the 37 basins.

With the computer program developed in this study, flood frequency analyses were carried out with the power transformation (with and without correction for kurtosis), with the log-Pearson type III, or LP3, distribution (with sample as well as with weighted skew), and with the mixed distribution, MD, method. These analyses indicated that the MD flood estimates derived in window 5, after detection and modification of any outliers/inliers, were generally satisfactory. Therefore, some results of analyses are presented only for windows 0 (in which no outlier/inlier detection and modification is attempted) and 5. Window 5 implies that outliers/inliers occurring on the average more often than in 3 samples out of 10 are not modified. Even with the small-sample bias, derivation of distribution parameters from the observed flood series, and acceptance of these parameters as representative of population parameters, the window 5 is believed to yield not only satisfactory flood estimates but also satisfactory distribution parameters.

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Table 13. Study Basins and Pertinent Data

<u>No.</u>	USGS No.	Stream and station name	<u>n</u>	<u>A</u>	<u>L</u>	<u>s</u>
	SANGAMON R	IVER BASIN				
1	05 571000	Sangamon River at Mahomet	31	362	56.41.	3.59
2	05 572000	Sangamon River at Monticello	71	550	80.04	2.75
3	05 572500	Sangamon River near Oakley	27	774	97.98	2.21
4	05 574000	S.F. Sangamon River near Nokomis	29	11.0	4.89	18.80
5	05 574500	Flat Branch near Taylorville	30	276	47.49	2.01
6	05 575500	S.F. Sangamon River near Kincaid	36	562	51.07	2.01
7	05 576000	S.F. Sangamon River near Rochester	30	867	84.82	1.32
8	05 576500	Sangamon River at Riverton	67	2618	164.83	1.48
9	05 577500	Spring Creek at Springfield	32	107	29.37	5.39
10	05 578500	Salt Creek near Rowell	37	335	53.80	2.59
11	05 579500	Lake Fork near Cornland	32	214	37.00	4.65
12	05 580000	Kickapoo Creek at Waynesville	32	227	36.08	6.23
13	05 580500	Kickapoo Creek near Lincoln	35	306	54.48	5.12
14	05 581500	Sugar Creek near Hartsburg	35	333	42.77	5.76
15	05 582000	Salt Creek near Greenview	38	1804	114.68	2.22
16	05 582500	Crane Creek near Easton	30	26.5	4.30	2.16
17	05 583000	Sangamon River at Oakford	62	5093	222.33	1.27
	ROCK RIVER	BASIN (in Illinois)				
1	05 435500	Pecatonica River at Freeport	66	1326	99.14	2.01
2	05 437000	Pecatonica River at Shirland	32	2550	118.50	2.01
3	05 437500	Rock River at Rockton	40	6363	178.14	.84
4	05 438250	Coon Creek at Riley	18	85.1	16.45	5.72
5	05 438500	Kishwaukee River at Belvidere	40	538	41.31	4.59
6	05 439500	S. Br. Kishwaukee R. near Fairdale	40	387.	40.29	2.27
7	05 440000	Kishwaukee River near Perryville	40	1099	52.97	4.07
8	05 440500	Killbuck Creek near Monroe Center	40	117	26.80	6.34
9	05 441000	Leaf River at Leaf River	40	103	18.27	10.45
10	05 443500	Rock River at Como	65	8755	266.76	1.00
11	05 444000	Elkhorn Creek near Penrose	40	146	38.97	4.28
12	05 445500	Rock Creek near Morrison	32	158	38.68	3.91
13	05 446500	Rock River at Joslin	40	9551	309.23	1.11
14	05 447000	Green River at Amboy	40	201	23.63	3.85
15	05 447500	Green River near Geneseo	43	1003	80.41	2.53
16	05 448000	Mill Creek at Milan	40	62.4	22.62	7.44
	LITTLE WAB	ASH RIVER BASIN				
1	03 379500	Little Wabash River below Clay City	65	1131	114.10	2.01
2	03 380475	Horse Creek near Keenes	19	97.2	26.38	4.07
3	03 380500	Skillet Fork near Wayne City	51	464	59.52	1.90
4	03 381500	Little Wabash River at Carmi	40	3102	207.10	1.16

Sensitivity of NO

The number of outliers/inliers, NO, that may be checked at each end of the ranked flood series was obtained from [n/10]; NO = 5 for n > 50. Frequency analyses were made with this NO as well as with other higher or lower numbers of outliers/inliers, designated as NO₁ and NO₂. The 100-year floods obtained in window 5 from LP3 with sample skew and from MD for 2 or 3 values of NO are given in Table 14 for 28 basins. The 100-year floods for window 0 as well as window 5 with NO are given for all the 37 basins. It is evident from the flood values for different values of NO in window 5 that these floods do not differ from each other very much, except in some cases where the observed flood series indicates more outliers/inliers than given by NO.. and NO . The NO can be used as a limiting guide in general. If the number of outliers/inliers is less than NO, the floods which are not outliers/inliers will not be detected or modified.

In the case of Salt Creek near Rowell in the Sangamon Basin (No. 10 and USGS No. 05 578500 in Table 14), NO = [37/10] or 3. The 4 lowest and 4 highest floods together with any modification of these floods with NO = 3 and NO₁ = 1 are given below:

L1 L2 LЗ L4 H4 HЗ H2 Н1 Observed flood, cfs 10,300 10,600 12,400 24,500 829 1040 1090 1310 Modified, $NO_1 = 1$ 762 Modified, NO = 3754 982 12,677 15,481 Because of the detection of 2 higher inliers, H2 and H3, the 100-year flood of 24,849 cfs with NO = 3 is higher than 23,068 cfs with NO₁ = 1 with the MD method. Similar results are obtained for a 1000-year flood - 41,562 cfs with NO = 3 and 37,849 cfs with $NO_1 = 1$.

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		100-year floods in cfs										
		Wind	Window O		Window	5	Window 5				Window	5
No .	USGS No.	LP3	МО	NO	LP3	MD	NO1	LP3	MD	NO2	LP3	MD
	SANGAMON RI	VER BASIN										
1	05 571000	17,954	16,291	3	16 , 490	16 , 369	2	17 , 954	16 , 291	1	17.954	16 , 291
2	05 572000	20,277	20,066	5	20,075	19 , 874	1	20,277	20,066			
3	05 572500	25 , 630	21,888	2	28,136	24,557						
4	05 574000	11,499	10,419	2	11,273	10,224						
5	05 574500	13,468	16,192	3	13,871	16,294	4	14,016	16 , 253	2	14,521	16,186
6	05 575500	25,318	25,039	3	25,240	24,962						
7	05 576000	22,345	20,785	3	22,203	20,775						
8	05 576500	38 , 917	53,173	5	33,931	44,041	3	33 , 915	43,601	1	35 , 686	45,887
9	05 577500	10,125	9,866	3	10,943	10,037	4	10,820	9,796	2	10,853	10,181
-86-10	05 578500	28,363	23,058	3	30,311	24,849	1	28,202	23,068			
11	05 579500	7 , 531	9,860	3	10,188	9,458	2	10,126	9 , 525	1	10,019	9 , 637
12	05 580000	25 , 793	24,284	3	25,312	24,137						
13	05 580500	25,438	23,678	3	25 , 013	23,562	4	25 , 673	23,771	2	25,266	23,762
14	05 581500	26,343	29,663	3	25 , 286	26,829						
15	05 582000	49,963	45,259	3	49,251	44,884						
16	05 582500	863	692	3	903	752	2	902	754	1	878	745
17	05 583000	78,851	89 , 279	5	70,037	69 , 535	3	67 , 647	67 , 266	1	65 , 702	65 , 362
	ROCK RIVER	BASIN (in	Illinois)									
1	05 435500	22,209	18,266	5	22,394	18,996	3	22,370	19,025	1	22,429	19,026
2	05 437000	21,981	19,893	3	22,234	20,542	1	22,195	20,494			
3	05 437500	35,881	32,527	4	35,900	32,411	2	35,900	32,411			
4	05 438250	4,795	5,927	1	3,943	4,720	2	3,655	4,483	3	4,110	4,629
5	05 438500	16 , 647	11,505	4	18,027	14,240	2	17,606	13,831	1	17,075	12,708
6	05 439500	, 10,186	8,823	4	9,914	9,643	2	, 9,668	9,510		-	
7	05 440000	24,980	18,665	4	26,412	22,608	2	25,325	21,890	1	24,786	21,618
8	05 440500	7,748	7,325	4	7,688	7,647	2	7,692	7,644			·
9	05 441000	11 , 459	10,368	4	12,024	10,898	2	, 12,018	10,908			
10	05 443500	58,567	58,555	5	58,584	59,051	3	58,478	58,982			

Table 14. 100-Year Floods with Different Values of NO

		100-year floods in cfs										
		Windo	Window 0		Window 5			Window 5			Window 5	
No.	USGS No.	LP3	MO	NO	LP3	MD	NO	LP3	MD	NO	LP3	MD
	ROCK RIVER	BASIN (in	Illinois)	(Cont:	inued)							
11	05 444000	7,300	7,829	4	7,232	7 , 787						
12	05 445500	5,741	6 , 348	3	5,183	5 , 687	1	5,589	6,192			
13	05 446500	60 , 315	49,036	4	62,690	54,700	2	61 , 857	53 , 703			
14	05 447000	6,425	7,156	4	6,445	7,063	3	6 , 354	7,180	2	6,318	7,068
15	05 447500	12,189	13 , 305	4	11,907	12 , 838						
16	05 448000	13,431	12,428	4	13,198	12,165	2	13,861	12,943	1	13,861	12 , 943
	LITTLE WABA	SH RIVER E	BASIN									
7-1	03 379500	55 , 586	56 , 494	5	55 , 863	55 , 863	3	57 , 101	56 , 445			
2	03 380475	17 , 873	19 , 893	1	8,858	8,476	2	11,078	11 , 388			
3	03 380500	36,719	39,943	5	35,593	37,723	3	33,695	35 , 498	1	31,902	33 , 194
4	03 381500	46,614	54 , 678	4	50,145	50 , 593	2	52 , 576	51,232			

Another example is the Kishwaukee River near Belvidere in the Rock River Basin (No. 5 and USGS No. 05 438500). The relevant data with NO = 4, NO₁ = 2, and NO₂ = 1 are:

	L1	L2	L3	L4	H4	НЗ	H2	Н1
Observed flood, cfs	935	955	1070	1090	9040	9200	9830	10,300
Modified, NO = 4	647	921		1134		9965	11 , 294	13,694
Modified, $NO_1 = 2$	635	911					11,115	13,397
Modified, NO ₂ = 1	630							13,018

The 100-year floods with number of outliers/inliers equal to 4, 2, and 1 are 14,240, 13,831, and 12,708 cfs, respectively, with the MD method. It is evident that NO = 4 includes practically all outliers/inliers, and that NO = 3 would have given similar results as NO = 4.

LP3 and MD Statistics

The distribution parameters for LP3 and MD, before and after modification of outliers/inliers (at level or window 0 and 5, respectively) are given in Table 15 for the 37 basins. The LP3 distribution can simulate 3 shapes on lognormal probability paper - convex, straight, and concave for positive, zero, and negative skew, respectively - as shown in figure 14. It cannot simulate symmetrical distributions with kurtosis \neq 3 but the PT method with kurtosis correction is satisfactory in such cases. However, if the distribution is not symmetrical and if the cumulative distribution exhibits an S-curve shape, the mixed distribution method, MD, provides satisfactory flood estimates. The MD becomes a normal distribution when $\mu_1 = \mu_2$ and $\sigma_1 = \sigma_2$. Some of the diverse shapes that can be simulated or fitted by the MD are shown in figure 14. The dotted lines indicate the two component

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ZERO WINDOW									F	IFTH W	INDOW							
		Out-		LP3				MD				LP3				MD		
No.	USGS No.	lier*	x	8	8	a	μ1	¥2	σ1	°2	x	ŝ	g	а	۴ı	μ ₂	σι	σ2
	SANGAMON F	RIVER BA	SIN				-	_	•	-					-	-	-	-
1	05 571000		3.607	.285	073	.751	3.513	3.891	.249	.174	3.608	.285	070	.743	3.510	3.891	.248	.175
2	05 572000	LI	3.697	.291	312	.651	3.607	3.867	.295	.192	3.696	.292	336	.561	3.573	3.855	.291	.204
3	05 572500	HI	3.755	.250	.398	.378	3.523	3.896	.089	.207	3.758	.258	.486	.362	3.536	3.885	.097	.235
4	05 574000		3.026	.359	.788	.555	2.830	3.270	.186	.373	3.025	.358	.771	.536	2.827	3.254	.182	.374
5	05 574500	LO,LI	3.560	.329	803	.060	2.739	3.612	.196	.259	3.560	.325	717	.051	2.712	3.606	.135	.263
6	05 575500		3.630	.338	051	.729	3.559	3.822	.327	.290.	3.630	.338	054	.664	3.541	3.807	.324	.293
7	05 576000	LI	3.712	.320	453	.557	3.541	3.928	.302	.182'	3.711	.321	463	.439	3.464	3.904	.278	.195
8	05 576500	LO,HO	4.144	.312	-1.227	.281	3.904	4.238	.442	.166	4.144	.291	-1.386	.159	3.719	4.225	.379	.182
9	05 577500		3.201	.386	325	.548	3.024	3.417	.376	.270	3.203	.384	203	.870	3.145	3.592	.376	.129
10	05 578500	HI	3.571	.367	.104		3.240	3.809	.196	.262	3.574	.377	.110	.409	3.234	3.809	.205	.276
11	05 579500	LO	3.269	.341	730		1.900	3.287	.140	.305	3.285	.298	.140	.000	1.849	3.285	.273	.297
12	05 580000	LI	3.589	.310	.456		3.416	3.763	.200	.303	3.587	.311	.417	.558	3.429	3.786	.215	.298
13	05 580500	LI,HO	3.626	.303	.345		3.413	3.735	.179	.295	3.624	.304	.299	.323	3.415	3.724	.190	.298
14	05 581500	LO	3.682	.277	.465		3.656	4.428	.244	.078	3.682	.273	.438	.964	3.656	4.396	.239	.051
15	05 581500	ЦU	4.071	.278	099		3.772	4.205	.163	.207	4.071	.279	128	.298	3.765	4.200	.165	.206
16	05 582500	T T 11 T		.329	525		1.977	2.530	.226	.147	2.297	.347	579	.583	2.103	2.569	.320	.138
17	05 582500			.329	525		4.221	2.530	.226	.147	4.309	.347	764	.380	4.046	2.569	.320	.138
	ROCK RIVER	BASIN	(in Ill:	inois)														
1	05 435500	LI,HI	3.767	.251	024	.810	3.691	4.094	.211	.094	3.767	.255	047	.868	3.714	4.115	.229	.088
2	05 437000			.208	345	.284		4.019	.096	.127	3.910	.216	410	.332	3.670	4.030	.140	.130
3	05 437500		4.128	.200	258	.291	3.877	4.231	.091	.128	4.126	.206	329	.460	3.955	4,272	.150	.114
4	05 438250	, HO	3.026	.399	926		2.793	3.191	.484	.201	3.018		-1.133	.120	2.193	3.130	.234	.239
5	05 438500			.319	310		3.351	3.868	.237	.097	3.554	.336	318	.660	3.390	3.873	.285	.139
6	05 439500			.269	869	.404		3.731	.227	.099	3.554		-1.018	.438	3.344	3.717	.296	.106
7	05 440000	,	3.855	.282	541	.524		4.094	.215	.086	3.856	.301	601	.609	3.702	4.094	.285	.108
8	05 440500	HI	3.337	.349	-1.011	.386		3.527	.343	.145	3.336	.356		.398	3.042	3.531	.376	.147
9	05 441000	HI	3.385	.382	755	.440		3.624	.354	.180	3.388	.385	712		3.147	3.641	.377	.173
10	05 443500	HI	4.362	.233	789		4.217	4.493	.245	.113	4.362	.234	800		4.212	4.490	.248	.117
11	05 444000		3.468		-1.049		3.208	3.579	.274	.139	3.468		-1.048		3.208	3.576	.271	.139
12	05 444000	LO, HO LO	3.333	.235	-1.049	.300		3.392	.274	.046	3.330	.232	-1.048	.292	3.305	3.423	.201	.040
13	05 445500	LI,HI	4.341	.218	421		4.174	4.528	.158	.048	4.341	.229	453	.544	4.188	4.524	.189	.1040
																		.164
14	05 447000	LO	3.387		-1.147		2.732	3.470	.077	.169	3.398	.261	-1.016		2.864	3.477	.146	
15	05 447500	HO	3.740	.212	940		3.375	3.803	.180	.142	3.740	.207	956		3.507	3.818	.222	.129
16	05 448000	HI,HO	3.401	.348	317	.206	2.902	3.530	.172	.252	3.399	.347	336	.240	2.942	3.543	.199	.244
	LITTLE WAE	ASH RIV	ER BASIN	1														
1	03 379500	LI,HI,H	HO 4.088	.315	324	.435	3.924	4.214	.322	.244	4.086	.316	318	.490	3.932	4.235	.315	.236
2	03 380475	HO	3.596	.231	.729	.489	3.533	3.657	.005	.311	3.579	.188	493	.303	3.726	3.514	.069	.187
3	03 380500	LI,HO,H	HI 3.886	.334	394	.108	3.454	3.939	.357	.290	3.887	.336	467	.612	3.781	4.053	.364	.190
4	03 381500	L0,L1,1	H04.147	.234	126	.519	4.131	4.164	.293	.143	4.151	.216	.306	.633	4.199	4.067	.235	.143

* LI, LO, HI, and HO denote low inlier, low outlier, high inlier, and high outlier respectively.

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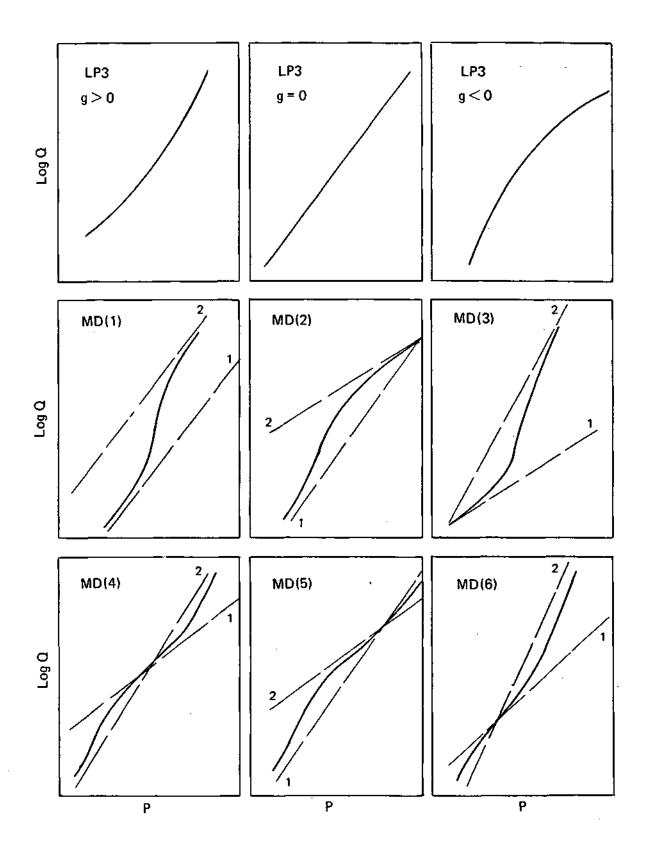


Figure 14. Probability curves for LP3 and MD

distributions and the solid curve the mixed distribution. The S-curve shapes can, to some extent, be caused or accentuated by the existence of outliers/inliers.

LP3 Statistics

It is generally felt that modification of any outliers/inliers in an observed flood series would change skew significantly. However, with the exception of 3 basins, the values of skew obtained without any modification of outliers/inliers and with modification of outliers/inliers detected up to level 5 or in window 5 are not much different from each other for the remaining 34 basins. The exceptions are: 05 579500, Lake Fork near Cornland, with a very low outlier, g changes from -0.730 in window 0 to 0.140 in window 5; 03 380475, Horse Creek near Keenes, a very high outlier, g changes from 0.729 in window 0 to -0.493 in window 5; and 03 381500, Little Wabash River at Carmi, with low outlier and inlier and a high outlier, g changes from -0.126 in window 0 to 0.306 in window 5.

Modification of high outlier(s) and/or low inlier(s) has the effect of making the skew value smaller in the algebraic sense, and the modification of high inlier(s) and/or low outlier(s) makes the skew value larger. The change in skew from window 0 to 5 can be explained generally by the type of outliers/inliers.

According to the U.S. Water Resources Council Bulletin 17, the regional skew for the 37 basins analyzed is about -0.4. The number of basins with lower and higher skew and the minimum and maximum values of skew in a major river basin are given on the next page (from information for window 5 in Table 15).

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River Basin	Number of sub-basi g < -0.4 g	lns with > -0.4	Min g	Max g
Sangamon	5	12	-1.386	0.771
Rock	11	5	-1.133	-0.047
Little Wabash	2	2	-0.493	0.306
Even after modification	of outliers/inliers,	the deri	ved values	of skew do
not indicate a regional	skew of -0.4. Value	es of skew	for basin	s with more

than 60 years of record are:

USGS No.	n	g(window 0)	g(window 5)
05 572000	71	-0.312	-0.336
05 576500	67	-1.227	-1.386
05 583000	62	-0.562	-0.764
05 435500	66	-0.024	-0.047
05 443500	65	-0.789	-0.800
03 379500	65	-0.324	-0.318

Again, a regional skew value of -0.4 is not indicated by the above 6 long-term stations.

MD Statistics

The mixed distribution has two component distributions, the parameters of mean and standard deviation carry subscripts 1 and 2, and the weight of the first distribution is given by *a*. The mean, μ_1 , of the first component distribution is smaller than μ_2 for the second distribution. The general shape of mixed distribution can be categorized from figure 14 with the relative values of σ_1 and σ_2 . Some distortions in these shapes can be caused by unequal weight of the two component distributions. A brief summary

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of the MD statistics is given below:

Item	Number of	sub-basins	in the basin of				
	Sangamon	Rock	Little Wabash				
a > .5	*9(8)	5(6)	1(2)				
a < .5	8(9)	11(10)	3(2)				
$^{\mu}1 < ^{\mu}2$	17(17)	16(16)	4(2)				
$\mu_1 > \mu_2$	0(0)	0(0)	0(2)				
$\sigma_1 > \sigma_2$	9(9)	12(13)	3 (3)				
$\sigma_1 < \sigma_2$	8(8)	4(3)	1(1)				

 \star for window 0 and in parentheses for window 5.

The flood series with $\sigma_1 > \sigma_2$ will have shapes similar to MD(2) and MD(5) and with $\sigma_1 < \sigma_2$ will resemble MD(3) and MD(6) in figure 14. The probability curves are affected largely by $\Delta \mu$ (or $\mu_1 - \mu_2$), $\Delta \sigma$ (or $\sigma_1 - \sigma_2$), and *a*. A few points of interest regarding the mixed distribution and kurtosis are

1. With
$$\mu_1 = \mu_2$$
 and $a = 0.5$

$$\mu = 0.5 \ (\mu_1 + \mu_2) = \mu_1 = \mu_2 \tag{1}$$

$$\sigma^2 = 0.5 \ (\sigma_1^2 + \sigma_2^2) \tag{2}$$

$$g = 0 \tag{3}$$

$$kt = 1.5 \ (\sigma_1^4 + \sigma_2^4) \ / \ \sigma^4 \tag{4}$$

in which kt is the kurtosis. The simulated distributions are symmetrical with kurtosis = 3 with σ_1/σ_2 or σ_2/σ_1 = 1, kt = 4.92 for a ratio of 3 or 1/3, and kt = 6.0 for a ratio of infinity or zero.

2. With $\sigma_1 = \sigma_2$ and a = 0.5 $\mu = 0.5 \ (\mu_1 + \mu_2)$ (5)

$$\sigma^{2} = 0.5(\sigma_{1}^{2} + \sigma_{2}^{2}) + 0.25(\Delta \mu)^{2} = \sigma_{1}^{2} + 0.25(\Delta \mu)^{2}$$
(6)

$$g = 0 \tag{7}$$

$$k = [12 \sigma_1^4 + 6 \sigma_1^2 (\Delta \mu)^4] / 4\sigma^4$$
(8)

in which σ_2 is replaced by σ_1 . The simulated distributions are symmetrical and yield kurtosis = ≤ 3 depending on the ratio of $\Delta \mu / \sigma_1$, keeping the mixed distribution unimodal.

The probability distributions for cases 1 and 2 correspond to MD(4) and MD(1), respectively, in figure 14. Only in these special cases, the kurtosis correction with the PT method may yield flood estimates comparable to those from the MD. However, the asymmetry of the observed flood distributions for the 37 study basins, in terms of mixed distribution parameters varying from those for cases 1 and 2, indicates that flood estimates from the MD will be better than from the PT with kurtosis correction (based on the assumption of symmetrical distribution).

Ratios of Q_{100}/Q_2 and Q_{1000}/Q_2

Ratios of a 100-year flood, Q_{100} , to a 2-year flood, Q_2 , and 1000-year flood, Q_{1000} , to Q_2 both for windows 0 and 5 and with LP3 sample skew and MD methods are given in Table 16. These ratios are plotted with respect to drainage area on a log-log paper in figures 15 and 16 for the Sangamon and in figures 17 and 18 for the Rock River basins for drainage areas exceeding 100 square miles.

Ratios for the Sangamon Basin

The ratios $Q_{100}^{\prime}/Q_2^{\prime}$ from LP3 in figure 15 indicate considerable scatter whereas they lie along two trend curves (one for the Salt Creek basins and

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Table 16. Ratios Q_{100}/Q_2 and Q_{1000}/Q_2

				Q ₁₀₀	/Q2		Q ₁₀₀₀ /Q ₂					
No.	USGS No.	A	<u>Wind</u> LP3	ow 0 MD	<u>Wind</u> LP3	iow 5 MD	<u>Wind</u> LP3	low 0 MD	<u>Wind</u> LP3	<u>ow 5</u> MD		
SANG	GAMON RIVE	R BASIN							-			
1	05 57100	00 362	4.47	4.01	4.42	4.03	7.03	5.74	7.07	5.76		
2	05 57200	0 550	3.93	3.83	3.89	3.79	5.71	5.99	5.60	5.74		
3	05 57250	0 774	4.68	4.13	5.15	4.74	8.56	6.09	9.99	7.35		
4	05 57400	00 11	12.07	11.65	11.82	11.46	36.65	23.90	35.37	23.40		
5	05 57450	0 276	3.35	4.15	3.50	4.21	4.16	6.57	4.46	6.71		
6	05 47550	0 562	5.89	5.83	5.88	5.80	10.39	10.13	10.34	10.07		
7	05 57600	0 867	4.10	3.65	4.09	3.65	5.81	5.36	5.76	5.25		
8	05 57650	0 2618	2.42	3.37	2.10	2.84	2.64	7.85	2.21	4.06		
9	05 57750	0 107	6.07	5.79	6.65	6.10	9.89	9.85	11.58	11.88		
10	05 57850	0 335	7.73	6.28	8.21	6.73	15.66	10.23	17.04	11.26		
11	05 57950	0 214	3.69	5.16	5.37	4.90	4.74	8.82	9.70	8.27		
12	05 58000	0 227	7.02	6.80	6.89	6.71	15.31	12.09	14.74	11.92		
13	05 58050	0 306	6.26	5.97	6.15	5.90	12.65	10.27	12.20	10.20		
14	05 58150	0 333	5.75	6.39	5.50	5.77	11.61	8.14	10.83	6.78		
15	05 58200	1804	4.19	3.68	4.13	3.64	6.55	5.38	6.38	5.30		
16	05 58250	26.5	4.05	2.84	4.22	3.17	5.61	3.73	5.81	4.74		
17	05 58300	0 5093	3.61	3.81	3.14	2.97	4.82	7.10	3.89	4.12		
ROCK	K RIVER BA	SIN										
1	05 43550	0 1326	3.79	3.22	3.82	3.32	5.83	4.07	5.86	4.53		
2	05 43700		2.62	2.22	2.64	2.32	3.39	2.79	3.39	2.95		
3	05 43750		2.62	2.24	2.61	2.27	3.44	2.82	3.39	2.81		
4	05 43825	50 85.1	3.93	4.67	3.21	3.84	4.84	11.33	3.66	5.87		
5	05 43850	0 538	4.49	3.12	4.83	3.74	6.76	3.81	7.41	5.12		
6	05 43950	0 387	2.60	1.98	2.49	2.21	3.04	2.41	2.81	3.49		
7	05 44000		3.29	2.14	3.44	2.68	4.33	2.55	4.50	4.13		
8	05 44050	0 117	3.12	2.65	3.08	2.78	3.65	3.75	3.55	4.60		
9	05 44100	103	4.23	3.44	4.44	3.63	5.55	5.00	5.94	5.89		
10	05 44350	0 8755	2.38	2.27	2.37	2.29	2.78	3.23	2.77	3.26		
11	05 44400		2.25	2.35	2.23	2.35	2.50	3.14	2.47	3.12		
12	05 44550		2.63	2.79	2.36	2.46	3.48	4.14	2.97	3.54		
13	05 44650		2.65	2.00	2.75	2.27	3.39	2.33	3.52	2.78		
14	05 44700		2.33	2.58	2.33	2.53	2.56	3.49	2.62	3.38		
15	05 44750		2.06	2.24	2.00	2.15	2.29	2.89	2.23	2.74		
16	05 44800	62.4	5.12	4.44	5.03	4.36	7.98	7.00	7.77	6.79		

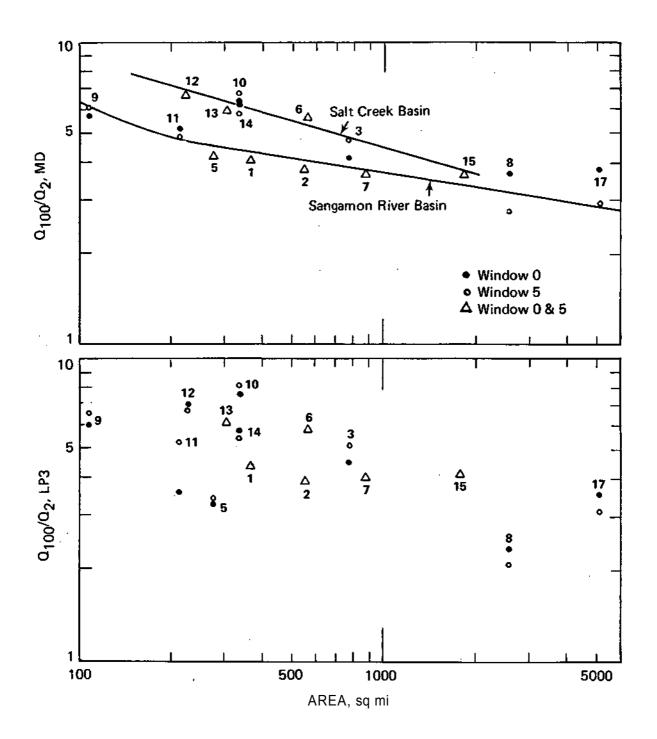


Figure 15. Q_{100}/Q_2 versus drainage area, Sangamon River Basin

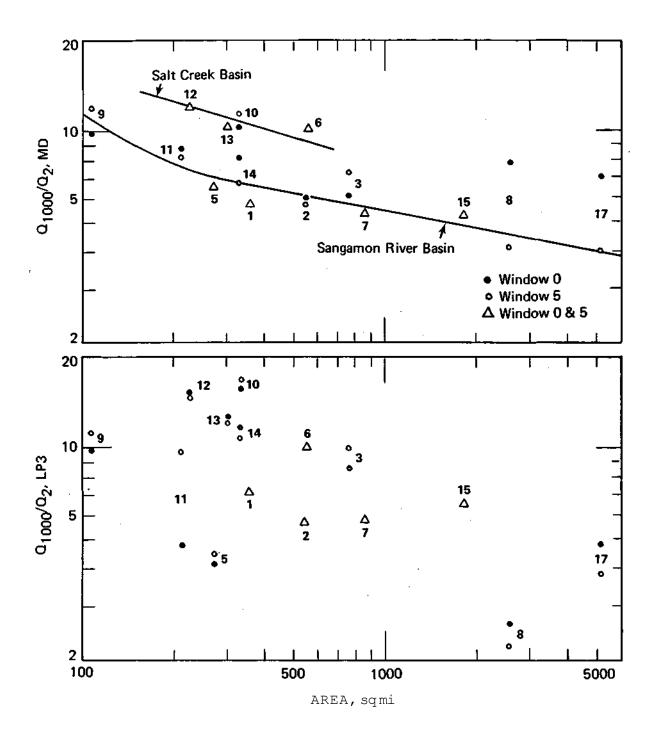


Figure 16. $Q_{\rm 1000}/Q_{\rm 2}$ versus drainage area, Sangamon River Basin

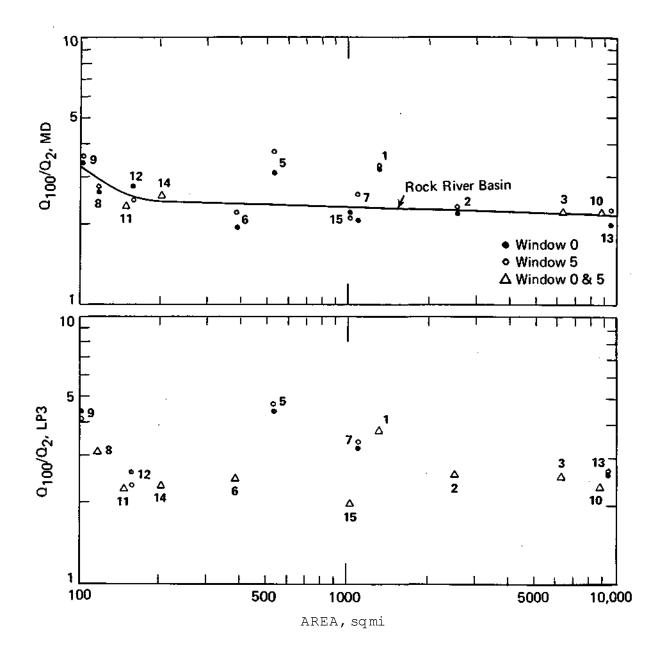


Figure 17. $Q_{100}/Q_2\,$ versus drainage area, Rock River Basin

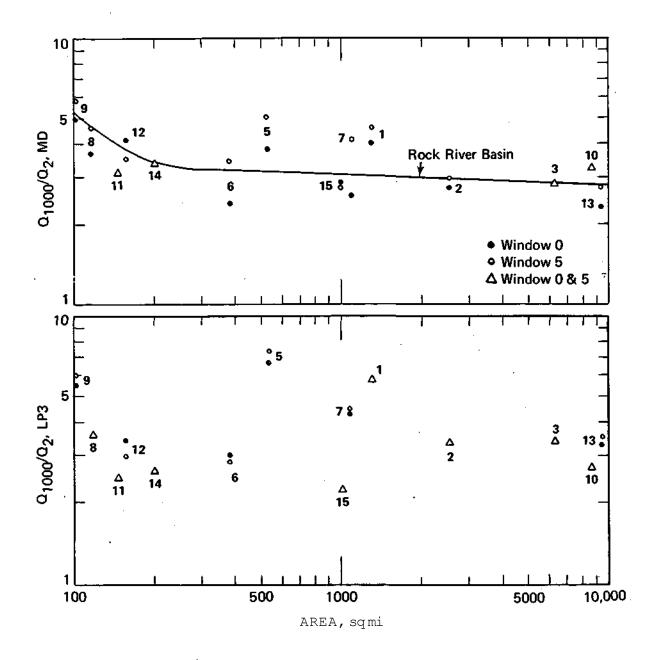


Figure 18. Q_{1000/Q_2} versus drainage area, Rock River Basin

the other for the Sangamon River basins) for the MD and window 5. The curves represent a decrease in the value of the ratio with increase in drainage area. Similar results are shown by the Q_{1000}/Q_2 plots in figure 16. The trend curves steepen as the drainage area becomes less than 200 square miles.

Ratios for the Book Basin

The ratios of Q_{100}/Q_2 and Q_{1000}/Q_2 , in figures 17 and 18, show that the MD method with window 5 indicates satisfactorily the decrease in these ratios with drainage area. However, the decrease is much smaller than in the Sangamon Basin. The trend curves steepen as area decreases below 200 square miles. The ratios with LP3 show considerable scatter.

Some Specific Examples

Examples of various types of outliers/inliers are discussed. The fitted probability curves with the MD and window 5, and computer output tables showing the modification of outliers/inliers from one window to the next are used to explain each example.

1. Lake Fork near Cornland: Low Outlier

The results obtained with the computer program are given in Table 17. It contains the 100-year flood estimates from five methods; observed and modified floods for NO points; statistics with PT, LP3, and MD methods for all the windows 0 through 6; and 2- to 1,000-year floods with the five methods for all the windows. The three high floods are not perceived as outliers or inliers up to window 5. The lowest flood is perceived as a significant low outlier but the next two floods as rather insignificant low outliers. The

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STATION DRAINAG				NEAR CORN s of Reco		(1948-19	79)				
L	EVEL NO.	0		12	3	4	5	6			
	ETHOD		100-Year Flood in cfs								
	ransform, PT		0 700	0 700	0 070	0 (20	10 077	11 510			
	t = 3.0			8,736			10,277	11,510			
Log Tra		9,838	9,838	9,838	9,942	10,361	10,624	11,035			
	Sample skew			7,531		9,625	10,188	10,666			
	Veighted skew					8,209					
Mixed	Distrib., MD	9,860	9,860	9,860	9,878	9,820	9,458	8,951			
Ту	rpe No.		Obser	ved and M	lodified	Floods i	n cfs				
Lo	w 1*	152	152	152	183	336	443	574			
	2*	548	548	548	548	548		712			
	3*	680	680	680	680	680	704	812			
	4	1,000									
	5	1,010									
Hi	gh 5	4,570									
	4	4,700									
	3*	4,900	4,900	4,900	4,900	4,900	4,900	4,936			
	2*	6,120	6,120	6,120	6,120	6,120	6,120	6,120			
	1*	8,930	8,930	8,930	8,930	8,930	8,930	8,876			
METHOD	STATISTICS			Values	of Stat	istics					
PΤ	mean	19.126	19.126	19.126	16.273	8.519	5.803	3.560			
	std dev	3.848			2.941	.899	.394	.111			
	skew	.036	.036	.036	.026	.001		.028			
	kurtosis,kt		4.045	4.045	3.932	3.433					
	5th moment				-1.448	919					
	lambda	.215	.215	.215	.181	.031	073	232			
LP3	mean	3.269	3.269	3.269	3.271	3.280	3.285	3.293			
	std dev	.341	.341	.341	.333	.309	.298	.283			
	sample skew	730	730	730	560	068	.140	.370			
	kurtosis,kt	5.438	5.438	5.438	4.852	3.477	3.136	2.933			
	5th moment	-12.793	-12.793	-12.793	-9.442	-1.618	.614	2.507			
MD	weight 'a'	.013	.013	.013	.010	.001	.000	0.000			
	mu1	1.900	1.900	1.900	1.967	1.849	1.849	1.849			
	mu2	3.287	3.287	3.287	3.285	3.281	3.285	3.293			
	sigma 1	.140	.140	.140	.171	.273	.273	.273			
	sigma 2	.305	.305	.305	.305	.306	.297	.283			
	Test Stat	3.258	3.258	3.258	3.342	.001	.004	.009			
* F	iah & low fl	oods con	sidered	for outl	ier deter	ction and	modific	ation			

 \star High & low floods considered for outlier detection and modification

STATION NO. 5579500 LAKE FORK NEAR CORNLAND DRAINAGE AREA 214.0 Sq Mi Years of Record 32 (1948-1979) VARIOUS RECURRENCE-INTERVAL FLOODS METHOD # Flood in cfs for Recurrence Intervals (Years) 2 10 25 50 100 500 1000 4,748 6,306 8,736 11,703 13,033 PT, kt=3.0 1,976 7,508 0 PT, sample kt 1,976 4,579 6,428 8,036 9,838 14,824 17,334 LP3, sample skew 2,042 4,681 5,915 6,758 7,531 9,092 9,676 weighted skew 1,965 4,862 6,486 7,722 8,964 11,864 13,114 MD, mixed dist. 9,860 14,527 16,866 1,912 4,731 8,139 6,580 PT, kt=3.0 1 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 2 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 3 1,965 4,742 6,340 7,589 8,878 12,039 13,475 PT, sample kt 1,965 4,587 6,460 9,942 15,096 17,719 8,097 2,006 LP3, sample skew 4,707 6,095 7,099 8,068 10,187 11,046 weighted skew 1,970 4,788 6,362 7,559 8,763 11,578 12,796 MD, mixed dist. 1,911 4,729 6,582 9,878 14,567 16,916 8,148 ΡT, kt=3.0 4 1,919 4,716 6,514 8,011 9,639 13,974 16,099 1,919 8,347 10,361 16,184 19,274 PT, sample kt 4,633 6,600 LP3, sample skew 1,920 4,716 6,511 8,004 9,625 13,932 16,040 weighted skew 1,989 4,589 6,023 7,112 8,209 10,787 11,908 MD, mixed dist. 1,911 4,707 6,547 8,106 9,820 14,478 16,810 PT, kt=3.0 5 1,896 4,688 6,632 8,337 10,277 15,854 18,798 6,672 1,896 4,654 10,624 16,994 20,485 PT, sample kt 8,496 LP3, sample skew 1,897 4,685 6,611 8,290 10,188 15,583 18,397 weighted skew 2,006 4,504 5,871 6,907 7,950 10,400 11,468 MD, mixed dist. 1,929 4,632 6,383 7,852 9,458 13,790 15,945 6 PT, kt=3.0 1,873 4,630 6,806 8,903 11,510 20,391 25,966 PT, sample kt 8,697 11,035 18,563 23,023 1,873 4,663 6,755 LP3, sample skew 1,886 4,628 8,500 10,666 17,232 20,872 6,655 4,412 6,666 7,641 9,927 10,923 weighted skew 2,034 5,696 MD, mixed dist. 7,493 8,951 12,826 14,730 1,963 4,528 6,149

= level number

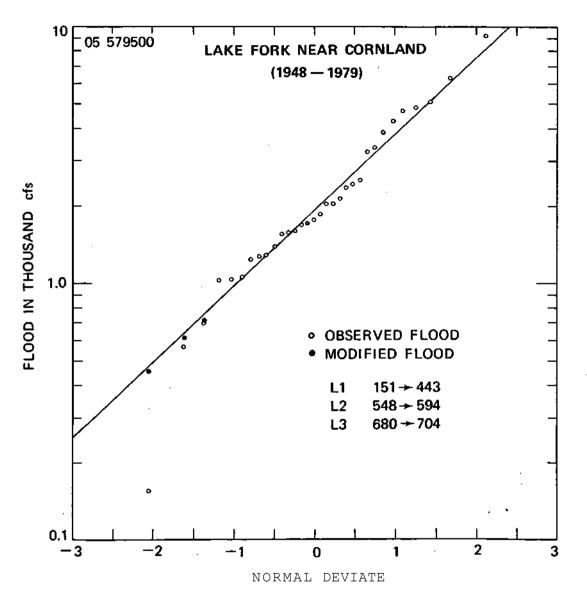


Figure 19. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Lake Fork near Cornland

three lowest floods of 152, 548, and 680 cfs are modified to 443, 594, and 704 cfs, respectively, in window 5. The modification of low outliers changes Q_{100} and Q_{1000} estimates of 7,531 and 9,676 cfs in window 0 to 10,188 and 18,397 in window 5 (because of sample skew changing from -0.730 to 0.140) with the LP3 and sample skew. The mean and standard deviation change from 3.269 and 0.341 to 3.285 and 0.298.

The PT statistics show that the power transformation reduces skew close to zero and the kurtosis decreases from 4.045 in window 0 to 3.162 in window 5; the kurtosis for a theoretical normal distribution is 3.0. Flood estimates with the PT are higher/lower with sample kurtosis than with kt=3.0 if sample kurtosis is higher/lower than 3.0. The PT 100-year flood estimate with kt=3.0 increases from 8,736 in window 0 to 10,277 in window 5. With sample kt, it increases from 9,838 cfs to 10,624 cfs.

The MD flood estimates are rather insensitive to modification of low outliers. The 100-year flood changes from 9,860 cfs in window 0 to 9,458 cfs in window 5 and a 1000-year flood changes from 16,866 cfs to 15,945 cfs. The MD statistics show that the effect of the first component distribution is negligible, the weight being a maximum of 0.013, and that the distribution is practically normal (which is indicated by the LP3 in between windows 4 and 5). The MD method seems to be the best for analyzing observed flood series with low outliers. The observed floods as well as the modified low floods in the 5th window and the probability curve fitted by the MD method are shown in figure 19.

2. Rock River at Rockton: Low and High Inliers

The results obtained with the computer program are given in Table 18.

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STATION NO. 5437500 ROCK RIVER AT ROCKTON DRAINAGE AREA 6363.0 Sq Mi Years of Record 40 (1940-1979) 3 5 6 LEVEL NO. 0 12 4 METHOD 100-Year Flood in cfs Power Transform, PT With kt = 3.034,432 34,432 34,432 34,432 34,465 34,544 35,249 30,933 30,933 30,933 30,933 31,137 31,740 33,100 With sample kt Log Transform 36,330 35,881 35,881 35,881 35,881 35,910 35,900 LP3, Sample skew LP3, Weighted skew 34,512 34,512 34,512 34,512 34,698 35,185 35,997 Mixed Distrib., MD 32,527 32,527 32,527 32,527 32,461 32,411 33,772 Observed and Modified Floods in cfs No. Type 4,258 Low 1* 5,400 5,400 5,400 5,400 5,183 4,692 2* 6,340 6,340 6,340 6,340 6,222 5,821 5,451 3* 6,340 6,340' 6,340 6,340 6,340 6,340 6,267 4* 6,880 6,880 6,880 6,880 6,880 6,880 6,880 5 7,450 23,800 5 High 4* 24,300 24,300 24,300 24,300 24,300 24,300 24,300 3* 25,400 25,400 25,400 25,400 25,400 25,400 25,874 2* 25,700 25,700 25,700 25,700 26,247 27,048 28,076 1* 30,000 30,000 30,000 30,000 30,434 31,950 30,000 METHOD STATISTICS Values of Statistics PΤ 58.809 58.809 58.809 58.809 63.875 79.332 79.342 mean std dev 8.686 8.686 8.686 8.686 9.801 13.388 13.733 skew -.051 -.051 -.051 -.051 -.052 -.054 -.047 kurtosis, kt 2.190 2.190 2.190 2.190 2.218 2.310 2.459 -.007 5th moment -.007 -.007 -.007 -.047 -.120 -.095 .310 .353 lambda .310 .310 .310 .322 .353 4.128 4.128 4.128 4.128 4.128 4.126 4.126 LP3 mean .200 .200 std dev .200 .200 .201 .206 .212 sample skew -.258 -.258 -.258 -.368 -.258 -.276 -.329 kurtosis, kt 2.211 2.211 2.211 2.211 2.259 2.414 2.617 5th moment -1.187 -1.187 -1.187 -1.187 -1.365 -1.921 -2.460 .291 .291 .291 weight 'a' .291 .485 MD .329 .460 mu1 3.877 3.877 3.877 3.955 3.972 3.877 3.896 mu2 4.231 4.231 4.231 4.231 4.242 4.272 4.270 sigmal .091 .091 .091 .091 .108 .150 .173 .128 .128 sigma2 .128 .128 .124 .114 .124 2.827 2.827 2.827 2.827 2.688 2.575 2.370 Test Stat

STATION NO. 5437500 ROCK RIVER AT ROCKTON DRAINAGE AREA 6363.0 Sq Mi Years of Record 40 (1940-1979) VARIOUS RECURRENCE-INTERVAL FLOODS METHOD # Flood in cfs for Recurrence Intervals (Years) 2 10 25 50 100 500 1000 PT, kt=3.0 13,866 23,613 28,125 31,335 34,432 41,335 44,224 0 13,866 24,099 27,333 29,272 30,933 PT, sample kt 34,121 35,304 LP3, sample skew 13,703 23,886 28,818 32,392 47,115 35,881 43,774 weighted skew 13,822 23,716 28,264 31,464 34,512 41,152 43,859 14,543 23,392 27,184 29,892 32,527 38,517 41,082 MD, mixed dist. PT, kt=3.0 1 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 2 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 3 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. 41,331 PT, kt=3.0 4 13,876 23,662 28,174 31,379 34,465 44,199 PT, sample kt 13,876 24,128 27,426 29,419 31,137 34,456 35,692 LP3, sample skew 13,710 23,947 28,880 32,443 35,910 43,720 47,010 weighted skew 13,816 23,796 28,388 31,619 34,698 41,402 44,134 MD, mixed dist. 14,487 23,489 27,231 29,890 32,461 38,286 40,768 PT, kt=3.0 5 13,898 23,783 28,295 31,484 34,544 41,312 44,125 13,898 24,195 PT, sample kt 27,677 29,842 31,740 35,490 36,899 LP3, sample skew 13,732 24,096 29,006 32,515 35,900 43,412 46,530 32,028 weighted skew 13,794 24,005 28,714 35,185 42,054 44,850 MD, mixed dist. 14,296 23,747 27,415 29,968 32,411 37,869 40,167 PT, kt=3.0 13,903 24,091 6 28,763 32,071 35,249 42,291 45,221 PT, sample kt 13,903 24,413 28,314 30,829 37,720 39,512 33,100 LP3, sample skew 13,762 24,392 29,385 32,930 36,330 43,803 46,875 weighted skew 13,791.24,350 29,249 32,703 35,997 43,170 46,091 MD, mixed dist. 14,189 24,067 28,137 31,002 33,772 40,026 42,688

The NO equals [40/10] or 4. However, in going from window 0 to 5, only two low inliers, L1 and L2, are detected and these are modified from their original values of 5,400 and 6,340 cfs to 4,692 and 5,821 cfs, respectively. Two high inliers, H1 and H2, are detected and are modified from 30,000 and 25,700 cfs to 30,434 and 27,048 cfs. Generally, the modification of low inliers should decrease the skew and of high inliers should increase the skew. With the LP3, the skew decreases from -0.258 in window 0 to -0.329 in window 5. Because of a small change in skew (as well as mean and standard deviation), the 100- and 1000-year floods of 35,881 and 47,115 cfs in window 0 change to 35,900 and 46,530 cfs in window 5, with LP3 and sample skew.

The PT statistics show that the power transformation reduces skew close to zero but the kurtosis changes from 2.190 in window 0 to 2.310 in window 5. Thus, the flood estimates with sample kurtosis are considerably smaller than with kt=3.0. However, the 100- and 1000-year flood estimates with kt=3.0 are close to those with LP3 and sample skew.

The MD flood estimates are rather insensitive to modification of inliers as are the estimates with the PT and LP3. The 100- and 1000-year floods change from 32,527 and 41,082 cfs in window 0 to 32,411 and 40,167 cfs in window 5. Flood estimates from the five methods are summarized below:

	100-year window 0	flood, cfs window 5	1000-year fl window 0	ood, cfs. window 5
PT, kt=3.0	34,432	34,544	44,224	44,125
sample kt	30,933	31,740	35,304	36,899
LP3, sample g	35,881	35,900	47,115	46,530
weighted g	34,512	35,185	43,859	44,850
MD	32,527	32,411	41,082	40,167

The MD statistics in window 5 show that neither $\mu_1 = \mu_2$ nor $\sigma_1 = \sigma_2$.

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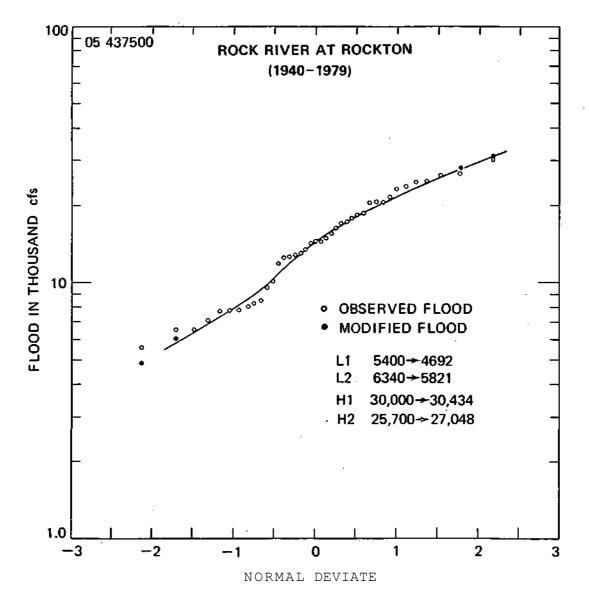


Figure 20. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Rock River at Rockton

Thus, the flood series is asymmetrical. Plots in figures 17 and 18 show that MD estimates for this basin lie on the well-defined regional curve for Q_{100}/Q_2 and Q_{1000}/Q_2 . Thus, the flood estimates with the MD are considered better than with other methods. The observed floods as well as the modified floods in the 5th window and the probability curve fitted by the MD method are shown in figure 20.

3. Flat Branch near Taylorville: Low Outliers and Inlier

The results obtained with the computer program are given in Table 19. The NO equals [30/10] or 3. In going from window 0 to 5, two low ouliers, L1 and L2, and a low inlier, L3, are detected and modified from their original values of 457, 660, and 1,770 cfs to 460, 841, and 1,370 cfs, respectively. Only one high oulier, H2, is detected in window 5 and it is modified from 11,300 cfs to 11,032 cfs which is relatively an insignificant modification. The LP3 sample skew changes from -0.803 in window 0 to -0.717 in window 5, the standard deviation from 0.329 to 0.325, and the mean remains unchanged. Because of a slight increase in skew (in the algebraic sense), the 100- and 1000-year floods are 13,468 and 16,691 cfs in window 0 and 13,871 and 17,697 cfs in window 5. In this example, the effects of low outlier and inlier practically balance each other.

The PT statistics indicate a skew very close to zero and a kurtosis of 3.256 in window 0 and 3.154 in window 5. Accordingly, the flood estimates with the sample kt are somewhat higher than with kt = 3.0, and the high flood estimates are higher than those with the LP3 and sample skew.

The MD statistics indicate a weight of only 0.06 to 0.05 for the first component distribution with a mean of 2.739 which is much smaller than

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STATION NO. 5574500 FLAT BRANCH NEAR TAYLORVILLE 276.0 Sq Mi Years of Record 30 (1950-1979) DRAINAGE AREA 2 0 1 3 4 5 6 LEVEL NO. METHOD 100-Year Flood in cfs Power Transform, PT With kt = 3.014,623 14,688 14,564 14,564 14,599 14,643 15,015 With sample kt 15,005 15,002 15,002 15,007 15,020 14,921 14,816 Log Transform LP3, Sample skew 13,468 13,468 13,549 13,616 13,821 13,871 15,148 16,725 16,575 LP3, Weighted skew 16,633 16,633 16,679 16,351 15,151 Mixed Distrib., MD 16,192 16,192 16,219 16,228 16,379 16,294 15,948 Observed and Modified Floods in cfs No. Type 1* 457 695 Low 457 457 457 457 460 2* 660 1,054 660 660 660 745 841 3* 1,770 1,770 1,649 1,551 1,440 1,370 1,324 4 1,850 5 1,860 High 5 7,540 8,620 4 3* 9,400 9,400 9,400 9,400 9,400 9,400 9,267 2* 11,300 11,300 11,300 11,300 11,300 11,032 10,612 1* 13,000 13,000 13,000 13,000 13,000 13,000 13,000 METHOD STATISTICS Values of Statistics 40.590 23.845 45.450 45.450 45.152 45.127 42.291 PTmean std dev 11.452 11.452 11.389 11.417 10.400 9.748 4.281 skew -.016 -.016 -.019 -.020 -.020 -.022 -.025 3.235 3.154 kurtosis, kt 3.256 3.256 3.218 3.185 2.908 -.558 -.543 -.519 -.500 -.513 -.237 5th moment -.558 lambda .325 .223 .337 .337 .336 .336 .318 3.559 3.568 LP3 mean 3.560 3.560 3.558 3.559 3.560 std dev .329 .329 .330 .331 .328 .325 .302 sample skew -.803 -.803 -.796 -.790 -.747 -.717 -.400 kurtosis, kt 4.302 4.252 4.208 4.108 4.302 4.046 3.159 5th moment -8.841 -8.841 -8.647 -8.479 -8.072 -7.872 -3.496 weight 'a' .060 .060 .061 .063 .053 .051 .055 MD 2.868 mu1 2.739 2.739 2.743 2.747 2.699 2.712 mu2 3.612 3.612 3.612 3.612 3.607 3.606 3.608 .196 .196 .005 sigma .194 .192 .137 .135 1 sigma2 .259 .259 .260 .260 .263 .263 .258 Test Stat 2.483 2.483 2.387 2.310 2.187 2.150 3.145

STATION NO. 5574500 FLAT BRANCH NEAR TAYLORVILLE DRAINAGE AREA 276.0 Sq Mi. Years of Record 30 (1950-1979)

VARIOUS RECURRENCE-INTERVAL FLOODS

METHOD #	Flc	od in c	fs for F	ecurrenc	e Interv	als (Yea	ars)
	2	10	25	50	100	500	1000
PT, kt=3.0 0	3,966	8,700	11,087	12,839	14,564	18,512	20,200
PT, sample kt	3,966	8,629	11,153	13,069	15,002	19,570	21,577
LP3, sample skew	4,015	8,777	10,865	12,240	13,468	15,842	16,691
weighted skew	3,833	9,199	12,158	14,395	16,633	21,827	24,057
MD, mixed dist.	3,902	8,617	11,448	13,743	16,192	22,570	25,627
PT, kt=3.0 1 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.			SAME	AS ABOV	E		
PT, kt=3.0 2	3,958	8,706	11,103	12,865	14,599	18,571	20,270
PT, sample kt	3,958	8,640	11,166	13,079	15,007	19,555	21,551
LP3, sample skew	4,004	8,790	10,902	12,299	13,549	15,976	16,848
weighted skew	3,824	9,205	12,178	14,427	16,679	21,910	24,158
MD, mixed dist.	3,901	8,618	11,454	13,759	16,219	22,605	25,683
PT, kt=3.0 3	3,952	8,710	11,115	12,882	14,623	18,609	20,315
PT, sample kt	3,952	8,649	11,175	13,082	15,005	19,527	21,510
LP3, sample skew	3,994	8,802	10,934	12,348	13,616	16,087	16,978
weighted skew	3,817	9,212	12,199	14,460	16,725	21,991	24,255
MD, mixed dist.	3,897	8,623	11,463	13,769	16,228	22,630	25,703
PT, kt=3.0 4	3,943	8,710	11,135	12,923	14,688	18,745	20,487
PT, sample kt	3,943	8,656	11,187	13,097	15,020	19,544	21,527
LP3, sample skew	3,976	8,805	10,999	12,477	13,821	16,496	17,483
weighted skew	3,820	9,161	12,110	14,341	16,575	21,766	23,998
MD, mixed dist.	3,874	8,631	11,514	13,864	16,379	22,938	26,119
PT, kt=3.0 5	3,937	8,675	11,094	12,879	14,643	18,707	20,454
PT, sample kt	3,937	8,631	11,137	13,025	14,921	19,376	21,324
LP3, sample skew	3,965	8,770	10,985	12,491	13,871	16,656	17,697
weighted skew	3,823	9,088	11,981	14,166	16,351	21,422	23,599
MD, mixed dist.	3,868	8,604	11,467	13,798	16,294	22,810	25,962
PT, kt=3.0 6 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	3,889 3,889 3,871 3,871 3,871 3,886	8,602 8,632 8,698 8,698 8,518		13,069 12,964 13,221 13,223 13,549	15,015 14,816 15,148 15,151 15,948	19,639 19,156 19,590 19,597 22,183	21,683 21,051 21,490 21,498 25,193

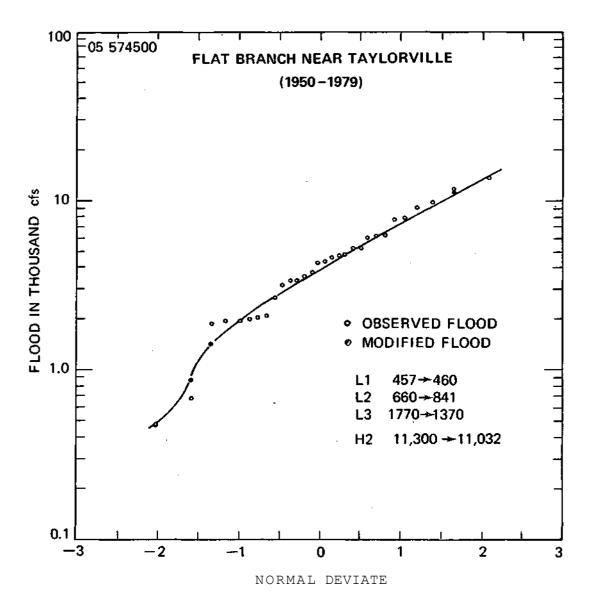


Figure 21. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Flat Branch near Taylorville

 $\mu_2 = 3.612$. The σ_1 is also smaller than σ_2 . Because of the small value of a and the large difference in μ_1 and μ_2 , the effect of the first component distribution is felt only in the beginning position of the fitted distribution as shown in figure 21. The LP3 cannot fit such a probability curve and the PT method may not be precise because the power transformed series is not exactly symmetrical (5th moment is not close to zero). Therefore, the flood estimates with the MD method are considered better than with the others.

4. Horse Creek near Keenes: High Outlier

The results obtained with the computer program are given in Table 20. The NO equals [19/10] or 1. Only one high outlier is indicated and it is modified from its value of 17,100 cfs to 9,170 cfs in the 4th window and 7,889 cfs in the 5th window. The LP3 statistics show that skew changes from 0.729 to -0.493 and standard deviation from 0.231 to 0.188 in going from window 0 to 5. The 100- and 1000-year floods are 17,873 and 35,635 cfs in window 0 and 8,858 and 10,756 cfs in window 5. Though the second highest observed flood in 19 years is 5,890 cfs, the 100- and 1000-year flood estimates are much lower than the observed flood of 17,100 cfs.

The PT statistics indicate a decrease in kurtosis from 4.004 in window 0 to 2.885 in window 5 and the corresponding 5th moment values are 2.264 and 0.877. The MD estimates of 100- and 1000-year floods are 19,893 and 36,193 cfs in window 0 and 8,476 and 11,823 cfs in window 5.

Because the flood estimates seem rather low and because a 19-year record is quite close to a 20-year record, analyses were made with NO = 2. The results are presented in Table 21. A summary of the flood estimates is:

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STATION NO. 3380475 HORSE CREEK NEAR KEENES DRAINAGE AREA 97.2 Sq Mi Years of Record 19 (1960–1979)									
I	EVEL NO.	0	1	2	3	4	5	6	
	ETHOD ransform, PT			100-Year	Flood i	n cfs			
With k With s	t = 3.0 ample kt	17,246 20,839	17,246 20,839	17,246 20,839		9,704 9,931		8,099 7,900	
LP3, W	nsform Sample skew Weighted skew Distrib., MD	11,191	11,191	11 , 191		9 , 464	8,858 9,150 8,476		
	pe No.	·		ed and M	-	-		·	
Lc	w 1* 2 3 4 5	1,550 2,200 2,220 2,270 2,330	1,550	1,550	1,550	1 , 550	1 , 550	1 , 550	
Hi	gh 5 4 3 2 1*	5,260 5,420 5,840 5,890 17,100	17,100	17 , 100	17,100	9,170	7,889	7,325	
METHOD	STATISTICS			Values	of Stat	tistics			
PΤ	mean std dev skew kurtosis,kt 5th moment lambda	2.729 .030 018 4.004 2.264 345	2.729 .030 018 4.004 2.264 345	.030 018	4.004	4.124 021 3.261	069 2.885	142.134 107	
LP3	mean std dev sample skew kurtosis,kt 5th moment	3.596 .231 .729 5.446 12.918	3.596 .231 .729 5.446 12.918	3.596 .231 .729 5.446 12.918	3.596 .231 .729 5.446 12.918	3.582 .195 300 3.072 -1.109	493 2.830		
MD	weight 'a' mu1 mu2 sigma1 sigma2 Test Stat	.489 3.533 3.657 .005 .311 7.645	.489 3.533 3.657 .005 .311 7.645	.489 3.533 3.657 .005 .311 7.645	.489 3.533 3.657 .005 .311 7.645	.162 3.716 3.556 .034 .202 2.729	3.726 3.514 .069 .187	.723 3.667 3.342 .111 .124 2.235	

STATION NO. 3380475 HORSE CREEK NEAR KEENES DRAINAGE AREA 97.2 Sq Mi Years of Record 19 (1960-1979) VARIOUS RECURRENCE-INTERVAL FLOODS METHOD # Flood in cfs for Recurrence Intervals (Years) 2 25 50 100 10 500 1000 PT, kt=3.0 7,895 10,928 13,771 17,246 28,927 0 3,775 36,270 PT, sample kt 3,775 7,613 11,190 15,146 20,839 48,257 73,728 LP3, sample skew 3,702 8,017 11,265 14,275 17,873 29,158 35,635 weighted skew 4,121 7,536 9,093 10,174 11,191 13,353 14,211 MD, mixed dist. 3,429 8,385 12,512 16,017 19,893 30,634 36,193 PT, kt=3.0 1 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 2 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PΤ, kt=3.0 3 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 4 3,918 6,635 7,908 8,820 9,704 1.1,688 12,525 PT, sample kt 12,224 3,918 6,596 7,944 8,941 9,931 13,216 LP3, sample skew 3,906 6,676 7,981 8,913 9,813 11,813 12,645 7,837 8,674 11,156 11,835 weighted skew 3,937 6,630 9,464 MD, mixed dist. 4,034 6,256 7,823 9,051 10,308 13,414 14,846 PT, kt=3.0 5 6,309 3,972 7,284 7,947 8,565 9,881 10,409 PT, sample kt 3,972 6,323 7,269 7,903 8,488 9,719 10,208 LP3, sample skew 3,926 6,420 7,481 8,197 8,858 10,225 10,756 weighted skew 3,898 6,462 7,607 8,401 9,150 10,756 11,401 MD, mixed dist. 4,081 6,196 7,002 7,672 8,476 10,758 11,823 PT, kt=3.0 6 3,999 6,156 7,009 7,577 8,099 9,185 9,613 7,900 8,783 PT, sample kt 3,999 6,193 6,967 7,460 9,121 LP3, sample skew 3,930 6,310 7,277 7,915 8,491 9,649 10,085 8,278 weighted skew 7,503 9,009 10,576 11,207 3,879 6,386 MD, mixed dist. 4,102 6,135 6,985 7,583 8,159 9,445 9,987

	STATION NO. 3380475 HORSE CREEK NEAR KEENES DRAINAGE AREA 97.2 Sq Mi Years of Record 19 (1960–1979)									
I	EVEL NO.	0	1	2	3	4	5	6		
	ETHOD			100-Year	Flood i	n cfs				
	ransform, PT $t = 3.0$	17 , 246	17 , 452	17 , 770	18,006	13,332	10 , 975	9,172		
With s Log Tra	ample kt nsform	20,839	20,922	21,108	21,273	14,194	11,105	9,012		
LP3, S	Sample skew	17,873	17 , 971	18,144	18,285	13,277	11,078	9,486		
	Weighted skew Distrib., MD	11,191 19,893		11,458 19,906	11,558 19,932	10,640 13,897	10,084 11,388	9,531 9,258		
	rpe No.	,		ed and M	-			·		
Lo	w 1*	1,550	1,550	1,550	1,550	1,550	1,550	1,550		
	2* 3	2,200 2,220	2,200	2,200	2,200	2,171	2,107	2,040		
	4 5	2,270 2,330								
ц	.gh 5	5,260								
111	- 4	5,420								
	3 2*	5,840 5,890	6,311	6,846	7,199	7,633	7,633	6,913		
	1*	17,100	17,100	17,100	17,100	12,166	9,694	8,090		
METHOD	STATISTICS			Values	of Stat	istics				
PT	mean	2.729	2.756	2.769	2.770	5.524	15.012	85.919		
	std dev skew	.030 018	.031 015	.032 011	.032 008	.209 .005	1.410 013	16.962 060		
	kurtosis, kt	4.004	3.948	3.877	3.834	3.450	3.105	2.805		
	5th moment	2.264	2.174	2.029	1.929	1.585	1.048	.550		
	lambda	345	341	339	339	105	.132	.442		
LP3	mean	3.596	3.598	3.600	3.601	3.594	3.588	3.581		
	std dev	.231	.232	.234	.235	.217	.206	.196		
	sample skew kurtosis,kt	.729 5.446	.710 5.317	.691 5.162	.683 5.067	.155 3.615	154 3.027	413 2.767		
	5th moment	12.918	12.343	11.659	11.243	3.210	225	-2.307		
MD	weight 'a'	.489	.476	.463	.454	.187	.090	.786		
	mul	3.533	3.533	3.532	3.531	3.544	3.729	3.655		
	mu2	3.657	3.657	3.658	3.659	3.606	3.574	3.309		
	sigma1	.005	.005	.005	.005	.005	.050	.139		
	sigma2	.311	.309	.307	.306 6.937	.239	.211	.121		
* [Test Stat	7.645	7.403	7.122			2.883 I modific	2.309		

STATION NO.	3380475	HORSE	CREEK	NEAR	KEENES		
DRAINAGE AREA	97.2	Sq Mi	Years	of Re	ecord	19	(1960-1979)

VARIOUS RECURRENCE-INTERVAL FLOODS

METHOD #	Flc	od in c	fs for H	Recurrenc	e Interv	als (Yea	rs)
	2	10	25	50	100	500	1000
PT, kt=3.0 0	3,775	7,895	10,928	13,771	17,246	28,927	36,270
PT, sample kt	3,775	7,613	11,190	15,146	20,839	48,257	73,728
LP3, sample skew	3,702	8,017	11,265	14,275	17,873	29,158	35,635
weighted skew	4,121	7,536	9,093	10,174	11,191	13,353	14,211
MD, mixed dist.	3,429	8,385	12,512	16,017	19,893	30,634	36,193
PT, kt=3.0 1	3,789	7,960	11,037	13,923	17,452	29,317	36,773
PT, sample kt	3,789	7,688	11,296	15,258	20,922	47,697	72,034
LP3, sample skew	3,720	8,076	11,344	14,367	17,971	29,241	35,689
weighted skew	4,137	7,593	9,174	10,273	11,308	13,511	14,387
MD, mixed dist.	3,425	8,460	12,564	16,036	19,866	30,476	35,953
PT, kt=3.0 2	3,803	8,043	11,187	14,144	17,770	30,019	37,751
PT, sample kt	3,803	7,785	11,440	15,434	21,108	47,482	71,037
LP3, sample skew	3,740	8,154	11,458	14,509	18,144	29,483	35,956
weighted skew	4,155	7,666	9,278	10,400	11,458	13,714	14,613
MD, mixed dist.	3,421	8,560	12,645	16,094	19,906	30,422	35,824
PT, kt=3.0 3	3,811	8,098	11,290	14,303	18,006	30,577	38,553
PT, sample kt	3,811	7,848	11,539	15,567	21,273	47,611	70,982
LP3, sample skew	3,752	8,206	11,540	14,618	18,285	29,718	36,244
weighted skew	4,167	7,715	9,347	10,485	11,558	13,849	14,763
MD, mixed dist.	3,413	8,626	12,699	16,138	19,932	30,371	35,742
PT, kt=3.0 4	3,879	7,507	9,674	11,438	13,332	18,322	20,763
PT, sample kt	3,879	7,401	9,779	11,843	14,194	20,954	24,557
LP3, sample skew	3,877	7,504	9,660	11,408	13,277	18,160	20,529
weighted skew	4,075	7,238	8,681	9,688	10,640	12,682	13,501
MD, mixed dist.	3,548	7,636	10,013	11,906	13,897	18,986	21,407
PT, kt=3.0 5	3,929	7,039	8,620	9,799	10,975	13,738	14,950
PT, sample kt	3,929	7,018	8,640	9,867	11,105	14,057	15,367
LP3, sample skew	3,922	7,060	8,670	9,873	11,078	13,910	15,154
weighted skew	4,005	6,942	8,275	9,205	10,084	11,974	12,735
MD, mixed dist.	3,983	6,832	8,578	9,962	11,388	14,925	16,554
PT, kt=3.0 6 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	3,973 3,973 3,931 3,927 4,045	6,639 6,645	7,631 7,852 7,871	8,347 8,693 8,724	9,172 9,012 9,486 9,531 9,258	10,765 10,423 11,184 11,267 11,107	11,415 10,987 11,865 11,967 11,907

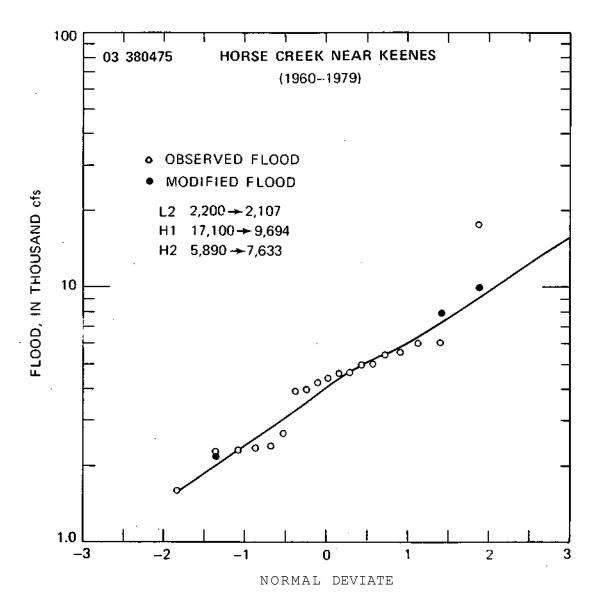


Figure 22. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Horse Creek near Keenes

	100-year	flood	1000-year	flood
	window 0	window 5	window 0	window 5
PT, kt = 3.0	17,246	10,975	36,270	14,950
Sample kt	20,839	11,105	73,728	15 , 367
LP3, sample g	17,873	11,078	35,635	15 , 154
weighted g	11,191	10,084	14,211	12,735
MD	19,893	11,388	36,193	16,554

With NO = 2, the changes in distribution statistics are less than with NO = 1. Also, the flood estimates are more in line with those indicated by storm frequency and runoff conditions (Singh, 1980). The observed floods as well as the modified floods in the 5th window and the probability curve fitted by the MD method are shown in figure 22.

5. Sangamon River near Oakley: High Inlier

The results obtained with the computer program are given in Table 22. The NO equals [27/10] or 2. Only one high inlier is indicated and it is modified from its value of 16,000 cfs in window 0 to 20,085 cfs in window 5. An insignificant low inlier is also indicated. The value changes from 2,390 cfs to 2,321 cfs in window 5. The LP3 statistics show that skew increases from 0.398 to 0.486 and standard deviation from 0.250 to 0.258 in going from window 0 to 5. The 100- and 1000-year floods, with LP3 and sample skew, increase from 25,630 and 46,891 cfs in window 0 to 28,136 and 54,588 cfs in window 5.

The PT statistics indicate an increase in kurtosis from 2.202 in window 0 to 2.311 in window 5 and the corresponding 5th moment values are 0.322 and 0.403. The 100- and 1000-year flood estimates with PT and sample kurtosis are 23,346 and 35,558 cfs for window 0 and 27,145 and 48,531 cfs for window 5. The flood estimates with kurtosis = 3.0 are much higher. The MD

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STATION NO. 5572500 SANGAMON RIVER NEAR OAKLEY DRAINAGE AREA 774.0 Sq Mi Years of Record 27 (1951-1977) LEVEL NO. 0 1 2 3 4 5 6. METHOD 100-Year Flood in cfs Power Transform, PT With kt = 3.032,191 32,191 32,191 32,191 34,458 36,775 38,599 With sample kt 2.3,346 23,346 23,346 23,346 24,824 27,145 30,550 Log Transform LP3, Sample skew 25,630 25,630 25,630 25,630 26,685 28,136 30,069 LP3, Weighted skew 18,472 18,472 18,472 18,472 18,840 19,357 20,109 Mixed Distrib., MD 21,888 21,888 21,888 21,888 22,929 24,557 26,840 No. Observed and Modified Floods in cfs Type 2,390 Low 1* 2,390 2,390 2,390 2,390 2,321 2,191 2,660 2* 2,660 2,660 2,660 2,660 2,660 2,565 3 3,020 4 3,020 5 3,120 5 11,800 High 4 13,200 3 13,700 2* 15,300 15,300 15,300 15,300 15,300 15,300 15,619 1* 16,000 16,000 16,000 16,000 17,766 20,085 22,841 METHOD STATISTICS Values of Statistics 2.409 2.409 2.409 2.225 2.287 \mathbf{PT} 2.409 2.282 mean std dev .017 .017 .017 .017 .014 .013 .015 .093 skew .093 .093 .093 .098 .097 .086 2.455 kurtosis, kt 2.202 2.202 2.202 2.202 2.232 2.311 5th moment .322 .322 .322 .322 .388 .403 .325 lambda -.402 -.402 -.402 -.402 -.427 -.439 -.426 3.755 3.755 3.755 3.755 3.757 3.758 3.759 LP3 mean .250 .250 .250 .250 .253 .258 .266 std dev sample skew .398 .398 .398 .398 .439 .486 .528 2.309 2.309 2.395 2.551 kurtosis, kt 2.309 2.309 2.773 5th moment. 1.997 1.997 2.934 3.612 1.997 1.997 2.373 weight 'a' MD .378 .378 .378 .378 .385 .362 .391 3.523 3.523 3.523 3.523 3.531 3.536 3.558 mu1 mu2 3.896 3.896 3.896 3.896 3.898 3.885 3.889 sigma1 .089 .089 .089 .089 .095 .097 .120 .207 .207 sigma2 .207 .207 .216 .253 .235 2.955 2.955 Test Stat 2.955 2.955 2.654 2.331 2.132

STATION NO. 5572500 SANGAMON RIVER NEAR OAKLEY DRAINAGE AREA 774.0 Sq Mi Years of Record 27 (1951-1977)									
	VARIOUS RECURRENCE-INTERVAL FLOODS								
METHOD	#	Floo 2	od in cf 10	s for Re 25	ecurrence 50	e Interva 100	als (Yea 500	ers) 1000	
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	נא נא נא	5,349 5,479 5,902	12 , 750	16,874 16,795 14,409	20,909 16,466	23,346 25,630 18,472	66,028 31,621 39,454 22,979 28,984	46,891 24,866	
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	1			SAME	AS ABOVE	2			
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	2			SAME	AS ABOVE]			
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	3			SAME	AS ABOVE	2			
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	, רא נא נא	5,342 5,476 5,927	1.1,736		20,995 21,604	24,824	76,489 34,974 41,819 23,511 30,746	40,045 50,090	
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.	ר א ר א ר א	5,338 5,465 5,951	13,106 12,582 11,943	18,104 17,803 14,971	26,421 22,363 22,554 17,186 20,978	36,775 27,145 28,136 19,357 24,557	87,423 41,021 45,122 24,268 33,731	54,588	
PT, kt=3.0 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.		5,339 5,445 5,969	13,302 12,920 12,228		27,522 24,253 23,820 17,790 22,622	38,599 30,550 30,069 20,109 26,840	93,404 51,083 49,547 25,385 37,733	147,001 63,687 60,637 27,618 42,944	

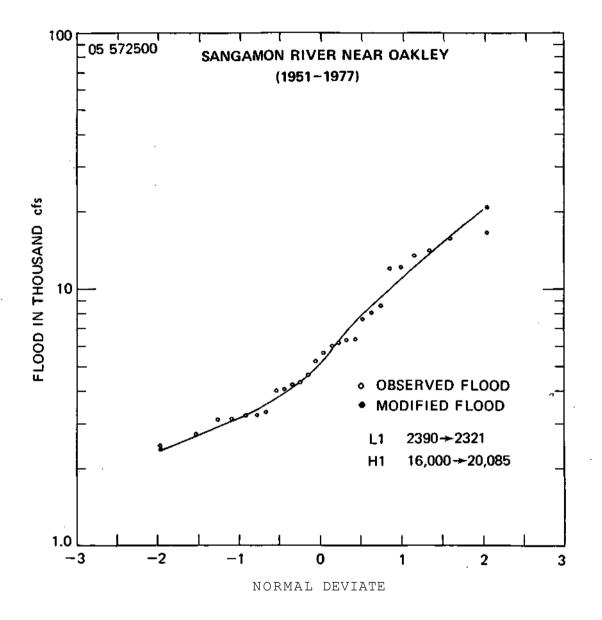


Figure 23. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Sangamon River near Oakley

estimates are 21,888 and 32,290 and 24,557 and 38,116 cfs, respectively. The Q_{100}/Q_2 and Q_{1000}/Q_2 curves in figures 15 and 16 show that the regional estimate lies somewhere in between the MD values for windows 0 and 5, and that estimates by LP3 with sample skew and PT are much higher. The observed floods as well as the modified floods in the 5th window and the probability curve fitted by the MD method are shown in figure 23.

6. Skillet Fork near Wayne City: High Outlier and High Inliers

The results obtained with the computer program are given in Table 23. The NO equals [51/5] or 5. Only one high outlier is indicated but there are 4 high inliers as shown in figure 24. Two rather insignificant low inliers are also detected. The high outlier and inliers are modified as shown below:

	H1	Н2	H3	H4	Н5
Window 0	51,000	22,800	20,000	18,500	18,000
Window 5'	37,862	26 , 139	23,188	21,157	19 , 593

The LP3 statistics show a minor change in skew, from -0.394 to -0.467 and in standard deviation from 0.334 to 0.336. The 100- and 1000-year floods change from 36,719 and 54,194 cfs in window 0 to 35,593 and 50,946 cfs in window 5. The change is rather small.

The PT statistics show that kurtosis decreases from 3.495 in window 0 to 2.884 in window 5 and the 5th moment decreases from 1.317 to -0.056. The corresponding 100- and 1000-year flood estimates with PT and sample kurtosis change from 40,644 and 69,165 cfs to 35,230 and 51,250 cfs. The MD estimates change from 39,943 and 66,909 cfs to 37,723 and 71,302 cfs.

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STATION NO. 3380500 SKILLET FORK NEAR WAYNE CITY DRAINAGE AREA 464.0 Sq Mi Years of Record 51 (1929-1979) 12 3 4 5 6 LEVEL NO. 0 METHOD 100-Year Flood in cfs Power Transform, PT With kt = 3.037,886 37,886 38,215 38,322 36,859 35,883 35,682 40,644 40,644 40,896 40,468 37,167 35,230 With sample kt 34,286 Log Transform LP3, Sample skew 36,719 36,719 37,068 37,322 36,275 35,593 35,852 36,897 37,165 LP3, Weighted skew 36,642 36,642 36,782 36,498 36,454 Mixed Distrib., MD 39,943 39,943 40,276 40,421 40,856 37,723.36,132 Observed and Modified Floods in cfs Туре No. 1* 858 858 858 858 858 858 927 Low 2* 1,450 1,450 1,450 1,450 1,450 1,450 1,471 3* 2,110 2,110 2,110 2,110 2,110 2,110 2,071 4* 2,860 2,860 2,860 2,842 2,654 2,518 2,415 5* 3,040 3,040 3,040 3,040 2,955 2,828 2,727 5* 18,000 18,000 18,000 18,482 19,240 19,593 19,895 High 4* 18,500 18,500 19,103 19,874 20,761 21,157 21,487 3* 20,000 20,000 20,778 21,673 22,760 23,188 23,579 2* 22,800 22,800 23,022 24,162 25,597 26,139 26,641 1* 51,000 51,000 51,000 49,290 41,984 37,862 35,540 METHOD STATISTICS Values of Statistics 19.560 20.012 24.431 28.125 28.459 PΤ 19.883 19.883 mean std dev 3.123 3.123 3.050 3.172 4.321 5.340 5.421 .007 skew .007 .005 .002 -.011 -.020 -.027 kurtosis, kt 3.495 3.495 3.471 3.370 3.055 2.884 2.758 5th moment 1.317 1.317 1.197 .906 .243 -.056 -.173 lambda .158 .158 .155 .159 .194 .218 .220 LP3 mean 3.886 3.886 3.887 3.888 3.888 3.887 3.887 .334 std dev .334 .334 .336 .336 .335 .336 sample skew -.394 -.394 -.388 -.389 -.437 -.467 -.444 kurtosis,kt 3.599 3.599 3.583 3.526 3.375 3.301 3.135 5th moment -3.744 -3.744 -3.705 -3.814 -4.412 -4.679 -4.173 weight 'a' .108 .098 .598 MD .108 .135 .206 .612 mu1 3.454 3.454 3.425 3.763 3.472 3.618 3.781 3.939 mu2 3.939 3.937 3.953 3.958 4.053 4.070 sigma1 .357 .357 .349 .337 .367 .364 .350 sigma2 .290 .290 .292 .190 .202 .287 .288 6.254 6.254 Test Stat 6.121 5.830 5.243 4.708 4.249

STATION NO. 3380500 SKILLET FORK NEAR WAYNE CITY DRAINAGE AREA 464.0 Sq Mi Years of Record 51 (1929–1979)									
VARIOUS RECURRENCE-INTERVAL FLOODS									
METHOD #				ecurrenc		•			
	2	10	25	50	100	500	1000		
PT, kt=3.0 C) 8,054	19,797	26,715	32,187	37,886	52,054	58,572		
PT, sample kt	8,054	19,422	27,065	33,527	40,644	59,739	69,165		
LP3, sample skew	8,098	19,837	26,478	31,568	36,719	48,884	54,194		
weighted skew	8,101	19,829	26,450	31,519	36,642	48,723	53,987		
MD, mixed dist.	8,007	19,623	27,031	33,219	39,943	57,990	66,909		
PT, kt=3.0 1 PT, sample kt									
LP3, sample skew weighted skew MD, mixed dist.			SAME	AS ABOV	Ε				
PT, kt=3.0 2	2 8,062	19,892	26,888	32,432	38,215	52,622	59,262		
PT, sample kt	8,062	19,531	27,232	33,734	40,896	60,092	69,563		
LP3, sample skew	8,104	19,931	26,656	31,825	37,068	49,496	54,939		
weighted skew	8,113	19,914	26,594	31,715	36,897	49,136	54,477		
MD, mixed dist.	8,025	19,716	27,202	33,450	40,276	58,578	67,630		
PT, kt=3.0 3	8 8,092	19,978	26,991	32,540	38,322	52,701	59,318		
PT, sample kt	8,092	19,695	27,272	33,594	40,468	58,604	67,396		
LP3, sample skew	8,125	20,037	26,818	32,032	37,322	49,864	55,358		
weighted skew	8,133	20,021	26,761	31,931	37,165	49,534	54,934		
MD, mixed dist.	8,080	19,898	27,401	33,628	40,421	58,549	67,476		
PT, kt=3.0 4	4 8,161	19,763	26,391	31,551	36,859	49,810	55,669		
PT, sample kt	8,161	19,721	26,439	31,714	37,167	50,611	56,733		
LP3, sample skew	8,165	19,908	26,417	31,344	36,275	47,720	52,630		
weighted skew	8,140	19,963	26,606	31,674	36,782	48,763	53,957		
MD, mixed dist.	8,028	19,939	27,555	33,929	40,856	59,518	68,773		
PT, kt=3.0 5	5 8,196	19,590	25,963	30,873	35,883	47,960	53,365		
PT, sample kt	8,196	19,679	25,861	30,529	35,230	46,354	51,250		
LP3, sample skew	8,176	19,799	26,135	30,881	35,593	46,379	50,946		
weighted skew	8,131	19,899	26,474	31,474	36,498	48,231	53,296		
MD, mixed dist.	8,457	18,873	25,197	30,881	37,723	59,361	71,302		
PT, kt=3.0 6	6 8,197	19,531	25,856	30,721	35,682	47,625	52,964		
PT, sample kt	8,197	19,710	25,623	29,977	34,286	44,245	48,531		
LP3, sample skew	8,151	19,782	26,191	31,025	35,852	47,008	51,776		
weighted skew	8,121	19,848	26,41.6	31,418	36,454	48,244	53,346		
MD, mixed dist.	8,407	19,150	25,281	30,392	36,132	53,146	62,528		

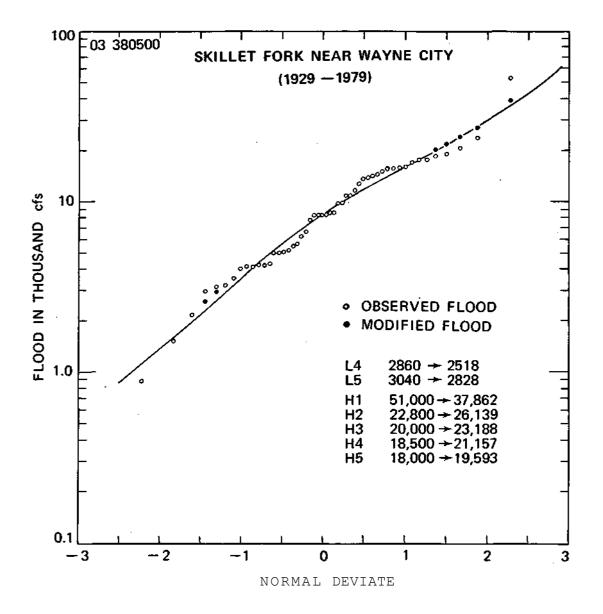


Figure 24. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Skillet Fork near Wayne City

For this basin, the effect of a rather high outlier is largely balanced by 4 high inliers. The 100-year flood estimates with different methods are very close but the 1000-year flood with the MD is about 1.3 to 1.4 times that from the others. The top flood of 51,000 cfs was caused by a 2-3 day storm producing about 10 inches of catchment rainfall; the estimated recurrence interval is 300 to 500 years. The MD gives a 500-year flood of 59,361 cfs and the observed top flood of 51,000 cfs would correspond to somewhat higher than a 300-year flood.

7. Kishwaukee River near Perryville: Low Inlier and High Inlier

The results obtained with the computer program are given in Table 24. The NO equals [40/4] = 4. One significant and one insignificant low inliers and one insignificant low outlier, and one significant and three less significant high inliers are shown in figure 25.

The LP3 statistics show a minor change in skew, from -0.541 to -0.601, and in standard deviation, from 0.282 to 0.301. The 100- and 1000-year floods are 24,980 and 32,832 cfs in window 0 and 26,412 and 34,545 cfs in window 5, with the sample skew. Modification of low inliers generally reduces the skew and of high inliers increases the skew. When both low and high inliers are present, the opposite effects are cancelled to some extent.

The PT statistics show that kurtosis increases from 1.912 in window 0 to 2.254 in window 5 and the absolute value of the 5th moment decreases from 0.685 to 0.373. A summary of 100- and 1000-year floods with different methods is given on the next page.

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STATION NO. 5440000 KISHWAUKEE RIVER NEAR PERRYVILLE DRAINAGE AREA 1099.0 Sq Mi Years of Record 40 (1940-1979)

L	EVEL NO.	0		12	3	4	5	6
	ETHOD			100-Year	Flood i	in cfs		
With k	ransform, PT t = 3.0 ample kt	22,164 18,562	22,164 18,562	22,409 18,864	22,593 19,261	23,519 20,577	24,682 22,140	25,953 23,939
LP3, S LP3, W	Sample skew Weighted skew Distrib., MD	24,980 26,385 18,665	24,980 26,385 18,665	25,179 26,515 18,663	25,232 26,822 18,930	25,741 27,682 20,605	26,412 28,698 22,608	27,034 29,960 25,516
Tyj	pe No.		Observ	ved and N	Modified	Floods :	in cfs	
Lot	w 1* 2* 3* 4* 5	2,020 2,080 2,340 2,360 2,620	2,020 2,080 2,340 2,360	2,020 2,080 2,340 2,360	1,789 2,080 2,340 2,360	1,483 2,080 2,340 2,403	1,281 1,980 2,340 2,505	1,096 1,787 2,299 2,582
Hio	gh 5 4* 3* 2* 1*	14,800 14,800 15,200 16,400 16,700	14,800 15,200 16,400 16,700	14,800 15,200 16,400 17,449	14,800 15,200 16,400 18,199	14,958 16,000 17,393 19,690	15,702 16,884 18,514 21,285	16,492 17,842 19,732 23,078
METHOD	STATISTICS			Values	of Stat	tistics		
РТ	mean std dev skew kurtosis,kt 5th moment lambda	224.414 71.137 158 1.912 685 .534	224.414 71.137 158 1.912 685 .534	201.455 62.554 152 1.936 633 .519	198.579 62.003 148 1.983 603 .517	154.926 46.616 130 2.110 491 .482	118.079 33.925 108 2.254 373 .443	95.519 26.616 088 2.426 268 .412
LP3	mean std dev sample skew kurtosis,kt 5th moment						3.856 .301 601 2.726 -4.093	
MD	weight 'a' mu1 mu2 sigma1 sigma2 Test Stat	.215 .086 3.651	.524 3.638 4.094 .215 .086 3.651	.220 .082 3.393	.563 3.661 4.103 .234 .082 2.847	4.099 .258 .098 2.102	4.094 .285 .108	2.381

STATION NO. 5440000 KISHWAUKEE RIVER NEAR PERRYVILLE DRAINAGE AREA 1099.0 Sq Mi Years of Record 40 (1940-1979)

VARIOUS RECURRENCE-INTERVAL FLOODS

METHOD #	Fl	ood in c	fs for R	ecurrenc	e Interv	als (Yea	rs)
	2	10	25	50	100	500	1000
PT, kt=3.0 0	7,930	14,948	18,025	20,154	22,164	26,504	28,268
PT, sample kt	7,930	15,454	17,108	17,934	18,562	19,616	19,962
LP3, sample skew	7,587	15,705	19,608	22,364	24,980	30,593	32,832
weighted skew	7,496	15,882	20,171	23,312	26,385	33,301	36,194
MD, mixed dist.	8,723	14,623	16,370	17,551	18,665	21,149	22,230
PT, kt=3.0 1 PT, sample kt LP3, sample skew weighted skew MD, mixed dist.			SAME	AS ABOV	Έ		
PT, kt=3.0 2	7,921	15,027	18,167	20,347	22,409	26,877	28,697
PT, sample kt	7,921	15,523	17,271	18,168	18,864	20,049	20,446
LP3, sample skew	7,590	15,764	19,716	22,515	25,179	30,916	33,214
weighted skew	7,503	15,931	20,250	23,416	26,515	33,498	36,421
MD, mixed dist.	8,582	14,700	16,411	17,566	18,663	21,159	22,296
PT, kt=3.0 3	7,920	15,105	18,288	20,499	22,593	27,130	28,980
PT, sample kt	7,920	15,574	17,458	18,464	19,261	20,658	21,140
LP3, sample skew	7,598	15,845	19,802	22,591	25,232	30,874	33,115
weighted skew	7,496	16,046	20,440	23,664	26,822	33,939	36,918
MD, mixed dist.	8,520	14,781	16,545	17,755	18,930	21,825	23,359
PT, kt=3.0 4	7,911	15,445	18,853	21,243	23,519	28,494	30,538
PT, sample kt	7,911	15,861	18,152	19,473	20,577	22,617	23,351
LP3, sample skew	7,637	16,135	20,199	23,051	25,741	31,447	33,695
weighted skew	7,515	16,380	20,977	24,361	27,682	35,186	38,334
MD, mixed dist.	8,468	15,175	17,408	18,997	20,605	24,966	27,499
PT, kt=3.0 5	7,906	15,862	19,554	22,171	24,682	30,228	32,527
PT, sample kt	7,906	16,223	18,987	20,676	22,140	24,987	26,051
LP3, sample skew	7,685	16,486	20,695	23,641	26,412	32,256	34,545
weighted skew	7,545	16,774	21,610	25,182	28,698	36,667	40,018
MD, mixed dist.	8,428	15,498	18,199	20,286	22,608	30,303	34,842
PT, kt=3.0 6	7,903	16,321	20,323	23,186	25,953	32,125	34,705
PT, sample kt	7,903	16,617	19,900	22,025	23,939	27,833	29,340
LP3, sample skew	7,739	16,880	21,214	24,224	27,034	32,886	35,146
weighted skew	7,566	17,253	22,390	26,200	29,960	38,509	42,111
MD, mixed dist.	8,488	15,644	18,914	21,803	25,516	38,081	44,734

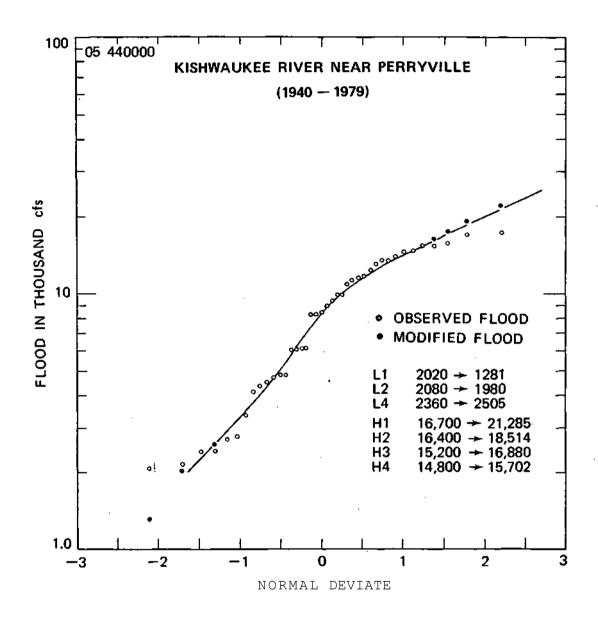


Figure 25. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Kishwaukee River near Perryville

	100-year		-	1000-year flood				
	window 0	window 5	window 0	window 5				
PT, kt = 3.0	22,164	24,682	28,268	32,527				
PT, sample kt	18,562	22,140	19,962	26,051				
LP3, sample skew	24,980	26,412	32,832	34,545				
LP3, weighted skew	26,385	28,698	36,194	40,018				
MD	18,665	22,608	22,230	34,842				

When both low and high inliers are present, the flood estimates are less sensitive to the modification of inliers for LP3 than with the MD.

8. Sangamon River at Riverton: Low Outliers and High Outliers

The results obtained with the computer program are given in Table 25. The NO equals [67/10] or a maximum of 5. Four low outliers and three high outliers (out of which H1 is a very significant high outlier) are shown in Figure 26. The modified values for these outliers in window 5 are also shown in the figure.

The LP3 statistics show that the skew decreases from -1.227 in window 0 to -1.386 in window 5 and the standard deviation decreases from 0.312 to 0.291. With sample skew, the 100- and 1000-year floods of 38,917 and 42,416 cfs in window 0 are replaced by 33,931 and 35,780 cfs, which are much lower than the observed flood of 68,700 cfs.

The PT statistics indicate that kurtosis and 5th moment decrease from 4.538 and 3.513 in window 0 to 3.231 and 0.141 in window 5. The 100- and 1000-year floods change from 54,386 and 84,445 cfs in window 0 to 41,641 and 53,410 cfs in window 5, with sample kurtosis. However, the corresponding MD estimates change from 53,173 and 123,725 cfs in window 0 to 44,041 and 63,018 cfs in window 5. Thus, only the 5th window 1000-year flood with

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STATION NO. 5576500 SANGAMON RIVER AT RIVERTON DRAINAGE AREA 2618.0 Sq Mi Years of Record 67 (1908-1979) 1 2 3 4 5 LEVEL NO. 0 6 METHOD 100-Year Flood in cfs Power Transform, PT With kt = 3.048,573 48,573 46,552 45,177 42,341 40,858 39,667 With sample kt 54,386 54,386 50,740 48,344 43,952 41,641 39,735 Log Transform LP3, Sample skew 38,917 38,917 37,621 36,719 34,823 33,931 34,489 LP3, Weighted skew 46,750 46,750 45,577 44,685 42,686 41,370 40,746 Mixed Distrib., MD 53,173 53,173 49,872 47,912 45,146 44,041 42,643 No. Observed and Modified Floods in cfs Type 1* 1,040 1,040 1,040 1,040 1,040 Low 1,040 1,280 2* 1,830 1,830 1,830 1,830 1,860 2,177 2,668 3* 2,540 2,540 2,540 2,540 2,778 3,151 3,640 4* 2,840 2,840 2,902 3,141 3,523 3,931 4,407 5* 4,260 4,260 4,260 4,260 4,260 4,599 5,060 5* 30,600 30,600 30,600 30,654 30,984 30,984 30,372 High 4* 32,900 32,900 32,900 32,900 32,900 32,900 31,786 3* 41,000 41,000 41,000 40,501 37,018 35,097 33,594 2* 44,200 44,200 44,200 44,200 40,418 37,998 36,154 68,700 68,700 60,108 54,410 47,205 43,513 40,834 1* METHOD STATISTICS Values of Statistics PT122.449 122.449 171.779 217.877 373.890 475.065 518.865 mean std dev 32.354 32.354 48.749 64.479 119.821 154.829 166.183 skew .088 .088 .063 .042 .007 -.016 -.034 kurtosis, kt 4.538 4.538 4.137 3.873 3.468 3.231 3.022 5th moment 3.513 3.513 2.265 1.474 .631 .141 -.123. lambda .408 .408 .453 .484 .553 .583 .594 4.144 4.144 4.143 4.143 4.142 4.144 LP3 4.148 mean std dev .312 .312 .310 .307 .300 .291 .276 sample skew -1.227 -1.227 -1.280 -1.313 -1.381 -1.386 -1.254 kurtosis, kt 5.791 5.791 5.802 5.841 5.968 6.069 5.498 5th moment -15.139 -15.139 -15.893 -16.532 -17.963 -19.160 -16.481 weight 'a' .281 .281 .264 .242 .159 MD .163 .166 mu1 3.904 3.904 3.877 3.847 3.704 3.719 3.753 mu2 4.238 4.238 4.239 4.237 4.228 4.225 4.227 .426 .442 .442 .433 .384 .379 .338 sigmal .182 sigma2 .166 .166 .168 .173 .185 .176 4.054 4.054 4.062 3.972 3.569 3.674 4.127 Test Stat

5576500 SANGAMON RIVER AT RIVERTON STATION NO. Years of Record DRAINAGE AREA 2618.0 Sq Mi 67 (1908-1979) VARIOUS RECURRENCE-INTERVAL FLOODS # Flood in cfs for Recurrence Intervals (Years) METHOD 2 10 25 50 100 500 1000 PT, kt=3.0 30,867 38,233 43,494 48,573 15,289 59,888 64,613 0 15,289 29,673 38,860 46,378 54,386 74,834 84,445 PT, sample kt LP3, sample skew 16,070 30,244 34,683 37,090 38,917 41,665 42,416 weighted skew 15,424 31,891 38,665 42,988 46,750 53,758 56,174 15,757 28,378 35,849 43,004 53,173 97,011 123,725 MD, mixed dist. PT, kt=3.0 1 PT, sample kt LP3, sample skew SAME AS ABOVE weighted skew MD, mixed dist. PT, kt=3.0 15,374 30,264 37,105 41,932 46,552 56,718 60,918 2 15,374 29,405 PT, sample kt 37,632 44,081 50,740 67,003 74,340 LP3, sample skew 16,119 29,784 33,867 36,022 37,621 39,941 40,551 weighted skew 15,439 31,529 37,997 42,070 45,577 51,997 54,171 MD, mixed dist. 15**,**757 28,300 35,391 41,698 49,872 85,042 107,980 15,429 29,823 36,311 PT, kt=3.0 3 40,854 45,177 54,615 58,489 PT, sample kt 15,429 29,182 36,746 42,513 48,344 62,162 68,247 LP3, sample skew 16,150 29,429 33,273 35,264 36,719 38,781 39,309 31,224 weighted skew 37,462 41,356 44,685 50,713 52,731 15,452 MD, mixed dist. 15,715 28,304 35,194 41,009 47,912 74,905 93,932 PT, kt=3.0 15,520 34,617 4 28,842 38,597 42,341 50,390 53,649 PT, sample kt 15,520 28,516 34,876 39,467 43,952 54,017 58,260 LP3, sample skew 16,189 28,604 31,963 33,638 34,823 36,812 36,424 weighted skew 15,460 30,479 36,213 39,727 42,686 47,927 49,642 MD, mixed dist. 15,540 28,240 34,844 39,911 45,146 58,361 64,829 PT, kt=3.0 15,563 28,268 33,682 37,388 40,858 48,271 51,256 5 PT, sample kt 15,563 28,114 33,816 41,641 49,985 53,410 37,819 LP3, sample skew 16,187 28,077 31,249 32,822 33,931 35,421 35,780 35,292 weighted skew 29,872 15,477 38,596 41,370 46,260 47,854 MD, mixed dist. 15,515 27,798 34,155 39,022 44,041 56,752 63,018 PT, kt=3.0 15,572 27,739 36,389 39,667 6 32,881 46,648 49,452 15,572 27,728 32,897 PT, sample kt 36,432 39,735 46,800 49,641 15,990 27,804 31,278 LP3, sample skew 33,117 34,489 36,504 37,042 weighted skew 15,403 29,196 34,548 37,886 40,746 45,961 47,726 MD, mixed dist. 15,525 27,384 33,423 37,995 42,643 54,073 59,419

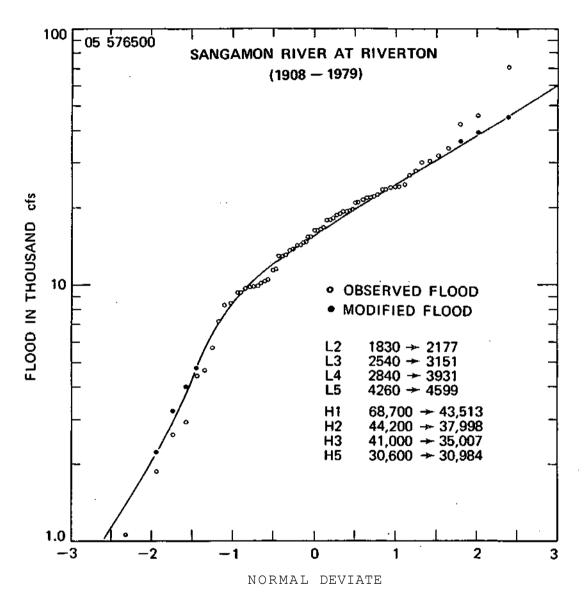


Figure 26. Observed and modified floods and the fitted mixed distribution curve (window 5) for the Sangamon River at Riverton

MD is close to the observed top flooded 68,700 cfs. The other methods yield estimates varying from 35 to 53 thousand cfs.

When the sample skew is very small in the algebraic sense, the LP3 flood estimates of 500, 1000, or higher recurrence-interval floods are not much higher than the 100-year flood. The Sangamon River at Riverton has flood data for 67 years. The MD estimates are considered better than those from the other four methods. However, the MD flood estimate in window 0 for high recurrence-interval floods can be very high. The MD does give good results after the outliers/inliers have been modified.

CONCLUSIONS

The main objectives of this study were: 1) the development of satisfactory tests for detecting outliers and inliers at various levels of significance in the two extreme tails of a suitably transformed flood series; 2) the extensive testing of available transformations in converting a number of observed flood series to series distributed approximately as N (μ , σ^2) and to determine the best transformation for general use; 3) the development and computerization of a flood-frequency methodology that not only detects and modifies outliers/inliers at different levels but also computes 2-year to 1000-year floods at those levels with the power transformation, log-Pearson type III, and mixed distribution methods; and 4) the overall conceptualization, theoretical basis, testing, and validation of a versatile and accurate new flood frequency method. These objectives have been met satisfactorily by the research, analyses, and comparative studies contained in this report. Some main conclusions, derived from this study, are given below.

1. An extensive testing of four methods or algorithms, for generating normally distributed random numbers, regarding their suitability, stability, and effectiveness in generating such numbers has indicated the Polar Method by Box, Muller, and Marsaglia to be the best.

2. Departure has been defined as the standard normal deviate corresponding to the plotting position of the high or low point of the series under consideration, minus the sample standard deviate for that point. The higher the absolute value of the departure, the more severe is the outlier/ inlier. The distribution of the departures for up to 5 points on both the

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high and the low end of various sample sizes has been determined from thousands of generated series. Both an extensive and a compact departure table have been developed for general testing of outliers at 0.01, 0.05, 0.10, 0.20, 0.30, and 0.40 levels of significance. Departures for only 0.01 and 0.05 levels are available in the literature for the top outlier and these are within 0.01 of the departures developed in this study. However, the statistical tests for inliers at any significance level and for outliers at 0.10 to 0.40 levels of significance and for up to five outliers/ inliers are not available in the literature at the present time. The developed departures allow a step-by-step detection and modification of outliers/inliers at various levels.

3. Generally, the literature has dealt with outliers - a flood significantly higher than that indicated by the trend of the rest of the data at the high end, or a flood significantly lower than that indicated by the rest of the data at the low end. The introduction and designation of inliers - a flood lower than that indicated by the rest of the data at the high end or higher than that indicated at the low end - in this study is a welcome addition and fills the information gap. Statistically, both outliers and inliers can occur. However, the absolute value of departure for an inlier is generally less than that for an outlier because the inlier cannot be less than the next lower flood in ranked series at the high end or more than the next higher flood at the low end.

4. Transformation of an observed flood series to an approximately normally distributed series is necessary for checking any outliers/inliers with statistical tests developed in this study. Three transformations power, Wilson-Hilferty, and 3-parameter lognormal - were tested on 28 flood

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series. The results indicate that the power transformation is superior to the others in terms of yielding consistent and satisfactory statistical parameters for the transformed series. Values of g in Table 26 for the power transformed series are very close to zero and those for the 5th are considerably lower than the values with log transformation only (e.g., for LP3).

5. Flood frequency methods have been put together in a computer program. These methods include power transformation method with kurtosis equal to 3.0 as for a normal distribution as well as with sample kurtosis, log-Pearson type III method with the sample skew as well as the weighted skew, and the mixed distribution. The kurtosis correction with the power transformation method is satisfactory if the transformed series approximates a symmetrical distribution. The relevant distribution statistics and measures of goodness of fit with the observed flood series and with the series after modification of outliers/inliers at various levels, as well as the 2-year to 1000-year floods at various levels, are presented in a tabular format. The output enables the analyst to follow the detection and modification of outliers/ inliers at various levels and to choose the level he thinks is the best to use for a particular basin.

 Results of flood frequency analyses of 37 observed flood series in Illinois indicate the following:

a) Absolute value of skew, g, with the power transformation (Table 26) is <0.05, 0.05 to 0.10, and 0.10 to 0.20 for 21, 8, and 8 basins in window 0 and for 19, 12, and 6 basins in window 5, respectively. The power transformation reduces the skew close to zero.

b) Kurtosis with the power transformation (Table 26) is < 3 for 26 basins and >3 for 11 basins in window 0, and <3 for 28 basins and >3 for

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Table 26. Values of g, kt, and 5th with Power and Log Transformation

USGS No.	Trans	g	Window 0 kt	5th	g	Window 5 g kt 5t				
SANGAMON RI	IVER BASIN				-					
05 571000	Power	-0.008	2.601	-0.248	-0.007	2.597	-0.262			
	Log	-0.073	2.621	-0.760	-0.070	2.617	-0.751			
05 572000	Power	-0.016	2.726	-0.430	-0.019	2.716	-0.412			
	Log	-0.312	2.991	-3.246	-0.336	3.003	-3.386			
05 572500	Power	0.093	2.202	0.322	0.097	2.311	0.403			
	Log	0.398	2.309	1.997	0.486	2.551	2.934			
05 574000	Power	0.116	2.433	0.765	0.111	2.444	0.716			
	Log	0.788	3.420	6.815	0.771	3.395	6.692			
05 574500	Power	-0.016	3.256	-0.558	-0.022	3.154	-0.513			
	Log	-0.803	4.302	-8.841	-0.717	4.046	-7.872			
05 575500	Power	-0.002	2.939	-0.064	-0.004	2.936	-0.071			
	Log	-0.051	2.944	-0.464	-0.054	2.942	-0.486			
05 576000	Power	-0.053	2.633	-0.014	-0.056	2.620	0.027			
	Log	-0.453	2.840	-3.205	-0.463	2.818	-3.192			
05 576500	Power	0.088	4.538	3.513	-0.016	3.231	0.1*41			
	Log	-1.227	5.791	-15.139	-1.386	6.069	-19.160			
05 577500	Power	-0.032	2.605	-1.034	-0.023	2.575	-0.768			
	Log	-0.325	2.959	-3.463	-0.203	2.736	-2.200			
05 578500	Power	0.020	2.207	0.451	0.018	2.267	0.274			
	Log	0.104	2.263	1.101	0.110	2.308	0.955			
05 579500	Power	0.036	4.045	-1.494	0.005	3.162	-0.593			
	Log	-0.730	5.438	-12.793	0.140	3.136	0.614			
05 580000	Power	0.046	2.596	0.742	0.038	2.671	0.657			
	Log	0.456	3.026	3.977	0.417	3.036	3.726			
05 580500	Power	0.040	2.474	0.946	0.030	2.571	0.835			
	Log	0.345	2.850	3.456	0.299	2.866	3.123			
05 581500	Power	0.012	3.079	0.970	0.015	3.013	0.887			
	Log	0.465	3.630	5.382	0.438	3.498	4.878			

US	GS No.	Trans	g	Window 0 kt	5th	g	Window 5 kt	5th
05	582000	Power Log	-0.014 -0.099	2.464 2.435	0.456 -0.146	-0.016 -0.128	2.484 2.458	0.391 -0.411
05	582500	Power Log	-0.136 -0.525	2.093 2.347	-0.629 -2.687	-0.108 -0.579	2.336 2.687	-0.425 -3.634
05	583000	Power Log	0.018 -0.562	3.820 3.421	3.900 3.027	-0.055 -0.764	2.792 3.317	0.416 -5.653
ROC	CK RIVER	BASIN						
05	435500	Power Log	-0.003 -0.024	2.387 2.387	0.006 -0.115	-0.005 -0.047	2.508 2.512	-0.046 -0.324
05	437000	Power Log	-0.077 -0.345	2.175 2.241	-0.209 -1.640	-0.071 -0.410	2.354 2.504	-0.188 -2.409
05	437500	Power Log	-0.051 -0.258	2.190 2.211	-0.007 -1.187	-0.054 -0.329	2.310 2.414	-0.120 -1.921
05	438250	Power Log	0.004 -0.926	4.245 4.274	1.947 -7.214	-0.064 -1.133	3.676 4.299	0.252 -9.109
05	438500	Power Log	-0.084 -0.310	1.936 2.068	-0.436 -1.593	-0.057 -0.318	2.239 2.379	-0.234 -2.007
05	439500	Power Log	-0.196 -0.869	2.231 2.923	-0.558 -5.292	-0.183 -1.108	2.334 3.632	-0.591 -8.388
05	440000	Power Log	-0.158 -0.541	1.912 2.244	-0.685 -2.744	-0.108 -0.601	2.254 2.726	-0.373 -4.093
05	440500	Power Log	-0.171 -1.011	2.396 3.520	-0.660 -7.537	-0.156 -1.062	2.498 3.797	-0.527 -8.691
05	441000	Power Log	-0.136 -0.755	2.235 3.047	-0.686 -5.650	-0.109 -0.712	2.372 3.039	-0.414 -5.334
05	443500	Power Log	-0.073 -0.789	2.581 3.403	-0.039 -6.455	-0.071 -0.800	2.591 3.462	-0.071 -6.709
05	444000	Power Log	-0.095 -1.049	2.702 4.137	-0.567 -9.658	-0.094 -1.048	2.697 4.175	-0.538 -9.924
05	445500	Power Log	0.002	3.572 3.573	0.525 -1.496	-0.006 -0.348	3.283 3.380	0.316 -2.824
05	446500	Power Log	-0.116 -0.421	1.945 2.167	-0.448 -2.271	-0.087 -0.453	2.205 2.472	-0.252 -2.916

Table 26. Concluded

USGS No.		Trans		Window 0		Window 5					
			g	kt	5th	g	kt	5th			
05	447000	Power Log	-0.115 -1.147	2.783 3.975	-0.494 -8.540	-0.092 -1.106	2.752 3.904	-0.427 -8.381			
05	447500	Power Log	-0.031 -0.940	3.117 4.133	0.036 -8.700	-0.042 -0.956	3.026 4.162	-0.065 -9.123			
05	448000	Power Log	-0.041 -0.317	2.468 2.544	-0.048 -1.824	-0.041 -0.336	2.502 2.580	0.014 -1.958			
LITTLE WABASH RIVER BASIN											
03	379500	Power Log	-0.014 -0.324	2.807 3.132	-0.647 -3.715	-0.015 -0.318	2.778 3.051	-0.472 -3.329			
03	380475	Power Log	-0.018 0.729	4.004 5.446	2.264 12.918	-0.013 -0.154	3.105 3.027	1.048 -0.225			
03	380500	Power Log	0.007 -0.394	3.495 3.599	1.317 -3.744	-0.020 -0.467	2.884 3.301	-0.056 -4.679			
03	381500	Power Log	0.007 -0.126	3.975 4.085	-1.418 -2.896	0.004 0.306	3.215 3.272	-0.377 2.286			

9 basins in window 5. The values range from 1.912 to 2.939 and 3.079 to 4.538 in window 0 and from 2.205 to 2.936 and 3.013 to 3.676 in window 5. The kurtosis range decreases in window 5 becuase of the modification of any outliers and inliers.

	c) Abso	olute	value	of	the	5th	with	the	power	tra	nsform	atio	n ('	Table	26)
are					Num	ıber	of ba	asins	with	the	5th	in	the	range	È
				<0.5	5		0.	5-1.0)		1.0-	2.0		>2	2.0
	Window	0		16)			13			5				3
	Window	5		26	5			11			0				0

The modification of outliers/inliers reduces significantly the absolute value of the 5th. The transformed series in window 5 are closer to normal distribution than are those in window 0.

d) The kurtosis correction with the power transformation method is reasonably valid if the 3rd and higher odd moments are close to zero. Though the values of g are close to zero for a majority of the transformed series, the 5th moment is not. Thus, the power transformed series are generally asymmetrical. The asymmetry is considered in the mixed distribution method.

e) The mixed distribution parameters a, μ_1 , μ_2 , σ_1 , and σ_2 for the 37 study basins in Table 15 and window 5 show that $0.4 \leq a \leq 0.6$ for 12 basins, $|\sigma_1 - \sigma_2| \leq 0.05$ for 1 out of 12 basins with *a* varying from 0.4 to 0.6 and $|\mu_2 - \mu_1| \leq 0.25$ for none out of 12 basins. Thus, the conditions of *a* - 0.5 and $\mu_1 - \mu_2$ or a = 0.5 and $\sigma_1 = \sigma_2$ are not satisfied. The analysis of power transformed series with or without correction for kurtosis is not the best solution because of the apparent asymmetry exhibited by the transformed series. The mixed distribution is the better answer to the problem.

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f) Plots of Q_{100}/Q_2 and Q_{1000}/Q_2 versus drainage area for the Sangamon and Rock River basins, with floods estimated from the mixed distribution and window 5, are well-defined and indicate a decrease in the ratio with increase in drainage area, except for areas less than 200 square miles. For smaller areas, the trend line steepens considerably. Corresponding data points with the LP3 and sample skew exhibit considerable scatter.

g) The flood estimates with the mixed distribution are generally found to be very satisfactory in window 5.

h) In the case of extreme high outliers, the storm statistics for the top 3 to 4 floods may be used in confirming the severity of the outlier with the methodology developed in a previous report (Singh, 1980).

i) The mixed distribution is highly versatile in simulating various observed distribution shapes. The method coupled with the detection and modification of outliers/inliers may perhaps be the best available at the present.

j) The regionalization of skew as recommended by the Water Resources Council and the use of LP3 may not be the best solution for the floodfrequency problem. The analyses presented in this report, together with the values of g in windows 0 and 5, do not suggest that regionalization of skew is worthwhile.

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